Essays in Risk Management

by

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Abstract

This thesis consists of five chapters that examines risk and uncertainty within two frameworks: foreign exchange market and real options. The first chapter is a preliminary part that overviews the structure of thesis. In the second chapter, I examine the impact of scheduled macroeconomic announcements on realised variance in the Canadian dollar/US dollar foreign exchange market. Information shocks as a whole are made up of public information shocks and private information shocks. I measure the public information shocks from the analyst forecast surprise and the private information shocks from volatility sensitivity to liquidity variables. I find that the realised variance is driven mainly by the latter rather than the former. However, my results for the most important announcements are not significant, which might be due to these being well-analysed publicly. Spread, as a proxy of private information shocks, is the most important liquidity measure, showing a significant increase around the arrival of announcements. My results are robust to joint effects of liquidity variables, considering announcements throughout the day (times other than 8:30 announcement), alternative measures of volatility (absolute return and modified absolute return), evaluation of announcements for US and Canada separately, examine the impact of surprise in model, and the economic classification of announcements.

In the third chapter, I aim to evaluate risk and uncertainty using real options technique. I develop a framework to evaluate representative agents' behaviour in a real options switching framework. I set up three models with revertible switching process under uncertainty and solve these using the alternating direction implicit algorithm. The models break down

into: cash-cost model, cash-time model, and projection model. The cash-cost model captures the cash expenses of switching whereas the cash-time model not only captures the cash cost but also the exact time cost, which is critical in horticulture. The projection model presents an approximation of cash-time model that has less computational complexity. The results of my sensitivity analyses indicate that increases in cost, time, volatility, drift, and discount rate have negative impacts on the switch frequency. If the correlation between two crops is positive, it has negative impacts on switch frequency, otherwise it has positive impacts. Differences between the models are more pronounced over longer periods.

In the fourth and fifth chapters, I extend the cash-time model from chapter three to evaluate orchardists' behaviour in the Hawke's Bay region. Chapter four examines the dataset thoroughly and provide a statistical review of orchards that will be modelled in chapter five. Orchardists have the incentive to switch from one type of apple to another as the apple profits change. In my model, orchardists have the option to carry on with the existing apple trees or to switch to competing apple types by uprooting the existing apple trees and planting new ones or grafting on the existing rootstock. The uprooting strategy is relatively expensive but is instantaneous, and results in young (unproductive) apple trees with a long life ahead of them. In contrast, the grafting strategy is less expensive and faster but continues with old trees. I compute the optimal land value at each age of apple trees from one-year to 33-years old. My results show that grafting is the optimal strategy when trees are young, whereas planting becomes optimal when they are old. Examining the apple dataset, I find that orchardists are biased against uprooting and grafting relative to my predictions. The deviation from what my model proposes and what orchardists follow in reality might be due to the assumption of my model and possible factors in the orchards that my model does not capture. My results show that the deviation from optimal policy for small orchardists is not significantly different from large orchardists.

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Chapter 1

Preliminaries

1.1 Abstract

This thesis evaluates risk and uncertainty by employing two approaches: market microstructure and real options framework. The second Chapter is an application of market microstructure approaches to evaluate realised variance being affected by the arrival of scheduled macroeconomic announcements. The third Chapter evaluates uncertainty using the real options approach. I construct a real options problem by deriving a partial differential equation for two competing products and solving it using the alternating direction implicit algorithm. Chapter four examines orchard data in the Hawkes Bay region and provides a statistical overview to evaluate orchardists behaviour in the following Chapter. In Chapter five, I develop the general model from Chapter three to a more complex and dynamic framework and examine if orchardists optimise their lands value according to the real options analysis.

1.2 Thesis Structure

This thesis evaluates risk and uncertainty in two approaches: (1) market microstructure and (2) real options. In the first approach, I evaluate the impact of macroeconomic announcements on realised variance in the Canadian foreign exchange market. Volatility shocks in the foreign exchange market are either due to having access to public information or private information. As a result, traders send bid or ask quotes in a wide range of execution aggressiveness. I use a valuable dataset that includes all the quotes behind the best quote that enables us to make a better analysis of what occurs behind the scenes. In this paper, liquidity variable (spread at the best quote, depth at the best quote and behind the best quote, absolute order flow, and absolute order imbalance) sensitivity is a proxy for private information shocks whereas financial analyst surprise at the ar-

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rival of public announcements in the market is a proxy for public information shocks. The former measure is derived from the Canadian FX market dataset and the latter from Bloomberg. After initial evaluation of volatility before and after the arrival of macroeconomic announcements, I set up my regression model to examine the sensitivity of realised variance to my explanatory variables. My results indicate that private information explains realised volatility better than public information. Also, spread, as a proxy of private information, reveals more significant results than the other variables. I break the pre-scheduled macroeconomic announcements into three samples: Full, Relatively Important, and Most Important samples. My results show that the first two samples provide more significant results than the last sample and this might be due to extensive public information about the Most Important announcements. As robustness checks, I test the impact of news at times other than 8:30, the effect of all liquidity measures jointly, the impact of macroeconomic announcements in US and Canada separately, a detailed analysis on surprise, an economic classification of announcements and rerun the model, and the magnitude change of liquidity measures in smaller intervals of three 5-minute intervals before and after the 8:30 announcement, between news days and no-news days. Overall, my results show that spread at the best quote and absolute order flow are the main liquidity measures with more significant results, confirming spreads importance.

The thesis proceeds with risk and uncertainty evaluation using the real options frameworks. The third Chapter constructs a real options switching problem under uncertainty and reversibility (the ability to go back to an original product) to optimise the switching between two competing products. The assumptions are: (1) products have an infinite lifetime and (2) there is only one switching strategy (removing the existing product and initiating the other product). In a case of options to switch, investors evaluate a cost-benefit analysis that normally does not follow the Marshallian investment trigger and generates hysteresis. I argue that according to the

type of activity, the time cost of switching may vary. For example, in livestock farming (e.g. beef and sheep), there is relatively little time cost and the switching process is fast, while with seasonal crops (e.g. beetroot and carrots) there is a time cost of one year, and in horticulture (e.g. apples and grapes) the time cost increases to 5 or 6 years, which is a critical time cost. I set up three models according to each case: a cash-cost model, a cash-time model, and a projection model. I expect that the three models would give qualitatively different outputs. The cash-cost model ignores the time cost while the cash-time model formally measures the time costs and the projection model approximates the time costs using the expected monetary cost of time. I then consider a range of scenarios. In each scenario, I examine the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place, across the three models and compare with the baseline model. I bootstrap samples and test if my statistics across three models are quantitatively different. Results show that the cash-cost model that is used throughout the literature overestimates the switch frequency. The projection model presents better results compared with cash-cost, but underperforms compared with cash-time due to approximating the time cost, which is very critical in horticulture. The cash-time model computes the exact time cost and its results are significantly different from the projection model when the time of waiting is substantial. The sensitivity results from the cashtime model are significantly different from the projection model at the 1% level in all scenarios except in the case when volatility is high. Results indicate that increases in cost, time, volatility, drift, and discount rate have negative impacts on the switch frequency. If the correlation between two crops is positive, it has negative impacts on switch frequency, otherwise it has positive impacts.

In Chapter four, I evaluate orchards in the Hawkes Bay region and provide a statistical review of the dataset in terms of land allocation to apple varieties and how this changes over time. I use the measured pa-

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rameters and orchardists actions in the following Chapter.

In Chapter five, I extend my model from the third Chapter and assume: (1) apple trees have a finite productive lifetime similar to what is seen in reality and (2) the model is more complex by having two competing switch strategies: the uproot-plant and the graft strategy. This Chapter not only constructs a complex dynamic real options problem but also compares orchardists choice with the optimal policy that is proposed by the model. In this regard, I examine the behaviour of orchardists in the Hawkes Bay region orchards. In my model, orchardists have the option to plant Braeburn or Queen apples, and they have the incentive to switch from one type of apple to another as revenues change. There are two strategies to switch: the uproot-plant and the graft strategies. The former is a relatively expensive and slow process but results in a rejuvenated orchard, whereas the latter is less expensive and faster but continues with existing apple trees. Hence, I keep track of apple tree age and optimise the land use value under uncertainty and reversibility. Results comprise land use value when keeping the existing apple trees, switching to the competing apple type by the uproot-plant strategy, and switching by grafting. The optimal land value is the maximum of these three values for each apple type and at each node of each age. In results, the optimal land use value (what my model proposes) is compared with what orchardists actually choose. I employ a conditional logit model to examine the difference between the optimal and actual behaviours.

There are situations in my dataset where orchardists' actions do not match the optimal land use policy proposed by my real options model. According to the current model set up, on average an orchardist with a hectare land area loses 0.5% of the corresponding net revenue of land value that by assuming the same cost of \$50 for both apple types, varies from \$220.93 to \$382.43 for Braeburn and from \$272.98 to \$693.77 for Queen. Orchardists have a tendency towards keeping the existing trees and do not

exercise their options to switch to the competing apple type either through the uproot-plant or the graft strategy. This hysteresis might govern their behaviour as long as there is not any force or a strong incentive to convince them to switch. Any financing or incentive from the government may lead orchardists to switch at the right time and decrease their inertia. My results from the conditional logit model show that orchardists are biased towards inertia. However, these results assume that orchardists do the right thing in the next period of time and theses costs would be small unless they are really in the money. The differences between the optimal land value based on the real options analysis and the actual value that orchardists achieve through their current policy are a proxy for how far they are from the policy advised by my model. However, the model results are based on some assumptions and changing the parameters will minimise the difference between optimal and actual. We test this assertion by changing the feeding parameters in sensitivity analysis section and find lower difference. This shows that orchardists might use different parameters than what we have used in my model. Also, a range of simplifications such as choosing only two apple types and ignoring the fact that orchardists have a portfolio of apple varieties might explain the hysteresis among orchardists that results in difference between theory and practice. Comparing small orchardists with large orchardists reveals the same attitude of being less willing to exercise switching options.

Chapter 2

Public and Private Information Shocks to Realised Variance in the US/Canadian FX Market

2.1 Introduction

Information asymmetry is one of the main drivers of trading in any market. Information asymmetry in its turn causes adverse selection and price changes. Hence, information shocks and volatility are strongly linked. Information shocks fall into two main categories: public information shocks and private information shocks. Public information shocks are typically measured as financial analyst surprise at the arrival of public announcements in the market¹ (see Balduzzi, Elton, and Green (2001), Pasquariello and Vega (2007), and Ben-Omrane and Hafner (2015)). On the other hand, private information shocks are a more subtle concept and not so easy to observe or capture. Private information shocks refer to any information that is not publicly available for all market participants. This can come from either direct access to information, or indirect access, such as interpretation capability or forecasting skill of agents in the market.

This chapter examines the limit order book in the Canadian dollar/US dollar currency market. My main contribution is that I use the limit order book to examine asymmetric information impact on liquidity that results in price fluctuations (volatility) in the FX market. I capture the impact of belief dispersion and private information shocks through liquidity measures (namely: spread at the best quote, depth at the best quote, depth at the second best quote, absolute order flow, and absolute order imbalance). At the same time, I use macroeconomic announcements (see Tables 2.1 and 2.2) to measure public information shocks. These two measures allow me to decompose the source of volatility shocks into public and private components.

¹ Analyst recommendations (ANR) display a list of analysts' recommendations, price targets, price target time periods, and a consensus rating for a selected equity. The consensus rating is based on analyst recommendations and compiled by Bloomberg reporters and researchers around the world. ANR allows you to determine the latest analyst sentiment and identify which analysts have the most accurate equity ratings (Bloomberg Help).

I conjecture that macroeconomic announcements not only include public information shocks, but also that the interpretation of these data lead to private information shocks. The limit order book is a valuable information source that records the behaviour of participants in the market. I expect that the sensitivity to liquidity measures, as a channel of private information, rises with the arrival of macroeconomic announcements. However, this sensitivity might vary according to the level of importance and how extensively the announcement has been analysed by traders in the market.

Most studies in the market microstructure literature are not specifically based on the limit order book; rather they are extensions of the Kyle (1985) model, which is based on a market maker. However, electronic limit order markets have outperformed dealer markets in many ways. For example, the access to instantaneous execution without the intervention of a market maker. Having said that, the information value of the limit order book and, in particular, the market activity around the arrival of information, remains an open question.

A range of studies shed light on this question and provide some answers (see Foucault, Moinas, and Theissen (2007), Lo and Sapp (2010, 2011), Jiang, Lo, and Valente (2014), Hoffmann (2014), Foucault, Hombert, and Rosu (2016)). These studies conclude that the limit order book contains asymmetric information because limit order submitters can potentially be informed traders. When informed traders submit orders, measures of depth and spread will reflect both private information and public information.

While limit order books are an established part of many markets, the presence of a limit order book for an FX market is particularly interesting. In a traditional FX market, there is no limit order book (i.e. it operates as a dealer market) and order flow serves as a proxy for private information. However, starting in 2000, an electronic limit order book was set

up by Reuters and Electronic Broking Services (EBS), allowing FX market participants to trade with each other. Traders post orders as either limit orders or market orders, and EBS executes the orders (bids against offers) according to price and time priority. Having the opportunity to observe bids/asks behind the best quotes would enable a trader to gauge the degree of asymmetric information present in the market, and evaluate the risk of market order submission. However, quotes behind the best quoted prices are not observable, and hence they are entirely private information. What is special about my dataset (Reuters D-3000) is that I have access to the dealers' full data including the hidden information (such as behind the best quote) and information on changes in orders (such as adding, deleting, or modifying quotes) in the limit order book. Further, since this is hidden, it could not be taken into account by traders, and hence would not prejudice their market order submission. I contribute to the literature by capturing private information that is carried by limit orders in addition to order flow as proposed in the FX literature. This chapter seeks to address the question of whether private information derived from the limit orders changes the liquidity and moves volatility. The empirical literature shows that public information arrival has a significant impact on volatility (see Bauwens, Ben-Omrane, and Giot (2005), Jiang, Lo, and Valente (2014), and Ben-Omrane and Hafner (2015)). However, this literature is still silent on whether private information, especially that coming from limit orders, affects volatility.

Recent exchange rate studies, instead of relying on macroeconomic determinants, have used the market microstructure approach as a data driven technique to link exchange rate changes to the limit order book as the channel of transferring information (see Evans and Lyons (2002), Lo and Sapp (2011), Lovcha and Perez-Laborda (2013), and Ben-Omrane and Hafner (2015)). However, the limit order book contains more information that can be extracted and used to explain and forecast future price changes due to having access to all the quotes. Lo and Sapp (2011) use

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the electronic order book to evaluate the determinants of order submission strategies (i.e. the choice of limit orders vs market orders) for traders in the Canadian dollar/US dollar currency pair. The aggressiveness of orders and the role of timing in the limit order market is investigated in two other papers (Lo and Sapp, 2008, 2010). I use the same electronic order book (Reuters D-3000) but answer a different question: I assert that volatility can be exacerbated by order aggressiveness. Foucault, Moinas, and Theissen (2007) state that the size of bid-ask spread, which is one of the liquidity measures derived from the limit order book, contains volatility information and can be representative of private information. They also propose that the limit order book includes information about the magnitude of future price fluctuations. Therefore, I incorporate not only the order flow, which has been the main focus of past studies, but also other liquidity measures from the limit order book to explain private information.

The remainder of the chapter is ordered as follows: In Section 2.2, I evaluate the data and descriptive statistics. I examine the Canadian FX data, macroeconomic announcements data, and how to construct the dependent and explanatory variables. In Section 2.3, I set up my main model and, by running the corresponding regression model, and analyse the relationship between realised variance and explanatory variables. In Section 2.4, I present robustness tests to evaluate the primary results of my regression model, namely: realised variance without lag, realised variance in the US and Canadian market separately, surprise variable analysis, alternative measures of volatility, the effect of all liquidity variables jointly, the effect of news at times other than 8:30, the magnitude of liquidity variable changes, and realised variance in economic samples. Finally, in Section 2.5, I provide a summary of significant results and an overall conclusion.

2.2 Data and Descriptive Statistics

2.2.1 The Canadian FX Data

This chapter uses high frequency data for the Canadian FX market, which comes from the Reuters D-3000 proprietary inter-dealer platform. The sample span is from January 4, 2005 to December 30, 2005. The data include information about the electronic order book, both market orders and limit orders, recorded continuously, 24 hours a day. The dataset contains 3 085 363 observations of transactions in the Canadian FX market. A majority of macroeconomic announcements in this time period occur at 8:30. Following Ederington and Lee (1995), Clare and Courtenay (2001), Sager and Taylor (2011), and Neely and Dey (2010), I focus on a 15-minute interval, because the impact of new information remains much more than normal for approximately 15-minutes after a macroeconomic announcement.

The trades in the dataset are sorted based on the date and the time of entering into the order book or being removed from the order book. The size of trade is recorded, along with an identification code showing whether it is a bid or ask; how much quantity is entered or removed from the order book; and whether it is a limit order or market order. I order the price and quantity of trade in different columns from the best to the worst, in which the "best" bid is the highest price to buy and the "best" ask is the lowest price to sell. Each price has a related quantity. My ordering enables me to evaluate the market condition (e.g. the size, volume, or frequency of trades) behind the best quote.

I use this dataset to measure realised variance (as the dependent variable) and explanatory variables such as spread at the best quote, depth at the best quote, depth at the second best quote, absolute order flow, and absolute order imbalance. The method of measuring these variables is

explained in Section 2.2.3. I exclude holidays and weekends from the dataset, as there is insignificant trade, and no announcements occur on these days. Comparing the Canadian FX dataset and the pre-scheduled macroeconomic announcements determines news days and no-news days.

I evaluate the market condition using different variables to find the behaviour of the market on news days and no-news days. I exploit the data to measure different factors including trade frequency, limit order frequency, trade volume, trade size, quote size, depth of limit order, and bid-ask spread. Moreover, I not only separate my examination into news days and no-news days, but also before and after the announcement. The results in Section 2.3.1 show that the behaviour of the Canadian FX market on news days is significantly different from no-news days. Similarly, the pattern of these variables is significantly different after the announcement compared to before the announcement.

2.2.2 Macroeconomic Announcement Data

Macroeconomic announcements intensify volatility in financial markets and, in particular, foreign exchange markets due to creating asymmetric information among traders. There will be a range of different expectations from the arrival of news in the market. Some news might confirm the prior expectations of some traders whereas it might reject the expectation of other traders. Macroeconomic announcements have different impacts on the price movement and the reaction of agents in exchange markets (Fleming and Remolona, 1999).

The pre-scheduled macroeconomic announcement data is the second dataset I use, along with my main trading dataset. The trade in the Canadian FX market is affected by macroeconomic announcements in both Canada and the US. Therefore, the entire set of macroeconomic announcements by both countries are collected from Bloomberg. The macroeconomic an-

nouncements vary according to their timing and the effect they have on trading in the market. This dataset helps me to analyse the Canadian FX market in three ways. Firstly, it helps to differentiate news days and nonews days. Secondly, it provides a proxy for surprise to evaluate the impact of public information shocks on volatility. Finally, it enables me to extend my analysis to examine the effect of announcements on volatility according to their type and importance.

The macroeconomic announcements data have different variables including the type of announcement, the date and time of the news release, actual values in the announcements, survey values derived from the expectation of analysts, and the relevance values that indicate the importance of news.¹ The date and time of macroeconomic announcements is a proxy to separate trading days into news days and no-news days as well as examining the effect of the macroeconomic announcements at a specific time, for instance at 8:30, which is the main interest of this chapter.

The difference between the actual values in the announcements and analyst forecasts measures the amount of surprise in the market.² The literature shows that as the amount of surprise increases, volatility surges (see Andersen, Bollerslev, Diebold, and Vega (2003, 2007), Evans and Lyons (2005), Neely and Dey (2010), King, Osler, and Rime (2013), Ben-Omrane and Hafner (2015), Gomes and Peraita (2016), and Boudt, Neely, Secru, and Wauters (2017)). I also clean the macroeconomic announcements data by deleting announcements with the same nature or if their corresponding forecast value, realised value, or announcement time is not reported.

¹ Bloomberg has a search engine to aggregate and generate breaking news and through different feeding sources including web and social media rank the announcements and release it by relevant value indicator.

² Analyst recommendations (ANR) display a list of analysts' recommendations, price targets, price target time periods, and a consensus rating for a selected equity. The consensus rating is based on analyst recommendations and compiled by Bloomberg reporters and researchers around the world. ANR allows you to determine the latest analyst sentiment and identify which analysts have the most accurate equity ratings (Bloomberg Help).

After cleaning the announcement data, there are 37 different types of US macroeconomic announcements, along with 16 different types of Canadian macroeconomic announcements during the estimation period.

The main focus of this chapter is to study the impact of the 8:30 announcements on the Canadian FX market. The impact of announcements at other times is considered in the robustness section. Moreover, I control for the effect of news that overlaps with other announcements by removing contaminated announcements (e.g. I exclude the 9:00 Bank of Canada rate decisions and 9:15 US Industrial Production month-on-month report, as their windows overlap). Hence I test the pure impact of the given announcement, which is not contaminated by the impact of other announcements.

Tables 2.1 and 2.2 report the summary statistics of pre-scheduled macroeconomic announcements in the US and Canada. The first column reports all the pre-scheduled macroeconomic announcements in the US and Canada. The following columns present the related attributes of news, including the source of news, the number of observations of each news item, the time of releases, and the summary statistics including the minimum, maximum, mean, and standard deviation of each indicator. Each event happens on a different date, but the time on a given date remains constant. I convert all times to Eastern Time so that the comparison and examination of trades is consistent.¹ Tables 2.1 and 2.2 report the timing of news announcements. The majority of pre-scheduled macroeconomic announcements arrive at 8:30, which is why this time is the focus of my study. As noted earlier, I broaden my analysis to all times in the robustness section. The pre-scheduled macroeconomic announcements are different in terms of the relevance value (the importance of news). In line with the literature (especially Pasquariello and Vega (2007)), I separate the

¹ The release time of news and the time of trade in the US and Canada are reported according to different time zones (GMT, which is Greenwich Mean Time, for Canada, and EST, which is Eastern Standard Time for US).

TABLE 2.1: Summary Statistics of Pre-scheduled Macroeconomic Announcements Survey (US).

This table provides summary statistics for macroeconomic announcements in the US and Canada in the year 2005. The columns present, in order: the type of macroeconomic announcement in the US, the source of announcement releases (the abbreviation of each source is defined below), the number of observations of each announcement, the time of the announcement, minimum, maximum, mean, and standard deviation of observations in each announcement.

Abbreviations: FR: Federal Reserve, BLS: Bureau of Labour Statistics, BEA: Bureau of Economic Analysis, CB: Census Bureau, DL: Department of Labour, B: Bloomberg, T: Treasury, CBB: Chicago Business Barometer, NAHB: National Association of Home Builders, K: Thousands, B: Billion.

Event	Source	Obs.	Time	Min	Max	Mean	Std.
US Announcements							
Bloomberg Consumer Confidence	В	10	17:00	-22	-8	-14.5	4.7668
Building Permits	CB	12	08:30	1985	2160K	2077.9K	49.8K
Business Inventories	CB	12	08:30	0.0010	0.0090	0.0038	0.0024
CPI MoM	BLS	12	08:30	-0.0040	0.0090	0.0025	0.0033
Capacity Utilization	FR	12	09:15	0.7890	0.8030	0.7953	0.0035
Change in Nonfarm Payrolls	BLS	12	08:30	-150K	225K	159.3K	101.1K
Chicago Purchasing Manager	CBB	12	10:00	52	62.500	58.617	3.289
Conf. Board Consumer Confidence	CB	12	10:00	88	106.30	99.050	5.706
Construction Spending MoM	CB	12	10:00	0.0030	0.0070	0.0049	0.0011
Consumer Credit	FR	12	15:00	4.1B	10B	6.4B	1.6B
Continuing Claims	DL	50	08:30	2573K	2939K	2664.7K	93.2K
Current Account Balance	BEA	4	08:30	-205B	-183B	-192.75B	9.2B
Durable Goods Orders	CB	12	08:30	-0.0150	0.0150	0.0034	0.0111
Empire Manufacturing	FR	12	08:30	1	25	15.983	6.1573
Employment Cost Index	BLS	4	08:30	0.0080	0.0100	0.0085	0.0010
FOMC Rate Decision (Upper Bound)	FR	8	14:15	0.0250	0.0425	0.0338	0.0061
Factory Orders	CB	12	10:00	-0.0230	0.0300	0.0058	0.0153
GDP Annualised QoQ	BEA	12	08:30	0.0330	0.0430	0.0368	0.0029
Housing Starts	CB	12	08:30	1903K	2090K	2013.6K	56.2K
ISM Manufacturing	ISM	12	10:00	51.400	58.500	55.383	2.4778
Import Price Index MoM	BLS	12	08:30	-0.0050	0.0140	0.0047	0.0068
Industrial Production MoM	FR	12	09:15	-0.004	0.01	0.0034	0.0031
Initial Jobless Claims	DL	52	08:30	309K	450K	329.8K	24.5K
Leading Index	CB	12	10:00	-0.0050	0.0080	0.0003	0.0040
Monthly Budget Statement	T	12	14:00	-100B	60B	-26.9B	50.8B
NAHB Housing Market Index	NAHB	12	13:00	61	71	67.833	2.9797
New Home Sales	CB	12	10:00	1125K	1350K	1253.4K	77.5K
Nonfarm Productivity	BLS	8	08:30	0.0150	0.0450	0.0238	0.0100
PCE Core YoY	BEA	11	08:30	0.0150	0.0190	0.0174	0.0016
Personal Consumption	BEA	10	08:30	0.0300	0.0430	0.0359	0.0042
Personal Income	BEA	12	08:30	-0.0260	0.0340	0.0041	0.0128
Personal Spending	BEA	12	08:30	-0.0020	0.0100	0.0045	0.0036
Philadelphia Fed Business Outlook	PFR	12	12:00	10	25	14.733	4.6582
Retail Sales Advance MoM	CB	12	08:30	-0.0140	0.0210	0.0038	0.0093
Trade Balance	СВ	12	09:30	-62.9B	-54B	-58.7B	2.6B
Unemployment Rate	BLS	12	08:30	0.0500	0.0540	0.0516	0.0015
Wholesale Inventories MoM	СВ	12	10:00	0.0030	0.0090	0.0056	0.0017

TABLE 2.2: Summary Statistics of Pre-scheduled Macroeconomic Announcements Survey (Canada).

This table provides summary statistics for macroeconomic announcements in the US and Canada in the year 2005. The columns present, in order: the type of macroeconomic announcement in Canada, the source of announcement releases (the abbreviation of each source is defined below), the number of observations of each announcement, the time of the announcement, minimum, maximum, mean, and standard deviation of observations in each announcement.

Abbreviations: BC: Bank of Canada, STCA: Statistics Canada, K: Thousands, B: Billion.

Event	Source	Obs.	Time	Min	Max	Mean	Std.
Canadian Announcements							
Bank of Canada Rate Decision	BC	8	09:00	0.0250	0.0325	0.0269	0.0029
Building Permits MoM	STCA	12	08:30	-0.0500	0.0400	-0.0082	0.0270
CPI YoY	STCA	12	07:00	0.0180	0.0330	0.0233	0.0046
Capacity Utilization Rate	STCA	4	08:30	0.8560	0.8720	0.8625	0.0072
Current Account Balance	STCA	4	08:30	3.4B	9B	6.14B	2.51B
Housing Starts	STCA	12	08:30	215K	230K	224.5K	4.48K
Industrial Product Price MoM	STCA	12	08:30	-0.0100	0.0090	0.0030	0.0047
Int'l Merchandise Trade	STCA	12	08:30	4.3B	6.9B	5.19B	0.787B
Labor Productivity QoQ	STCA	4	08:30	0.0010	0.0050	0.0030	0.0016
Manufacturing Sales MoM	STCA	12	08:30	-0.0060	0.0100	0.0038	0.0046
New Housing Price Index MoM	STCA	12	08:30	0.0030	0.0050	0.0039	0.0007
Quarterly GDP Annualised	STCA	4	08:30	0.0170	0.0360	0.0263	0.0078
Raw Materials Price Index MoM	STCA	12	08:30	-0.0200	0.0400	0.0148	0.0201
Retail Sales MoM	STCA	12	08:30	-0.0050	0.0140	0.0031	0.0065
Unemployment Rate	STCA	12	07:00	0.0660	0.0730	0.0687	0.0019
Wholesale Trade Sales MoM	STCA	12	08:30	0.0020	0.0070	0.0050	0.0015

announcements into three samples, based on the relevance value that is reported by Bloomberg. The first sample includes the Full set of announcements in the US and Canada. The second sample restricts attention to the Relatively Important announcements: change in non-farm payrolls, initial jobless claims, FOMC rate decisions, annualised GDP Quarter-on-Quarter (QoQ), ISM manufacturing, CPI Month-on-Month (MoM), the conference board consumer confidence, durable goods orders, retail sales advance MoM, new home sales, housing starts, industrial production MoM, unemployment rate, factory orders, personal income, and personal spending, in the US. For Canada, the relatively important announcements are bank of Canada rate decisions, CPI Year-on-Year (YoY), annualised quarterly GDP, and the unemployment rate in Canada. Finally, the third sample restricts the announcements even further to consider only the Most Important announcements: change in non-farm payrolls, initial jobless claims, FOMC rate decisions (upper bound), annualised GDP QoQ, ISM manufacturing, and CPI MoM in the US as well as bank of Canada rate decisions and CPI YoY in Canada. In the robustness section, I separate the announcements into four samples based on their economic function: output, prices, employment, and rate decision.

2.2.3 Constructing Variables

In this section, I describe the construction of variables for my study. The variables of interest include realised variance as the dependent variable, along with a number of explanatory variables.

Realised Variance

Conventional GARCH models (as a parametric measure of volatility) are unable to explain jumps in volatility, and the low autocorrelation in squared

returns when evaluated out of sample (Kliber and Bedowska-Sojka, 2010). However, I can use the aggregated high frequency returns, which are an observable estimate of latent volatility and a good estimate from which to forecast future volatility. The dependent variable in my model is realised variance (*RV*). I examine *Realised Variance* in the 15-minute interval before and after the 8:30. I choose the 8:30 announcement instead of other announcements because a large portion of important news will come into the market at this time, and by focusing on announcements at the same time of day, it is robust to intraday patterns in volatility responses. However, in Section 2.4.6, I analyse the impact of news at times other than 8:30. I apply the Hansen and Lunde (2006) method to measure the realised variance, which is the sum of squares of the difference in log mid-quote prices in each one-minute interval during a 15-minute interval:

$$RV_t = \sum_{\tau=1}^{15} r_{t-\tau}^2 = \sum_{\tau=1}^{15} (\log(mq_{t-\tau}) - \log(mq_{t-\tau-1}))^2 * 1000,$$
 (2.1)

where RV_t is the realised variance during each 15-minute interval, r_{τ} is the return over each 1-minute interval, mq_{τ} is the mid-quote (average of the best bid and best ask at the end of each one-minute interval), τ is the counter of each one-minute interval, and 1000 is a rescaling for comparability with other variables. If there is an interval without any reported quote, I use the quote from the previous interval. Choosing the one-minute interval is ad hoc and changing it does not affect my results.

In traditional realised variance calculations, the limit order book is not available, and the returns used to calculate realised variance are based on transaction prices. However, in this study, I do not need to rely on transaction prices since I have access to the limit order book with detailed information. In this regard, I model realised variance as an ex-post and observable volatility in the FX market using the limit order book.

In the robustness section, I apply two alternative measures to see

whether the results change. The first alternative to realised variance is absolute return. Many scholars have used the absolute return as a proxy for volatility in the literature (see Jones, Kaul, and Lipson (1994), Andersen and Bollerslev (1997, 1998), Granger and Sin (2000), Ahn, Bae, and K. Chan (2001), Forsberg and Ghysels (2007), Foucault, Moinas, and Theissen (2007)). I define absolute return as:

$$|R_t| = |\log(mq_{t-1}) - \log(mq_{t-16})| * 1000,$$
 (2.2)

where $|R_t|$ is the absolute return for each 15-minute interval, and mq_t is the average of the best bid and the best ask at time t.¹

In high frequency data, there is a possibility that mid-quotes differ from their true value due to market microstructure effects (Hasbrouck, 1993). Therefore, I propose a third measure that takes into consideration the possibility of negative correlation of returns inducing predictability in future returns. The third measure of volatility is the modified absolute return. I compute the 15-minute interval absolute returns during the active time of the trading day (6:00 a.m – 5:00 p.m). Then I regress the AR values on their lagged values and retain the absolute values of residuals of this regression as my third proxy for volatility. This measure is used by Forsberg and Ghysels (2007) and Foucault, Moinas, and Theissen (2007). I define the third measure R^{resi} as:

$$|R_t| = \beta_0 + \beta_1 |R_{t-15}| + \epsilon_t$$

 $R_t^{resi} = |\epsilon_t| * 1000,$ (2.3)

where $|R_t|$ is the absolute return for each 15-minute interval, $|R_{t-15}|$ is the 15-minute lag absolute return, ϵ_t is the regression residual, and R_t^{resi} is the time series of modified absolute returns.

¹ In all my measures, I use the quote from the preceding time period because the effect of each quote will be revealed in the subsequent time period.

Explanatory Variables

A key explanatory variable in my model set up is the existence of news in the market. I define a dummy variable, which is equal to one if there is news and equal to zero if there is not any news. This measure examines the presence of news and its effect on the market. Another explanatory variable is the amount of surprise or shock in the market. This can be calculated by finding the difference between the realised value and analyst forecast, divided by the standard deviation of each sample. As a result, in this process, all the macroeconomic variables with different measurement units will be standardised, since different units of measurement will shift both means and standard deviations.

$$Sur_{k,t} = \left| \frac{Actual_{k,t} - Survey_{k,t}}{S.d(Actual_k - Survey_k)} \right|, \tag{2.4}$$

where $Actual_{k,t}$ is the realised value of announcement k at time t, $Survey_{k,t}$ is the analyst forecasts of announcement k at time t, and S.d is the standard deviation of the difference in the numerator. Many scholars have exploited this method to measure public information shocks in financial markets (see Balduzzi, Elton, and Green (2001), Pasquariello and Vega (2007), and Ben-Omrane and Hafner (2015)). I will examine the surprise with the positive and negative sign as well as zero surprise in the robustness section (in the literature there are studies that evaluate the sign of the surprise: see Gosnell, Keown, and Pinkerton (1996), and Andersen, Bollerslev, Diebold, and Vega (2003)).

Among the other explanatory variables, the bid-ask spread variable plays an important role in the literature and many scholars use this proxy to explain market conditions. Brekke and Oksendal (1988) consider the bid-ask spread as a signal of liquidity and transaction costs in the market. They argue that the inventory management cost for market makers decreases when many uninformed traders are present in the market. This

situation results in a lower level of adverse selection, and consequently a narrower bid-ask spread. In my analyses, the spread is an explanatory variable, being measured as the difference between the best ask quote and the best bid quote at the end of each 15-minute interval. It is defined as:

$$S_t = ap_{0,t-1} - bp_{0,t-1}, (2.5)$$

where $ap_{0,t-1}$ is the best ask quote one minute before t, and $bp_{0,t-1}$ is the best bid quote one minute before t, both taken from the electronic order book. I use this variable as a liquidity measure to explain realised variance in the market. Pasquariello and Vega (2007) use the change in bid-ask spread to examine if anonymity matters in an electronic limit order book. However, I use this measure to capture the sensitivity of realised variance, which counts as a private information shock.

The explanatory variables $D_{0,t}$ and $D_{1,t}$ are measured as the summation of quantity quoted at the best quote and the second best quote at the end of each 15-minute interval respectively:

$$D_{0,t} = aq_{0,t-1} + bq_{0,t-1}$$

$$D_{1,t} = aq_{1,t-1} + bq_{1,t-1},$$
(2.6)

where $D_{0,t}$ is the depth at the best quote; $D_{1,t}$ is the depth at the second best quote; $aq_{0,t-1}$ is the quantity at the best ask price one minute before t; $bq_{0,t-1}$ is the quantity at the best bid price one minute before t; $aq_{1,t-1}$ is the quantity at the second best ask price one minute before t; and $bq_{1,t-1}$ is the quantity at the second best bid price one minute before t. Depth at the best quote is an active and aggressive quote and has a high likelihood of getting exercised. However, in a market with asymmetric information, traders hide their quotes behind the best quote to avoid this probability due to uncertainty (Jiang, Lo, and Valente, 2014). As a result, I expect that a decline in liquidity will have a negative impact on the depth at the

best quote and a positive impact on the depth behind the best quote. The uncertainty among traders that is caused by having different levels of information leads to an illiquid market when they quote further away from the best quotes. The primary assumption here is that traders are waiting for some information to be released by scheduled announcements. In this regard, there is an uncertain situation for traders but the level of uncertainty varies among traders. Some of the traders are more expert and have better understanding of the market whereas some of traders are not as experienced.

Recent studies suggest that order flow matters in determining exchange rates and in explaining fluctuations in the market. Evans and Lyons (2002) state that order flow is a conduit for aggregating heterogeneous expectations about the future state of the economy and different interpretations of news in real time. The order flow will be driven by a broad set of macroeconomic information, which is the concern of exchange rate theories (Rime, Sarno, and Sojli, 2010). In line with the literature, I include order flow as an explanatory variable in my model to explain realised variance in the market. The value of order flow is the difference between the buy volume and the sell volume during the 15-minute interval. I use the absolute value of order flow as:

$$|AOF_t| = \sum_{\tau=1}^{15} |Q_{t-\tau-1,t-\tau}^{Ask} - Q_{t-\tau-1,t-\tau}^{Bid}|, \qquad (2.7)$$

where $|AOF_t|$ is the absolute order flow of each 15-minute interval, $Q_{t-\tau-1,t-\tau}^{Ask}$ is the quantity quoted on the ask side between times $t-\tau-1$ and $t-\tau$, and $Q_{t-\tau-1,t-\tau}^{Bid}$ is the quantity quoted in the bid side between times $t-\tau-1$ and $t-\tau$.

Finally, the last explanatory variable is order imbalance, which is informative about price movements. Order imbalance has a significant effect on price and liquidity for two reasons. Firstly, in line with Kyle's theory

of price information, order imbalance is a proxy for private information that can decrease the liquidity in the market temporarily (and in an extreme case, permanently). Secondly, the inventory problem faced by the market maker, which is adjusting the bid-ask spreads according to inventory, is exacerbated due to extreme order imbalances (Chordia, Roll, and Subrahmanyam, 2002). I measure the order imbalance by finding the difference between the bid-depth and ask-depth at the end of each 15-minute interval.

$$|AOI_t| = \left| Depth_{t-1}^{Ask} - Depth_{t-1}^{Bid} \right|, \tag{2.8}$$

where $|AOI_t|$ is the absolute order imbalance of each 15-minute interval, $Depth_{t-1}^{Ask}$ is the ask depth at one minute before t, and $Depth_{t-1}^{Bid}$ is the bid depth at one minute before t. In the following section, I use all these measurements in a regression model to explain the realised variance in the Canadian FX market.

2.3 Empirical Analysis

In this section, I evaluate the impact of 8:30 macroeconomic announcements on realised variance. I evaluate *Realised Variance* changes on news days and no-news days as well as before and after the announcements. I set up my main regression model, and evaluate the relationship and sensitivity of *Realised Variance* to my explanatory variables.

2.3.1 The Model

I split each trading day into two 15-minute intervals. The first interval examines market conditions before 8:30, and the second interval examines market conditions after 8:30. This process is applied for both news days and no-news days.

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TABLE 2.3: Difference in Difference t-test Results for News Days vs No-news Days and before vs after Announcements.

This table presents the mean of realised variance before and after 8:30, news days and no-news days, and in three samples: full sample, relatively important, and the most important. The t-test evaluates whether there is a significant difference between news days and no-news days as well as before and after announcements separately. The number of observations for Full sample, relatively important (RI) sample, and most important (MI) sample are 330, 240, and 176 respectively. I use UNIVARIATE procedure in SAS to test if my samples before and after 8:30 are random with normal distribution. Results show that all three samples before and after 8:30 have normal distribution.

Note: The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively.

Realised Variance	Before	After	After-Before
Panel A: Full Sample News (Full)			
News	0.1103	0.2920	0.1817*** (0.0292)
No News	0.0934	0.1031	0.0097 (0.3850)
Diff(News and No-News)	$0.0169 \\ (0.0153)$	0.1889*** (0.0498)	0.1720*** (0.0518)
Panel B: Relatively Important News (RI)			
News	0.1107	0.2807	0.1700^{***} (0.0321)
No News	0.0973	0.1031	0.0058 (0.6010)
Diff(News and No-News)	0.0134 (0.0161)	0.1776^{***} (0.0442)	0.1642*** (0.0467)
Panel C: Most Important News (MI)			
News	0.0883	0.4590	0.3707*** (0.0891)
No News	0.0973	0.1031	0.0058 (0.6010)
Diff(News and No-News)	-0.0090 (0.0189)	0.3559^{***} (0.0661)	0.3649*** (0.0676)

Table 2.3 presents the results. I test the hypothesis of no effect from news for all trading days based on three different samples of announcements. The first sample includes the *Full* set of pre-scheduled macroeconomic announcements; the second sample includes the *Relatively Important* pre-scheduled macroeconomic announcements; and finally, the last sample includes the *Most Important* pre-scheduled macroeconomic announcements. I test *Realised Variance* changes between news days and no-news days, as well as before and after the announcements on news days to see if they are significant or not.

Table 2.3 shows that after 8:30, there is a significant difference in realised variance. This conclusion holds for all three samples of announcements. The difference is much higher for the *Most Important* news sample compared with the other two samples. In addition, the difference between news days and no-news days is insignificant before 8:30, while this difference is higher and significant after 8:30. The difference in difference between news days and no-news days is significant for all three samples. My result for the *Most Important* sample is much higher compared with the other two samples. These volatility increases, however, could be due to either public information shocks or private information shocks.

2.3.2 Regression Analysis

In this section, my goal is to evaluate the impact of public information shocks and private information shocks by setting up a regression analysis. Before setting up the regression model, I check the sample data before and after 8:30 to see if they are stationary. Table 2.4 shows the unit root test for realised variance before and after 8:30. Results imply that there is no evidence of a unit root and both samples before and after the 8:30 are stationary and they do not have a trend as well.

I aim to break down the total effect of the arrival of macroeconomic

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Table 2.4: Realised Variance Unit Root Test Before and After 8:30

This table presents unit root tests to find if the samples before and after 8:30 are stationary. There are three different types of test to check the stationary. Results show that the samples are stationary without any trend and need of having difference. Note that the ADF test in SAS presents three tests: Rho, Tau, and F tests.

Augmented Dickey-Fuller Unit Root Tests								
	Before							
Туре	Lags	Rho	Pr < Rho	Tau	Pr < Tau	F	Pr > F	
	0	-58.49	<.00	-6.06	<.00			
Zero Mean	1	-24.66	0.00	-3.54	0.00			
	2	-17.54	0.00	-2.95	0.00			
	0	-152.51	0.00	-12.22	<.00	74.64	0.00	
Single Mean	1	-135.04	0.00	-8.14	<.000	33.16	0.00	
J	2	-185.11	0.00	-7.54	<.00	28.44	0.00	
	0	-153.16	0.00	-12.22	<.00	74.66	0.00	
Trend	1	-136.45	0.00	-8.16	<.00	33.28	0.00	
	2	-188.22	0.00	-7.55	<.00	28.50	0.00	
				Aft	er			
Туре	Lags	Rho	Pr < Rho	Tau	Pr < Tau	F	Pr > F	
	0	-98.08	<.00	-8.70	<.00			
Zero Mean	1	-44.65	<.00	-4.75	<.00			
	2	-30.36	<.00	-3.71	0.00			
	0	-164.18	0.00	-13.45	<.00	90.48	0.00	
Single Mean	1	-137.58	0.00	-8.28	<.00	34.28	0.00	
2	2	-158.37	0.00	-7.18	<.00	25.76	0.00	
	0	-166.14	0.00	-13.54	<.00	91.72	0.00	
Trend	1	-142.87	0.00	-8.39	<.00	35.21	0.00	
	2	-171.04	0.00	-7.32	<.00	26.81	0.00	

announcements into public and private by setting up a regression model. In this regard, I define dummy variables to captures the effect of the existence/absence of news and before/after the news. A group of liquidity variables (i.e. spread, depth at the best quote, depth at the second best quote, absolute order flow, and absolute order imbalance) capture the effect of private information shocks whereas surprise is the proxy for capturing the impact of public information shocks. The model is set up as:

$$RV_{t} = \beta_{0} + \beta_{1} \mathbf{1}_{News} + \beta_{2} \mathbf{1}_{Post} + \beta_{3} \mathbf{1}_{News} \mathbf{1}_{Post} + \theta_{1} RV_{t-15} + \theta_{2} RV_{t-15} \mathbf{1}_{News}$$

$$+ \theta_{3} RV_{t-15} \mathbf{1}_{Post} + \theta_{4} RV_{t-15} \mathbf{1}_{News} \mathbf{1}_{Post} + \delta_{1} X_{t} + \delta_{2} X_{t} \mathbf{1}_{News} + \delta_{3} X_{t} \mathbf{1}_{Post}$$

$$+ \delta_{4} X_{t} \mathbf{1}_{News} \mathbf{1}_{Post} + \gamma Sur_{t} \mathbf{1}_{Post} + \epsilon_{t},$$

$$(2.9)$$

where RV_t is the realised variance; $\mathbf{1}_{News}$ is the news dummy equal to 1 or 0 in the event of news and no-news respectively; $\mathbf{1}_{Post}$ indicates that I evaluate the impact of explanatory variables after 8:30; RV_{t-15} is the 15-minute lag realised variance; X_t denotes a liquidity measure, which could be spread at the best quote (S_0) , depth at the best quote (D_0) , depth at the second best quote (D_1) , absolute order flow (AOF), or absolute order imbalance (AOI), and captures the sensitivity of Realised Variance to liquidity changes. Surprise (Sur), as a proxy for public information shocks, captures the impact of realised variance after the arrival of prescheduled macroeconomic announcements. The interaction of the news dummy, $\mathbf{1}_{News}$, and post 8:30 dummy, $\mathbf{1}_{Post}$, represents the difference-indifference impact that is the main focus of my analysis. In other words, this explanatory variable evaluates the sensitivity difference between news days and no-news days along with before and after the announcements. If $\delta_4 > 0$, then sensitivity of volatility to the microstructure variable rises after news announcements, and the private information effect is present. Conversely, if $\gamma > 0$, then Realised Variance increases from surprises, consistent with public information shocks being present. Tables 2.5, 2.6, and 2.7 report my results of the regression model for each liquidity measure

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using the *Full*, *Relatively Important*, and *Most Important* samples. I examine the impact of all liquidity measures jointly in Section 2.4.5. Newey and West (1987) standard errors are applied to take care of autocorrelation and heteroskedasticity in the time series regression model.

Table 2.5 indicates that prior to the announcements, there are no significant differences between news days and no-news days for the Full sample. I capture this effect with the Dummy variable $\mathbf{1}_{News}$ and its interactions with RV_{t-15} and X_t . Similarly, there are no significant results after 8:30 on no-news days. I capture this effect with the Dummy variable $\mathbf{1}_{Post}$ and its interactions with RV_{t-15} and X_t . The sensitivity of realised variance to surprise as a proxy for public information shocks is almost zero ($\gamma \simeq 0$). However, spread sensitivity (as a proxy for private information) has a significant increase with the arrival of macroeconomic announcements, and this is consistent with the existence of private information. In other words, realised variance is highly sensitive to bid-ask spread, as a proxy for private information shocks, following the arrival of news. A high level of information asymmetry, due to the new information, increases the risk of trading in the market. Information asymmetry results in adverse selection, and, in turn, dealers place their quotes carefully due to lack of enough information. Informed traders send their quotes with enough confident whereas uninformed traders are not sure about the right price or volume quotes. This finding, in the first place, is consistent with the literature that pre-scheduled macroeconomic announcements have a significant impact on realised variance due to repositioning and repricing by the traders in the market (see Andersen and Bollerslev (1997), Fleming and Remolona (1999), Balduzzi, Elton, and Green (2001), and Evans and Lyons (2008)). However, my results also imply that private information shocks are the key driving force as opposed to public information shocks.

I evaluate the same regression model for two alternative samples, *Relatively Important* (Table 2.6) and *Most Important* (Table 2.7). The results

TABLE 2.5: Realised Variance at 8:30 Announcements Using the Full Sample.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 330 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liquidity Measure (Full Announcements Sample)						
variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI		
Intercept	0.0642*** (0.0181)	0.0775*** (0.0184)	0.0786*** (0.0176)	0.0687*** (0.0146)	0.0681*** (0.0136)		
1_{News}	0.0397 (0.0301)	0.0239 (0.0241)	0.0158 (0.0266)	0.0256 (0.0205)	0.0325 (0.0204)		
1_{Post}	-0.0317 (0.0272)	-0.0201 (0.0233)	-0.0253 (0.0250)	-0.0181 (0.0190)	-0.0134 (0.0187)		
$1_{News}1_{Post}$	0.0397 (0.0577)	0.1464^{**} (0.0729)	0.1110 (0.0743)	0.1235** (0.0614)	0.1236** (0.0646)		
RV_{t-15}	0.3821*** (0.0405)	0.3826*** (0.0395)	0.3810*** (0.0406)	0.3807^{***} (0.0422)	0.3847*** (0.0393)		
$\mathrm{RV}_{t-15}1_{News}$	-0.1217 (0.1655)	-0.1255 (0.1637)	-0.1211 (0.1651)	-0.1510 (0.1729)	-0.1353 (0.1630)		
$RV_{t-15}1_{Post}$	0.1443 (0.0985)	0.1695 (0.1108)	0.1452 (0.1008)	0.0953 (0.1676)	0.1745 (0.1118)		
$RV_{t-15} 1_{News} 1_{Post}$	0.1455 (0.3381)	0.2825 (0.3555)	0.1854 (0.3764)	0.2922 (0.4017)	0.2638 (0.3472)		
X_t	0.0095 (0.0429)	-0.0018 (0.0015)	-0.0017 (0.0011)	$0.0000 \\ (0.0002)$	-0.0006 (0.0016)		
$X_t 1_{News}$	-0.0464 (0.0704)	0.0005 (0.0021)	0.0016 (0.0020)	0.0001 (0.0002)	-0.0020 (0.0030)		
$X_t 1_{Post}$	0.0612 (0.0700)	0.0005 (0.0018)	0.0014 (0.0023)	0.0002 (0.0004)	-0.0015 (0.0022)		
$X_t 1_{News} 1_{Post}$	0.2337** (0.1081)	-0.0059 (0.0059)	0.0015 (0.0066)	-0.0004 (0.0005)	-0.0030 (0.0068)		
$\mathrm{Sur}_t 1_{Post}$	-0.0039 (0.0208)	0.0106 (0.0247)	0.0132 (0.0253)	0.0165 (0.0271)	$0.0108 \ (0.0251)$		
Adjusted R ²	0.2930	0.1630	0.1561	0.1527	0.1628		

are insignificant for both samples, *Relatively Important* and *Most Important*. This is consistent with the conjecture that there is extensive public analysis about the *Relatively Important* and *Most Important* announcements, which decreases the value of private analysis. To summarise, bid-ask spread sensitivity as a proxy for private information shocks shows a significant relationship with realised variance at the 5% level in the *Full* sample. However, the bid-ask spread coefficients for the *Relatively Important* and *Most Important* samples are insignificant. Liquidity variables are a sign of existing private information in the market that results in a more volatile market. Uninformed traders are exposed to financial losses in this market and they are better off to avoid trading by quoting further away from the best quote or exiting from the market.

2.4 Robustness and Extensions

In Section 2.3, I apply the Hansen and Lunde (2006) method to measure realised variance in the Canadian FX market. The results indicate that the arrival of new information has a significant impact on realised variance primarily due to private information shocks. Public information shocks have an insignificant impact. Spread is the key variable for measuring private information shocks. Other liquidity measures in the model do not produce significant results. In the following, I examine more tests to find if there is any change in my results. The robustness tests consist of: evaluating realised variance by removing the lag from the regression model, testing realised variance in the US and Canada separately, more tests on the surprise variable, alternative measures of volatility, the effect of all liquidity variables jointly, the effect of news at times other than at 8:30, the magnitude of liquidity variable changes, evaluation in economic samples, and finally a summary of all cases.

TABLE 2.6: Realised Variance at 8:30 Announcements Using the RI Sample.

Results of regression of dependent variable (RV) realised variance on liquidity measure: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 240 observations in RI announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liquidity Measure (RI Announcements Sample)						
	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI		
Intercept	0.0642*** (0.0183)	0.0775*** (0.0185)	0.0786*** (0.0178)	0.0687*** (0.0148)	0.0681*** (0.0137)		
1_{News}	0.0221 (0.0308)	0.0191 (0.0258)	0.0184 (0.0288)	0.0194 (0.0226)	0.0230 (0.0213)		
1_{Post}	-0.0317 (0.0274)	-0.0201 (0.0234)	-0.0253 (0.0252)	-0.0181 (0.0191)	-0.0134 (0.0189)		
$1_{News}1_{Post}$	0.0504 (0.0628)	0.1098 (0.0755)	0.0657 (0.0701)	0.1067 (0.0683)	0.1126 (0.0705)		
RV_{t-15}	0.3821*** (0.0408)	0.3826*** (0.0398)	0.3810*** (0.0409)	0.3807*** (0.0425)	0.3847*** (0.0396)		
$\mathrm{RV}_{t-15}1_{News}$	0.0028 (0.2098)	-0.0008 (0.2110)	-0.0080 (0.2156)	-0.0140 (0.2181)	-0.0031 (0.2105)		
$\mathrm{RV}_{t-15}1_{Post}$	0.1443 (0.0992)	0.1695 (0.1117)	0.1452 (0.1015)	0.0953 (0.1688)	0.1745 (0.1127)		
$\text{RV}_{t-15}1_{News}1_{Post}$	-0.1032 (0.4005)	-0.0755 (0.4285)	-0.2461 (0.4852)	0.0124 (0.4602)	-0.0339 (0.4075)		
X_t	0.0095 (0.0432)	-0.0018 (0.0015)	-0.0017 (0.0011)	$0.0000 \\ (0.0002)$	-0.0006 (0.0016)		
$X_t 1_{News}$	-0.0125 (0.0737)	-0.0002 (0.0024)	0.0001 (0.0019)	0.0000 (0.0003)	-0.0018 (0.0034)		
$X_t 1_{Post}$	0.0612 (0.0705)	0.0005 (0.0018)	0.0014 (0.0023)	0.0002 (0.0004)	-0.0015 (0.0022)		
$X_t 1_{News} 1_{Post}$	0.1396 (0.1015)	0.0010 (0.0062)	0.0116 (0.0117)	0.0001 (0.0009)	-0.0006 (0.0065)		
$\operatorname{Sur}_t 1_{Post}$	0.0185 (0.0225)	0.0278 (0.0313)	0.0257 (0.0313)	0.0239 (0.0354)	0.0287 (0.0313)		
Adjusted R ²	0.3359	0.1848	0.2090	0.1808	0.1901		

TABLE 2.7: Realised Variance at 8:30 Announcements Using the MI Sample.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 176 observations in MI announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liquidity Measure (MI Announcements Sample)						
V 111111111111111111111111111111111111	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI		
Intercept	0.0642*** (0.0188)	0.0775*** (0.0190)	0.0786*** (0.0183)	0.0687*** (0.0152)	0.0681*** (0.0141)		
1_{News}	0.0162 (0.0433)	-0.0006 (0.0294)	0.0003 (0.0396)	-0.0093 (0.0266)	0.0077 (0.0242)		
1_{Post}	-0.0317 (0.0282)	-0.0201 (0.0241)	-0.0253 (0.0259)	-0.0181 (0.0197)	-0.0134 (0.0194)		
$1_{News}1_{Post}$	0.1007 (0.1713)	0.1791 (0.1717)	0.0494 (0.1427)	0.1917 (0.1638)	0.2195 (0.1731)		
RV_{t-15}	0.3821*** (0.0420)	0.3826*** (0.0409)	0.3810*** (0.0421)	0.3807*** (0.0438)	0.3847*** (0.0408)		
$\mathrm{RV}_{t-15}1_{News}$	-0.0650 (0.2098)	-0.0423 (0.2234)	-0.0317 (0.2106)	-0.2452 (0.2259)	-0.0554 (0.2098)		
$\mathrm{RV}_{t-15}1_{Post}$	0.1443 (0.1021)	0.1695 (0.1149)	0.1452 (0.1045)	0.0953 (0.1739)	0.1745 (0.1160)		
$\text{RV}_{t-15}1_{News}1_{Post}$	1.7000 (1.0856)	1.7010 (1.8767)	3.3298** (1.4723)	2.8365** (1.5110)	2.5675* (1.5887)		
X_t	0.0095 (0.0444)	-0.0018 (0.0015)	-0.0017 (0.0012)	$0.0000 \\ (0.0002)$	-0.0006 (0.0017)		
$X_t 1_{News}$	-0.0429 (0.1193)	0.0008 (0.0022)	$0.0004 \\ (0.0041)$	0.0010 (0.0006)	-0.0010 (0.0032)		
$X_t 1_{Post}$	0.0612 (0.0726)	0.0005 (0.0018)	0.0014 (0.0023)	0.0002 (0.0004)	-0.0015 (0.0023)		
$X_t 1_{News} 1_{Post}$	0.1227 (0.1412)	0.0201 (0.0333)	0.230 (0.0154)	0.0004 (0.0041)	-0.0028 (0.0287)		
Sur_t1_{Post}	-0.0414 (0.1113)	-0.1309 (0.1172)	-0.2385 (0.1523)	-0.1417 (0.1312)	-0.1414 (0.1378)		
Adjusted R ²	0.4906	0.4139	0.4903	0.3984	0.3986		

2.4.1 Realised Variance Regression Without Lag

To clarify the impact of explanatory variables on realised variance, I evaluate the regression model for all three samples by excluding the realised variance lag. Table 2.8 shows that the results do not change after excluding realised variance lag and spread at the best quote remains the only significant variable among all liquidity variables as a proxy of private information shock. Surprise as a public information shock does not reveal any significant results even after excluding the realised variance lag. Tables 2.9 and 2.10 present the results for the relatively important and the most important news. There are not any significant results for these two samples since they are highly analysed by traders in the market.

2.4.2 Realised Variance in the US and Canadian Market Separately

In this section I examine the impact of 8:30 announcements on realised variance in the US and Canadian market separately. Table 2.1 determines the news days according to the US macroeconomic announcements and similarly Table 2.2 determines news days based on the Canadian macroeconomic announcements. I apply the same regression model (2.9) for the Full sample in each country. Table 2.11 reports the results for realised variance when I use only the US macroeconomic announcements. Spread as a private information shock is the only liquidity variable that is significant at the 5% level. Surprise as a public information shock does not reveal any significant results. A big portion of macroeconomic announcements are coming from the US and the results are not very different from Table 2.5. Table 2.12 reports the results when I use the Canadian macroeconomic announcements and results are neither significant for private information shocks nor for the public information shocks. It might imply that the US

TABLE 2.8: Realised Variance at 8:30 Announcements Using the Full Sample, without Lag.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 330 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liq	Liquidity Measure Without Lag (Full Announcements Sample)							
	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI				
Intercept	0.0972*** (0.0187)	0.0210*** (0.0210)	0.0185*** (0.0185)	0.0155*** (0.0155)	0.0147*** (0.0147)				
1_{News}	0.0257 (0.0264)	0.0108 (0.0248)	0.0023 (0.0233)	0.0139 (0.0190)	0.0224 (0.0181)				
1_{Post}	-0.0133 (0.0298)	-0.0063 (0.0274)	0.0037 (0.0271)	-0.0150 (0.0210)	0.0065 (0.0205)				
$1_{News}1_{Post}$	0.0577 (0.0524)	0.1674** (0.0728)	0.1156 (0.0745)	0.1692** (0.0572)	0.1477** (0.0628)				
X_t	0.0092 (0.0469)	-0.0017 (0.0023)	-0.0018 (0.0013)	0.0003 (0.0004)	0.0008 (0.0023)				
$X_{t}1_{News}$	-0.0460 (0.0732)	0.0003 (0.0028)	0.0015 (0.0021)	-0.0002 (0.0004)	-0.0038 (0.0035)				
$X_t 1_{Post}$	0.0649 (0.0747)	0.0017 (0.0031)	0.0001 (0.0025)	0.0006 (0.0004)	-0.0008 (0.0040)				
$X_t 1_{News} 1_{Post}$	0.2337** (0.1111)	-0.0038 (0.0055)	0.0041 (0.0066)	-0.0007 (0.0005)	-0.0005 (0.0066)				
$\operatorname{Sur}_t 1_{Post}$	-0.0003 (0.0217)	0.0159 (0.0260)	0.0176 (0.0264)	0.0198 (0.0282)	0.0159 (0.0262)				
Adjusted R ²	0.2724	0.1338	0.1350	0.1321	0.1346				

TABLE 2.9: Realised Variance at 8:30 Announcements Using the RI Sample, without Lag.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 240 observations in RI announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liq	Liquidity Measure Without Lag (RI Announcements Sample)						
	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI			
Intercept	0.0972*** (0.0188)	0.1102*** (0.0211)	0.1122*** (0.0186)	0.0949*** (0.0156)	0.0974*** (0.0148)			
1_{News}	0.0171 (0.0290)	0.0149 (0.0275)	0.0144 (0.0251)	0.0160 (0.0215)	0.0219 (0.0197)			
1_{Post}	-0.0133 (0.0300)	-0.0063 (0.0275)	0.0037 (0.0272)	-0.0150 (0.0211)	0.0065 (0.0206)			
$1_{News}1_{Post}$	$0.0470 \\ (0.0570)$	0.1070^* (0.0646)	0.0251 (0.0703)	0.1307** (0.0610)	0.1142* (0.0619)			
X_t	0.0092 (0.0472)	-0.0017 (0.0023)	-0.0018 (0.0013)	0.0003 (0.0004)	0.0008 (0.0024)			
$X_t 1_{News}$	-0.0123 (0.0776)	-0.0004 (0.0030)	-0.0001 (0.0021)	-0.0002 (0.0004)	-0.0034 (0.0039)			
$X_t 1_{Post}$	0.0649 (0.0751)	0.0017 (0.0032)	0.0001 (0.0025)	0.0006 (0.0004)	-0.0008 (0.0040)			
$X_t 1_{News} 1_{Post}$	0.1377 (0.1060)	0.0016 (0.0064)	0.0150 (0.0103)	-0.0004 (0.0010)	0.0005 (0.0068)			
Sur_t1_{Post}	0.0220 (0.0247)	0.0311 (0.0335)	0.0274 (0.0326)	0.0285 (0.0373)	0.0326 (0.0336)			
Adjusted R ²	0.3116	0.1603	0.1960	0.1612	0.1619			

TABLE 2.10: Realised Variance at 8:30 Announcements Using the MI Sample, without Lag.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 176 observations in MI announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liq	uidity Measure V	Vithout Lag (MI A	Announcements Sa	ample)
	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI
Intercept	0.0972*** (0.0191)	0.1102*** (0.0215)	0.1122*** (0.0190)	0.0949*** (0.0159)	0.0974*** (0.0151)
1_{News}	0.0211 (0.0392)	-0.0080 (0.0289)	-0.0139 (0.0407)	-0.0302 (0.0246)	0.0055 (0.0218)
1_{Post}	-0.0133 (0.0306)	-0.0063 (0.0281)	0.0037 (0.0278)	-0.0150 (0.0215)	0.0065 (0.0210)
$1_{News}1_{Post}$	0.1863 (0.1800)	0.2459 (0.1696)	0.3390** (0.1601)	0.4707*** (0.1980)	0.3701** (0.1926)
X_t	0.0092 (0.0481)	-0.0017 (0.0023)	-0.0018 (0.0013)	0.0003 (0.0004)	0.0008 (0.0024)
$X_t 1_{News}$	-0.0989 (0.1286)	0.0008 (0.0027)	0.0015 (0.0046)	0.0009 (0.0007)	-0.0031 (0.0040)
$X_t 1_{Post}$	0.0649 (0.0765)	0.0017 (0.0032)	0.0001 (0.0026)	0.0006 (0.0004)	-0.0008 (0.0041)
$X_t 1_{News} 1_{Post}$	0.1986 (0.1506)	0.0326 (0.0283)	0.0157 (0.0183)	-0.0021 (0.0045)	0.0156 (0.0301)
$\operatorname{Sur}_t 1_{Post}$	0.0166 (0.1014)	-0.0903 (0.1121)	-0.1306 (0.1360)	-0.0753 (0.1178)	-0.0735 (0.1168)
Adjusted R ²	0.4399	0.3705	0.3569	0.3187	0.3190

TABLE 2.11: Realised Variance at 8:30 Announcements in US, Using the Full Sample.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 321 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liquidity Measure in US (Full Announcements Sample)						
Variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI		
Intercept	0.0743*** (0.0185)	0.0879*** (0.0190)	0.0889*** (0.0182)	0.0748*** (0.0144)	0.0756*** (0.0143)		
1_{News}	0.0248 (0.0305)	0.0007 (0.0260)	-0.0009 (0.0266)	0.0110 (0.0205)	0.0194 (0.0208)		
1_{Post}	-0.0418 (0.0274)	-0.0317 (0.0238)	-0.0332 (0.0255)	-0.0258 (0.0189)	-0.0219 (0.0193)		
$1_{News}1_{Post}$	$0.0606 \ (0.0570)$	0.1752^{***} (0.0722)	0.1293^* (0.0729)	$0.1473^{***} $ (0.0601)	0.1430** (0.0636)		
RV_{t-15}	0.3689*** (0.0413)	0.3694*** (0.0401)	0.3673*** (0.0414)	0.3745*** (0.0421)	0.3727*** (0.0401)		
$RV_{t-15}1_{News}$	-0.0172 (0.1588)	-0.0085 (0.1618)	-0.0111 (0.1611)	-0.0300 (0.1695)	-0.0251 (0.1564)		
$RV_{t-15}1_{Post}$	0.1671* (0.0968)	0.1883* (0.1062)	0.1623^* (0.0995)	0.1176 (0.1505)	0.1935^* (0.1067)		
$\text{RV}_{t-15}1_{News}1_{Post}$	-0.0628 (0.3231)	0.0334 (0.3520)	-0.0335 (0.3600)	0.0526 (0.3806)	0.0240 (0.3411)		
X_t	-0.0049 (0.0470)	-0.0024 (0.0015)	-0.0022** (0.0011)	-0.0001 (0.0002)	-0.0010 (0.0017)		
$X_t 1_{News}$	-0.0387 (0.0739)	0.0020 (0.0023)	0.0020 (0.0019)	0.0001 (0.0002)	-0.0024 (0.0029)		
$X_t 1_{Post}$	0.0697 (0.0723)	0.0011 (0.0018)	0.0014 (0.0022)	0.0003 (0.0004)	-0.0013 (0.0022)		
$X_t 1_{News} 1_{Post}$	0.2362** (0.1105)	-0.0065 (0.0058)	0.0021 (0.0066)	-0.0004 (0.0004)	-0.0013 (0.0065)		
$\operatorname{Sur}_t 1_{Post}$	-0.0056 (0.0205)	0.0096 (0.0245)	0.0124 (0.0248)	0.0159 (0.0264)	0.0102 (0.0247)		
Adjusted R ²	0.2826	0.1462	0.1424	0.1407	0.1459		

macroeconomic announcements are more important and include more private information shocks compared to the Canadian announcements.

TABLE 2.12: Realised Variance at 8:30 Announcements in CA, Using the Full Sample.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 118 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liquidity Measure in CA (Full Announcements Sample)						
· Marabaco	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI		
Intercept	0.0624*** (0.0183)	0.0919*** (0.0194)	0.0884*** (0.0186)	0.0761*** (0.0147)	0.0786*** (0.0145)		
1_{News}	0.0459 (0.0313)	0.0075 (0.0251)	0.0059 (0.0276)	0.0157 (0.0207)	$0.0208 \ (0.0214)$		
1_{Post}	-0.0586 (0.0397)	-0.0171 (0.0325)	-0.0161 (0.0482)	-0.0227 (0.0224)	-0.0103 (0.0269)		
$1_{News}1_{Post}$	0.0656 (0.0654)	0.1465^* (0.0776)	0.0994 (0.0873)	0.1310** (0.0635)	0.1229^* (0.0685)		
RV_{t-15}	0.3645*** (0.0432)	0.3650*** (0.0406)	0.3610*** (0.0427)	0.3701*** (0.0416)	0.3720*** (0.0396)		
$\mathrm{RV}_{t-15}1_{News}$	-0.1106 (0.1699)	-0.1121 (0.1652)	-0.1054 (0.1673)	-0.1347 (0.1730)	-0.1263 (0.1647)		
$\mathrm{RV}_{t-15}1_{Post}$	$0.1972* \\ (0.1164)$	0.2159* (0.1257)	0.1772 (0.1348)	-0.1023 (0.2803)	$0.2155* \\ (0.1250)$		
$\text{RV}_{t-15}1_{News}1_{Post}$	0.1134 (0.3533)	0.2796 (0.3706)	0.1929 (0.3969)	0.5223 (0.4709)	$0.2640 \\ (0.3611)$		
X_t	0.0459 (0.0431)	-0.0029^* (0.0016)	-0.0019^* (0.0011)	-0.0001 (0.0002)	-0.0018 (0.0018)		
$X_t 1_{News}$	-0.1009 (0.0744)	0.0019 (0.0022)	0.0018 (0.0020)	0.0001 (0.0002)	-0.0004 (0.0032)		
$X_t 1_{Post}$	0.1588 (0.1581)	0.0005 (0.0022)	0.0006 (0.0040)	0.0013 (0.0010)	-0.0015 (0.0028)		
$X_t 1_{News} 1_{Post}$	0.1523 (0.1798)	-0.0063 (0.0061)	0.0027 (0.0076)	-0.0015 (0.0010)	-0.0036 (0.0071)		
$\operatorname{Sur}_t 1_{Post}$	-0.0067 (0.0210)	0.0085 (0.0250)	0.0115 (0.0257)	0.0155 (0.0276)	$0.0085 \ (0.0253)$		
Adjusted R ²	0.2855	0.1556	0.1481	0.1519	0.1551		

2.4.3 Surprise Variable Analysis

In this section, I provide more analysis on surprise variable as a proxy for public information shocks. The public information shocks do not show any significant results throughout the study. I drop this variable from my regression analysis to check if liquidity variables behave differently. Moreover, this is a test to examine what effect a purely private information shock might have on volatility. Table 2.13 shows that deleting the surprise from the regression model (2.9) does not change the behaviour of liquidity variables and their significance level. The reason might be due to the less important role of surprise in results and deleting this variable from the regression model does not cause any significant changes.

Next, I separate positive and negative surprises into two groups. I measure surprise as the difference between actual and survey values that is estimated by financial analyst (surprise=actual-survey). However, a positive difference has not the same meaning for all announcements. For example, more actual inflation than survey value is a bad news whereas more actual production than survey value is a good news. To have announcements with the same impact direction (e.g. good news), I set up actual minus survey for good news and survey minus actual for bad news. I count all announcements in Tables 2.1 and 2.2 as good news except Business Inventories, Wholesale Inventories MoM, Continuing Claims, Initial Jobless Claims, Unemployment Rate, Employment Cost Index, Unemployment Rate, PCE Core YoY, Import Price Index MoM CPI MoM, Industrial Production Price MoM, New Housing Price Index MoM, Raw materials Price Index MoM, and CPI YoY. In this regard, if the sign for all announcements (i.e. good/bad news) is positive we have positive shock otherwise is negative shock.

I revise the regression model (2.9) and separate the surprise variable into two components: positive and negative surprise. I define two dummy

TABLE 2.13: Realised Variance at 8:30 Announcements Using the Full Sample with no Surprise.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 330 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liquidity Measure (Full Announcements Sample)						
Value	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI		
Intercept	0.0642*** (0.0181)	0.0775*** (0.0183)	0.0786*** (0.0176)	0.0687*** (0.0146)	0.0681*** (0.0136)		
1_{News}	0.0397 (0.0300)	0.0239 (0.0241)	0.0158 (0.0266)	0.0256 (0.0205)	0.0325 (0.0204)		
1_{Post}	-0.0317 (0.0271)	-0.0201 (0.0232)	-0.0253 (0.0250)	-0.0181 (0.0190)	-0.0134 (0.0187)		
$1_{News}1_{Post}$	0.0351 (0.0507)	0.1604*** (0.0554)	0.1289** (0.0575)	0.1412*** (0.0496)	0.1376*** (0.0476)		
RV_{t-15}	0.3821*** (0.0405)	0.3826*** (0.0394)	0.3810*** (0.0406)	0.3807*** (0.0421)	0.3847*** (0.0392)		
$\mathrm{RV}_{t-15}1_{News}$	-0.1217 (0.1652)	-0.1255 (0.1634)	-0.1211 (0.1648)	-0.1510 (0.1727)	-0.1353 (0.1627)		
$\text{RV}_{t-15}1_{Post}$	0.1443 (0.0983)	0.1695 (0.1107)	0.1452 (0.1006)	0.0953 (0.1673)	0.1745 (0.1117)		
$\text{RV}_{t-15}1_{News}1_{Post}$	0.1421 (0.3379)	0.2940 (0.3535)	0.1990 (0.3758)	0.3026 (0.4029)	0.2751 (0.3456)		
X_t	0.0095 (0.0428)	-0.0018 (0.0014)	-0.0017 (0.0011)	$0.0000 \\ (0.0002)$	-0.0006 (0.0016)		
$X_t 1_{News}$	-0.0464 (0.0702)	0.0005 (0.0021)	0.0016 (0.0020)	0.0001 (0.0002)	-0.0020 (0.0030)		
$X_t 1_{Post}$	0.0612 (0.0699)	0.0005 (0.0018)	0.0014 (0.0023)	0.0002 (0.0004)	-0.0015 (0.0022)		
$X_t 1_{News} 1_{Post}$	0.2328*** (0.1064)	-0.0060 (0.0058)	0.0013 (0.0066)	-0.0003 (0.0004)	-0.0031 (0.0067)		
Adjusted R ²	0.2952	0.1649	0.1575	0.1535	0.1646		

TABLE 2.14: Realised Variance at 8:30 Announcements Using the Full Sample with Positive and Negative Surprises.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 311 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liquidity Measure (Full Announcements Sample)						
variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI		
Intercept	0.0634*** (0.0183)	0.0804*** (0.0194)	0.0814*** (0.0189)	0.0687*** (0.0147)	0.0699*** (0.0145)		
1_{News}	0.0438 (0.0306)	0.0236 (0.0251)	0.0155 (0.0277)	0.0256 (0.0205)	0.0329 (0.0212)		
1_{Post}	-0.0308 (0.0273)	-0.0229 (0.0241)	-0.0281 (0.0259)	-0.0181 (0.0190)	-0.0152 (0.0194)		
$1_{News}1_{Post}$	0.0391 (0.0542)	0.1563** (0.0648)	$0.1276^* \ (0.0706)$	$0.1405^{***} (0.0567)$	0.1334** (0.0574)		
RV_{t-15}	0.3779*** (0.0421)	0.3781*** (0.0411)	0.3759*** (0.0425)	0.3809*** (0.0422)	0.3813*** (0.0407)		
$\mathrm{RV}_{t-15}1_{News}$	-0.1330 (0.1681)	-0.1364 (0.1655)	-0.1307 (0.1674)	-0.1511 (0.1740)	-0.1476 (0.1647)		
$\mathrm{RV}_{t-15}1_{Post}$	0.1485 (0.0993)	0.1740 (0.1116)	0.1503 (0.1017)	0.0951 (0.1678)	0.1779 (0.1125)		
$\text{RV}_{t-15}1_{News}1_{Post}$	0.2022 (0.3334)	0.3680 (0.3512)	0.2801 (0.3615)	0.3860 (0.3855)	0.3492 (0.3452)		
X_t	0.0191 (0.0411)	-0.0020 (0.0015)	-0.0018 (0.0012)	$0.0000 \\ (0.0002)$	-0.0008 (0.0016)		
$X_t 1_{News}$	-0.0612 (0.0702)	0.0006 (0.0021)	0.0016 (0.0020)	0.0001 (0.0002)	-0.0020 (0.0030)		
$X_t 1_{Post}$	0.0516 (0.0691)	0.0007 (0.0018)	0.0015 (0.0023)	0.0002 (0.0004)	-0.0013 (0.0022)		
$X_t 1_{News} 1_{Post}$	0.2422 (0.1065)	-0.0057 (0.0056)	0.0010 (0.0066)	-0.0004 (0.0005)	-0.0026 (0.0063)		
$Sur_t 1_{Pos} 1_{Post}$	-0.0015 (0.0174)	0.0093 (0.0232)	0.0102 (0.0241)	0.0133 (0.0249)	0.0094 (0.0235)		
$\operatorname{Sur}_t 1_{Neg} 1_{Post}$	0.0492 (0.0463)	0.0715 (0.0484)	0.0722 (0.0523)	0.0776 (0.0534)	$0.0704 \\ (0.0485)$		
Adjusted R ²	0.2924	0.1686	0.1617	0.1623	0.1680		

variables for each case. If the surprise is positive dummy is one otherwise is zero and a dummy in the same manner for negative surprises. Table 2.14 shows the results when I split the surprise into positive and negative. Results are robust with the main regression model and separating news will not change the significant level of other variables. Surprise as a public information shock did not reveal any significant role compared to private information shock and this separation will not change the relationship. Among all liquidity variables, spread as a proxy for private information shock is significant at the 5% level.

2.4.4 Alternative Measures of Volatility

In this section, I evaluate two alternative measures of realised variance: AR and R^{resi} . I set up my regression model using (2.9), but I substitute AR and R^{resi} (as described in Section 2.2.3) in place of RV as the dependent variable.

Many studies use the AR measure to examine ex-post volatility (see Jones, Kaul, and Lipson (1994), Ahn, Bae, and K. Chan (2001), and Foucault, Moinas, and Theissen (2007)). In line with this literature, I use the AR measure along with the three different announcement samples: Full, $Relatively\ Important$, and $Most\ Important$. However, the results are significant only for the Full and $Relatively\ Important$ samples. Tables 2.15 and 2.16 report these results.

Table 2.15 shows that AR is sensitive to spread and absolute order flow after the announcements on news days. As before, I capture this effect by interaction of $\mathbf{1}_{News}$ and $\mathbf{1}_{post}$ with explanatory variables. Spread (the difference between the best bid quote and the best ask quote) sensitivity has a positive coefficient and absolute order flow (the difference between the buy volume and the sell volume during the 15-minute interval) has a negative coefficient for the arrival of new information. I note that

TABLE 2.15: Absolute Return at 8:30 Announcements Using the Full Sample.

Results of regression of dependent variable (AR) absolute return on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 330 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors. All numbers are scaled by 1000.

Variables		Liquidity Meas	ure (Full Announ	cements Sample)	
variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI
Intercept	0.1130* (0.0670)	0.2200*** (0.0620)	0.2900*** (0.0680)	0.1890*** (0.0380)	0.1820*** (0.0420)
1_{News}	0.1720^* (0.0990)	0.0370 (0.0780)	-0.0400 (0.0870)	$0.0360 \\ (0.0530)$	$0.0670 \\ (0.0570)$
1_{Post}	0.1790* (0.0980)	0.1280* (0.0780)	0.0210 (0.1090)	0.1070* (0.0630)	0.1510*** (0.0610)
$1_{News}1_{Post}$	-0.2300^* (0.1320)	-0.0900 (0.1070)	-0.0900 (0.1390)	-0.0300 (0.0890)	-0.0900 (0.0890)
AR_{t-15}	225.2920 (212.9000)	186.3580 (213.3000)	241.9630 (207.3000)	403.0470 (255.7000)	166.5910 (214.2000)
$AR_{t-15}1_{News}$	-99.8100 (243.9000)	-78.6100 (247.5000)	$ \begin{array}{c} -133.9100 \\ (242.9000) \end{array} $	-375.5000 (301.0000)	-65.9700 (248.2000)
$AR_{t-15}1_{Post}$	$ \begin{array}{c} -316.3100 \\ (259.2000) \end{array} $	$ \begin{array}{c} -220.6100 \\ (260.7000) \end{array} $	$ \begin{array}{c} -334.4500 \\ (255.9000) \end{array} $	-618.1600** (296.5000)	-181.6400 (262.6000)
$AR_{t-15}1_{News}1_{Post}$	334.6920 (309.3000)	218.4160 (318.6000)	340.7050 (310.2000)	669.8240** (359.9000)	218.9390 (318.5000)
X_t	0.2410* (0.1290)	-0.0066 (0.0079)	-0.0200*** (0.0065)	-0.0024 (0.0015)	0.0017 (0.0090)
$X_t 1_{News}$	-0.4800^{**} (0.2380)	0.0016 (0.0092)	0.0140* (0.0081)	0.0028* (0.0016)	-0.0098 (0.0110)
$X_t 1_{Post}$	-0.1600 (0.2300)	-0.0003 (0.0085)	0.0180** (0.0094)	0.0039** (0.0020)	-0.0100 (0.0100)
$X_t 1_{News} 1_{Post}$	0.5560* (0.3060)	0.0045 (0.0110)	0.0005 (0.0120)	-0.0041^{**} (0.0020)	0.0140 (0.0130)
$\operatorname{Sur}_t 1_{Post}$	0.0190 (0.0270)	0.0320 (0.0310)	0.0390 (0.0320)	0.0250 (0.0340)	0.0310 (0.0310)
Adjusted R ²	0.0070	0.0028	0.0069	0.0033	0.0033

in both cases, the coefficient suggests volatility rises more when liquidity falls (i.e. spread rises, or order flow declines).

The results for spread and absolute order flow are significant at the

TABLE 2.16: Absolute Return at 8:30 Announcements Using the RI Sample.

Results of regression of dependent variable (AR) absolute return on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 240 observations in RI announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors. All numbers are scaled by 1000.

Variables	Liquidity Measure (RI Announcements Sample)							
variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI			
Intercept	0.1130* (0.0680)	0.2200*** (0.0620)	0.2900*** (0.0690)	0.1890*** (0.0390)	0.1820*** (0.0420)			
1_{News}	0.1360 (0.0900)	0.0310 (0.0780)	-0.0700 (0.0870)	0.0210 (0.0560)	$0.0560 \\ (0.0580)$			
1_{Post}	0.1790^* (0.0990)	0.1280^* (0.0780)	0.0210 (0.1100)	0.1070* (0.0630)	0.1510** (0.0620)			
$1_{News}1_{Post}$	-0.2100^* (0.1310)	-0.0700 (0.1130)	-0.1200 (0.1440)	-0.0200 (0.1050)	-0.0800 (0.0970)			
AR_{t-15}	$225.2920 \\ (214.7000)$	186.3580 (215.1000)			166.5910 (216.1000)			
$AR_{t-15}1_{News}$	-36.9800 (252.3000)	$ -5.4600 \\ (252.5000) $			17.7010 (253.2000)			
$AR_{t-15}1_{Post}$	$ \begin{array}{c} -316.3100 \\ (261.4000) \end{array} $	$ \begin{array}{c} -220.6100 \\ (263.0000) \end{array} $	-334.4500 (258.1000)	-618.1600** (299.1000)	-181.6400 (264.9000)			
$AR_{t-15}1_{News}1_{Post}$	193.2070 (337.9000)	67.0230 139.5820 (349.0000) (334.4000)		471.0560 (379.4000)	51.6330 (348.1000)			
X_t	$0.2410^* $ (0.1310)	-0.0066 (0.0080)	-0.0200** (0.0066)	-0.0024 (0.0016)	0.0017 (0.0091)			
$X_t 1_{News}$	-0.3600 (0.2040)	0.0008 (0.0094)	0.0170** (0.0084)	0.0029* (0.0018)	-0.0100 (0.0110)			
$X_t 1_{Post}$	-0.1600 (0.2320)	-0.0003 (0.0086)	0.0180** (0.0095)	0.0039** (0.0020)	-0.0100 (0.0100)			
$X_t 1_{News} 1_{Post}$	0.4480* (0.2830)	0.0039 (0.0110)	0.0140 (0.0130)	-0.0038 (0.0024)	$0.0160 \\ (0.0130)$			
$\operatorname{Sur}_t 1_{Post}$	$0.0350 \\ (0.0370)$	0.0440 (0.0420)	0.0340 (0.0390)	0.0370 (0.0450)	$0.0460 \\ (0.0430)$			
Adjusted R ²	0.00810	0.0022	0.0143	0.0023	0.0028			

Table 2.17: Absolute Return at 8:30 Announcements Using the MI Sample.

Results of regression of dependent variable (AR) absolute return on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 176 observations in MI announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors. All numbers are scaled by 1000.

Variables		Liquidity Meas	ure (MI Announ	cements Sample)		
variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI	
Intercept	0.1130 (0.0700)	0.2200*** (0.0640)	0.2900*** (0.0710)	0.1890*** (0.0400)	0.1820*** (0.0440)	
1_{News}	-0.0200 (0.2070)	0.0220 (0.1320)	-0.2100 (0.1090)	-0.0200 (0.0740)	0.0570 (0.1260)	
1_{Post}	0.1790* (0.1030)	0.1280 (0.0810)	0.0210 (0.1140)	$0.1070^* \ (0.0660)$	0.1510** (0.0640)	
$1_{News}1_{Post}$	-0.2600^* (0.3220)	-0.2700 (0.2380)	0.1390 (0.2630)	-0.0200 (0.2990)	-0.2100 (0.2270)	
AR_{t-15}	225.2920 (221.8000)	186.3580 (222.3000)			166.5910 (223.2000)	
$AR_{t-15} 1_{News}$	-51.5400 (364.6000)	$ \begin{array}{c} -150.3000 \\ (324.4000) \end{array} $	-353.3900 (291.9000)	-639.0400 (366.8000)	-129.3400 (332.0000)	
$AR_{t-15}1_{Post}$	$ \begin{array}{c} -316.3100 \\ (270.1000) \end{array} $	$ \begin{array}{c} -220.6100 \\ (271.7000) \end{array} $	-334.4500 (266.6000)	-618.1600** (309.1000)	$ \begin{array}{c} -181.6400 \\ (273.6000) \end{array} $	
$AR_{t-15}1_{News}1_{Post}$	493.3440 (680.6000)	273.3010 (632.5000)	485.3050 (621.5000)	914.8790 (790.1000)	475.3600 (703.5000)	
X_t	0.2410* (0.1310)	-0.0066^* (0.0083)	-0.0200^{***} (0.0068)	-0.0024 (0.0016)	0.0017 (0.0094)	
$X_t 1_{News}$	0.1940 (0.5780)	0.0058 (0.0120)	0.0450*** (0.0160)	0.0072^{***} (0.0025)	-0.0025 (0.0240)	
$X_t 1_{Post}$	-0.1600 (0.2400)	-0.0003 (0.0088)	0.0180^* (0.0098)	0.0039** (0.0020)	-0.0100 (0.0100)	
$X_t 1_{News} 1_{Post}$	-0.0900 (0.6140)	0.0420 (0.0340)	-0.0300 (0.0240)	-0.0061 (0.0049)	$0.0460 \\ (0.0440)$	
Sur_t1_{Post}	$0.1470 \\ (0.2060)$	0.0410 (0.2050)	-0.0023 (0.2180)	0.0580 (0.2000)	0.0540 (0.2080)	
Adjusted R ²	0.0089	0.0042	0.0045	0.0005	-0.0006	

10% and 5% level respectively. The results indicate that when I use the AR measure, spread and absolute order flow become important microstructure variables (as compared to only spread when I use the *Realised Variance* measure). Surprise sensitivity as a proxy for public information shocks remains insignificant in all cases. The results for the *Relatively Important* sample (Table 2.16) show that the spread is the only significant liquidity measure. However, the (unreported) results for the *Most Important* sample are insignificant.

In high frequency data, there is a possibility that mid-quotes differ from their fair value due to market microstructure effects (Hasbrouck, 1993). My third measure, which is modified absolute return (R^{resi}), regresses the absolute return of each interval on its lag value during each trading day (7:00 a.m-5:00 p.m) and retains the residuals (see Section 2.2.3). The absolute value of the retained residuals before and after 8:30 is the modified absolute return that is used as a proxy for measuring volatility (Foucault, Moinas, and Theissen, 2007). Tables 2.18 and 2.19 report my results for using the R^{resi} measure for the Full and Relatively Important samples respectively.

Table 2.18 shows that the second alternative measure of volatility (R^{resi}) after the announcements is sensitive to depth at the second best quote and absolute order flow, using the Full sample. I postulate that with the arrival of macroeconomic announcements and the existence of asymmetric information, traders avoid quoting at the best price to wait for the market to expose hidden information. Hence, a rise in the depth at the second best quote (D_1) is also a symptom of illiquidity. In this regard, I note that realised variance has a significantly higher sensitivity to the depth behind the best quote (D_1) . The results indicate that depth at the second best quote (D_1) and absolute order flow (AOF) are significant at the 1% and 5% level (respectively) for the Full sample. Absolute order flow declines when the market is less liquid. Hence my negative coeffi-

TABLE 2.18: Modified Absolute Return at 8:30 Announcements Using the Full Sample.

Results of regression of dependent variable (R^{resi}) modified absolute return on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 330 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors. All numbers are scaled by 1000.

Variables	Liquidity Measure (Full Announcements Sample)							
variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI			
Intercept	0.0760** (0.0330)	0.0570** (0.0260)	0.0380 (0.0290)	0.0780*** (0.0280)	0.0670*** (0.0290)			
1_{News}	0.1230** (0.0640)	0.0930* (0.0530)	0.1450*** (0.0440)	0.0910*** (0.0360)	0.1160*** (0.0380)			
1_{Post}	0.0062 (0.0530)	$0.0440 \\ (0.0370)$	0.0910** (0.0460)	-0.0200 (0.0350)	$0.0130 \\ (0.0390)$			
$1_{News}1_{Post}$	-0.0800 (0.0830)	-0.0200 (0.0710)	-0.2000^{***} (0.0760)	0.0290 (0.0550)	-0.0100 (0.0590)			
\mathbf{R}^{resi}_{t-15}	373.1200** (165.1000)	260.8360 (261.1000)	354.7160** (174.7000)	345.0530** (182.6000)	336.5190 (276.0000)			
$\mathbf{R}^{resi}_{t-15}1_{News}$	-365.8800* (209.8000)	$ \begin{array}{c} -254.5600 \\ (290.0000) \end{array} $	$ \begin{array}{c} -340.9100 \\ (217.2000) \end{array} $	-332.0600 (241.4000)	-336.3600 (303.8000)			
$R^{resi}_{t-15}1_{Post}$	155.7180 (225.0000)	254.6440 (302.5000)	221.1350 (217.1000)	17.1030 (250.0000)	195.5450 (313.4000)			
$\mathbf{R}_{t-15}^{resi}1_{News}1_{Post}$	2.4110 (275.8000)	-103.6400 (345.7000)	$ \begin{array}{c} -61.9600 \\ (270.0000) \end{array} $	139.9820 (311.5000)	-15.4400 (353.2000)			
X_t	-0.0400 (0.0470)	0.0038 (0.0042)	0.0039 (0.0027)	-0.0003 (0.0007)	0.0011 (0.0045)			
$X_t 1_{News}$	-0.0700 (0.1740)	-0.0006 (0.0078)	-0.0060 (0.0040)	0.0003 (0.0007)	-0.0070 (0.0070)			
$X_t 1_{Post}$	0.0420 (0.0990)	-0.0078 (0.0055)	-0.0100^{***} (0.0042)	0.0018** (0.0008)	-0.0001 (0.0120)			
$X_t 1_{News} 1_{Post}$	0.2140 (0.1970)	0.0051 (0.0095)	0.0260*** (0.0069)	-0.0019^{**} (0.0009)	0.0028 (0.0130)			
$\operatorname{Sur}_t 1_{Post}$	0.0140 (0.0220)	0.0250 (0.0260)	0.0300 (0.0270)	0.0280 (0.0290)	$0.0240 \\ (0.0260)$			
Adjusted R ²	0.0079	0.0021	0.0089	0.0020	0.0024			

TABLE 2.19: Modified Absolute Return at 8:30 Announcements Using the RI Sample.

Results of regression of dependent variable (R^{resi}) modified absolute return on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 240 observations in RI announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors. All numbers are scaled by 1000.

Variables		Liquidity Mea	sure (RI Announc	ements Sample)		
V 11111 1C3	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI	
Intercept	0.0730** (0.0330)	0.0530** (0.0260)	0.0350 (0.0290)	0.0740*** (0.0280)	0.0640** (0.0290)	
1_{News}	0.1110* (0.0610)	0.0420 (0.0660)	0.1020** (0.0440)	$0.0710^* \ (0.0410)$	0.1010** (0.0420)	
1_{Post}	0.0094 (0.0540)	0.0470 (0.0370)	0.0950** (0.0460)	-0.0200 (0.0350)	$0.0160 \\ (0.0390)$	
$1_{News}1_{Post}$	-0.1000 (0.0840)	0.0110 (0.0840)	-0.2300*** (0.0800)	0.0340 (0.0680)	-0.0200 (0.0660)	
\mathbf{R}^{resi}_{t-15}	401.0010** (166.5000)	282.9250 (265.8000)	381.8710** (176.9000)	373.5410** (185.4000)	360.2420 (280.6000)	
$\mathbf{R}_{t-15}^{resi}1_{News}$	$ \begin{array}{c} -313.8700 \\ (256.9000) \end{array} $	$ \begin{array}{rrr} -197.0500 & -306.4500 \\ (325.2000) & (262.7000) \end{array} $		$ \begin{array}{c} -315.0000 \\ (274.0000) \end{array} $	$ \begin{array}{c} -289.5700 \\ (339.1000) \end{array} $	
$\mathbf{R}_{t-15}^{resi}1_{Post}$	124.1140 (229.7000)	229.2650 (309.3000)	187.9950 (222.3000)	$ \begin{array}{c} -23.4600 \\ (251.5000) \end{array} $	169.2870 (319.6000)	
$\mathbf{R}_{t-15}^{resi}1_{News}1_{Post}$	43.6650 (316.4000)	-51.1200 -14.3200 (377.2000) (309.5000)		225.8660 (342.6000)	40.5760 (386.2000)	
X_t	-0.0400 (0.0470)	0.0040 (0.0043)	0.0039 (0.0027)	-0.0003 (0.0007)	0.0012 (0.0046)	
$X_t 1_{News}$	-0.0800 (0.1710)	0.0056 (0.0100)	-0.0017 (0.0041)	0.0005 (0.0008)	-0.0071 (0.0088)	
$X_t 1_{Post}$	0.0440 (0.1000)	-0.0077 (0.0056)	-0.0100** (0.0042)	0.0018** (0.0009)	0.0003 (0.0120)	
$X_t 1_{News} 1_{Post}$	0.2210 (0.1950)	-0.0025 (0.0120)	0.0330*** (0.0082)	-0.0025^* (0.0014)	0.0021 (0.0150)	
$\operatorname{Sur}_t 1_{Post}$	0.0320 (0.0330)	0.0390 (0.0390)	0.0310 (0.0360)	0.0450 (0.0390)	$0.0400 \\ (0.0390)$	
Adjusted R ²	0.0126	0.0053	0.0182	0.0041	0.0044	

cient is consistent: volatility becomes more sensitive to liquidity declines post-announcement. Similarly, Table 2.19 shows that the R^{resi} measure is sensitive to depth at the second best quote and absolute order flow after

TABLE 2.20: Modified Absolute Return at 8:30 Announcements Using the MI Sample.

Results of regression of dependent variable (R^{resi}) modified absolute return on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 176 observations in MI announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors. All numbers are scaled by 1000.

Variables	Liquidity Measure (MI Announcements Sample)							
variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI			
Intercept	0.0720** (0.0350)	0.0530** (0.0280)	0.0360 (0.0300)	0.0740*** (0.0290)	0.0630** (0.0300)			
1_{News}	$0.0800 \\ (0.0750)$	$0.0450 \\ (0.0700)$	0.0440 (0.0680)	0.0210 (0.0480)	0.0390 (0.0660)			
1_{Post}	0.0095 (0.0570)	0.0470 (0.0390)	0.0930** (0.0480)	-0.0200 (0.0360)	0.0160 (0.0410)			
$1_{News}1_{Post}$	-0.0700 (0.2150)	-0.1600 (0.1580)	-0.0800 (0.1860)	$0.0300 \\ (0.2300)$	-0.1400 (0.1620)			
\mathbf{R}^{resi}_{t-15}	102.8500** (175.7000)	289.3630 (280.1000)	384.8020** (187.1000)	377.7640** (195.8000)	364.3410 (297.0000)			
$R^{resi}_{t-15}1_{News}$	$ \begin{array}{c} -249.1500 \\ (268.0000) \end{array} $	$ \begin{array}{c} -206.1300 \\ (371.27000) \end{array} $	-343.4100 (307.5000)	-653.0300^{*} (382.1000)	$ \begin{array}{c} -285.4200 \\ (379.4000) \end{array} $			
$R^{resi}_{t-15}1_{Post}$	130.7780 (242.4000)	231.5660 (326.1000)	191.9710 (234.9000)	-19.5200 (264.7000)	173.5970 (337.7000)			
$R^{resi}_{t-15}1_{News}1_{Post}$	593.8680 (632.3000)	311.2790 (717.7000)	550.4130 (644.9000)	1153.7840* (700.1000)	892.0710 (666.6000)			
X_t	-0.0400 (0.0500)	0.0038 (0.0045)	0.0037 (0.0028)	-0.0003 (0.0007)	0.0012 (0.0049)			
$X_t 1_{News}$	0.5270** (0.2430)	-0.0003 (0.0086)	0.0030 (0.0100)	0.0029 (0.0018)	0.0058 (0.0140)			
$X_t 1_{Post}$	0.0420 (0.1050)	-0.0075 (0.0059)	-0.0099^{**} (0.0044)	0.0018** (0.0009)	0.0004 (0.0130)			
$X_t 1_{News} 1_{Post}$	-0.3800 (0.2620)	0.0370 (0.0260)	0.0120 (0.0160)	-0.0038 (0.0043)	0.0320 (0.0320)			
$\operatorname{Sur}_t 1_{Post}$	0.1490 (0.1690)	0.0520 (0.1730)	0.0390 (0.1810)	0.0700 (0.1680)	0.0830 (0.1640)			
Adjusted R ²	0.3100	0.0206	0.0129	0.0105	0.0229			

announcements when I use the *Relatively Important* sample. Depth at the second best quote is still significant at the 1% level, but absolute order flow is significant at the 10% level.

Overall, these results reinforce my main findings: volatility changes significantly with the arrival of macroeconomic announcements in the market. Among public information shocks and private information shocks, the latter show a significant impact on volatility. Although these alternative volatility measures lead to similar findings, the answer varies to the follow up question: which proxy for private information shocks captures this effect? Using the Realised Variance measure, the sensitivity comes through the spread. However, using the AR measure, spread and absolute order flow are the key measures. Using the R^{resi} measure, depth at the second best quote and absolute order flow are the proxies that capture the effect. The results confirm my main assertion that volatility is related to liquidity measures that are a proxy for private information shocks. However, the main results (using the Realised Variance measure) are most relevant. The reason for this is that when using the Realised Variance measure, I compute the returns at the best quotes for each one-minute interval and then aggregate the squared returns, while when using the AR and R^{resi} measures, I only look at the beginning and the end of 15-minute intervals. In other words, these measures miss the impact of quotes within the 15-minute interval, which, in high frequency data, have significant impacts. Hence I would advocate the use of spread as a measure for private information over absolute order imbalance or depth.

2.4.5 The Effect of All Liquidity Variables Jointly

In this section, I examine the effect of all liquidity measures jointly. So far, I have examined three different measures of volatility along with three different samples of macroeconomic announcements. I now implement the same regression model using (2.9) with the *Realised Variance* measure of volatility along with the three different samples. However, I now include all the liquidity measures. Spread at the best quote (S_0) and depth at the second best quote (S_0) are the main microstructure measures that are sig-

nificant when I use the *Realised Variance* measure along with the Full and MI samples. Table 2.21 reports my results.

Table 2.21 shows that when I examine the effect of all liquidity measures jointly using the *Realised Variance* measure, spread and depth at the second best quote are the microstructure variables that exhibit greater impact on *Realised Variance* after an announcement. When I use the *Full* sample the results are significant only for spread at the 1% level. However, when I use the *Most Important* sample, the results are significant for both spread and depth at the second best quote at the 10% level. The *Relatively Important* sample does not reveal any significant results. Surprise, as a measure of public information shocks, is insignificant for all my samples. These results are broadly consistent with my analysis in Section 2.3.2 when I use liquidity measures individually. Again, I note that rises in spread or depth at the second best quote are symptoms of illiquidity. The main distinction between the two sets of results is the finding of a private information shock using the *Most Important* sample.

Overall, the results show that when I test the effect of all the liquidity measures jointly, spread at the best quote (S_0) is the main explanatory variable that captures the private information shock effect. Depth at the second best quote (D_1) is another liquidity measure that has a significantly stronger relationship with volatility after the arrival of macroeconomic announcements but only in the *Most Important* sample (at the 10% level). Overall, with the arrival of macroeconomic announcements, as noted earlier, private information shocks are the main source of volatility increases in the market.

2.4.6 The Effect of News at Times Other Than at 8:30

So far, I have only examined the price changes in the Canadian FX market when there is an 8:30 announcement. A question arising here is: what if I

TABLE 2.21: The Effect of All Liquidity Variables Jointly on Realised Variance.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 330 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Explanatory	Samples			Explanatory	Samples			
Variables	Full	RI	MI	─ Variables	Full	RI	MI	
Intercept	0.0960*** (0.0315)	0.0960*** (0.0321)	0.0960*** (0.0345)	$\mathrm{D}_{1,t}$	-0.0017 (0.0017)	-0.0017 (0.0017)	-0.0017 (0.0018)	
1_{News}	0.0153 (0.0427)	0.0148 (0.0471)	0.0131 (0.0670)	$\mathrm{D}_{1,t}1_{News}$	0.0016 (0.0023)	$0.0000 \\ (0.0024)$	-0.0003 (0.0048)	
1_{Post}	-0.0622 (0.0518)	-0.0622 (0.0528)	-0.0622 (0.0568)	$\mathrm{D}_{1,t}1_{Post}$	0.0012 (0.0025)	0.0012 (0.0026)	0.0012 (0.0028)	
$1_{News}1_{Post}$	0.0853 (0.0895)	-0.0400 (0.1041)	-0.2323 (0.1607)	$\mathbf{D}_{1,t}1_{News}1_{Post}$	$0.0008 \ (0.0063)$	0.0054 (0.0081)	0.0163^* (0.0088)	
R_{t-15}	0.3356*** (0.0590)	0.3356*** (0.0601)	0.3356*** (0.0647)	AOF_t	0.0002 (0.0002)	0.0002 (0.0002)	0.0002 (0.0003)	
$R_{t-15}1_{News}$	-0.1257 (0.1901)	0.0086 (0.2394)	-0.3805 (0.2436)	$AOF_t 1_{News}$	-0.0002 (0.0003)	-0.0002 (0.0003)	0.0011 (0.0007)	
$R_{t-15}1_{Post}$	0.1431 (0.1568)	0.1431 (0.1597)	0.1431 (0.1719)	$AOF_t 1_{Post}$	$0.0001 \\ (0.0004)$	$0.0001 \\ (0.0004)$	0.0001 (0.0004)	
$R_{t-15}1_{News}1_{Post}$	0.3092 (0.3822)	-0.1574 (0.4677)	2.9951 (1.9775)	$AOF_t 1_{News} 1_{Post}$	-0.0001 (0.0005)	$0.0006 \\ (0.0008)$	0.0021 (0.0040)	
$S_{0,t}$	0.0182 (0.0469)	0.0182 (0.0478)	0.0182 (0.0514)	AOI_t	0.0064 (0.0051)	0.0064 (0.0052)	0.0064 (0.0055)	
$S_{0,t}1_{News}$	-0.0610 (0.0750)	-0.0187 (0.0844)	-0.1541 (0.1263)	$ ext{AOI}_t 1_{News}$	-0.0097 (0.0070)	-0.0084 (0.0075)	-0.0110 (0.0104)	
$S_{0,t}1_{Post}$	0.0401 (0.0780)	0.0401 (0.0794)	0.0401 (0.0855)	$ ext{AOI}_t 1_{Post}$	-0.0143 (0.0095)	-0.0143 (0.0097)	-0.0143 (0.0104)	
$S_{0,t}1_{News}1_{Post}$	0.2664*** (0.1157)	0.1591 (0.1142)	0.2520^* (0.1525)	$ ext{AOI}_t 1_{News} 1_{Post}$	0.0166 (0.0162)	-0.0109 (0.0263)	-0.0464 (0.0501)	
$D_{0,t}$	-0.0065 (0.0045)	-0.0065 (0.0046)	-0.0065 (0.0049)	$\operatorname{Sur}_t 1_{Post}$	-0.0070 (0.0220)	0.0072 (0.0267)	-0.1508 (0.1226)	
$\mathrm{D}_{0,t}1_{News}$	0.0069 (0.0054)	0.0056 (0.0058)	0.0082 (0.0067)	Adjusted R ²	0.2688	0.3337	0.5597	
$\mathbf{D}_{0,t}1_{Post}$	$0.0100 \\ (0.0073)$	$0.0100 \\ (0.0074)$	0.0100 (0.0080)					
$\mathbf{D}_{0,t}1_{News}1_{Post}$	-0.0184 (0.0138)	0.0140 (0.0252)	0.0416 (0.0526)					

include other announcements outside the 8:30 interval? Does this assumption change my results, or can I still assert that it is the arrival of private information that affects volatility?

The pre-scheduled macroeconomic announcements are released at different times. The list of all news and their arrival times are reported in Tables 2.1 and 2.2. According to the reported list, a majority of news releases are at 8:30 and the rest are at different times. Depending on the country and the nature of the macroeconomic indicator, I have news at 7:00, 8:30, 9:15, 10:00, and 14:15. To examine the effect of news as a whole, I include all the announcements outside the 8:30 interval (as well as the 8:30 announcements). As noted previously in Section 2.2.2, I remove overlapping announcements. I set up my regression model using equation (2.9) to examine realised variance as my main volatility measure through the Full, Relatively Important, and Most Important samples. Table 2.22 shows the results of using the Realised Variance measure for three different samples when I examine the impact of announcements at times other than 8:30. The results of using the *Realised Variance* measure for the *Full* sample reveals significant results for spread at the best quote (S_0) , depth at the best quote (D_0) , and absolute order imbalance (AOI). The spread at the best quote is significant at the 1% level while depth at the best quote and absolute order imbalance are significant at the 5% level. The Relatively Important sample reveals significant results for spread at the best quote (S_0) and depth at the second best quote (D_1) at the 10% and 5% level respectively. The Most Important sample shows significant results for depth at the best quote (D_0) at the 10% level.

Overall, these results are consistent with my findings for the 8:30 announcement using the full sample. The effect of private information shocks appear in spread at the best quote, depth at the best quote, and absolute order imbalance for the Full sample but for the $Relatively\ Important$ samples it comes through the spread at the best quote and depth behind the best

TABLE 2.22: The Effect of News at Times other than 8:30 on Realised Variance.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 330 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Explanatory	Samples			Explanatory	Samples		
Variables	Full	RI	MI	Variables	Full	RI	MI
Intercept	0.0531** (0.0241)	0.0531** (0.0243)	0.0494** (0.0230)	$\mathrm{D}_{1,t}$	0.0014 (0.0011)	0.0014 (0.0012)	0.0007 (0.0011)
1_{News}	0.0550^* (0.0342)	0.0580 (0.0384)	0.0763 (0.0515)	$\mathrm{D}_{1,t}1_{News}$	-0.0014 (0.0019)	-0.0040^{**} (0.0021)	-0.0042^* (0.0026)
1_{Post}	-0.0397 (0.0327)	-0.0397 (0.0329)	-0.0461 (0.0342)	$\mathrm{D}_{1,t}1_{Post}$	-0.0001 (0.0015)	-0.0001 (0.0015)	$0.0001 \\ (0.0014)$
$1_{News}1_{Post}$	0.0366 (0.0621)	0.0007 (0.0626)	-0.0356 (0.1369)	$\mathbf{D}_{1,t}1_{News}1_{Post}$	0.0003 (0.0040)	0.0070** (0.0033)	0.0032 (0.0043)
RV_{t-15}	0.3369*** (0.0958)	0.3369*** (0.0963)	0.3191*** (0.1006)	AOF_t	0.0002 (0.0003)	0.0002 (0.0003)	0.0002 (0.0003)
$\mathrm{RV}_{t-15}1_{News}$	-0.1654 (0.1458)	-0.0692 (0.1692)	0.2458 (0.2485)	$AOF_t 1_{News}$	$0.0001 \\ (0.0004)$	$0.0005 \\ (0.0007)$	0.0002 (0.0007)
$RV_{t-15}1_{Post}$	0.1401 (0.2522)	0.1401 (0.2536)	0.3923 (0.3287)	$AOF_t 1_{Post}$	$0.0001 \\ (0.0005)$	$0.0001 \\ (0.0005)$	-0.0001 (0.0006)
$RV_{t-15}1_{News}1_{Post}$	0.0374 (0.3473)	-0.3207 (0.3387)	-0.1088 (0.6561)	$AOF_t 1_{News} 1_{Post}$	-0.0004 (0.0006)	-0.0006 (0.0008)	$0.0001 \\ (0.0015)$
$S_{0,t}$	-0.0299 (0.0558)	-0.0299 (0.0561)	-0.0075 (0.0473)	AOI_t	0.0041 (0.0029)	0.0041 (0.0030)	$0.0005 \ (0.0033)$
$S_{0,t}1_{News}$	-0.0154 (0.0894)	0.0359 (0.0990)	0.0761 (0.1178)	$\mathrm{AOI}_t 1_{News}$	-0.0058 (0.0047)	-0.0044 (0.0058)	-0.0065 (0.0080)
$S_{0,t}1_{Post}$	0.0187 (0.0638)	0.0187 (0.0641)	0.0180 (0.0576)	AOI_t1_{Post}	$-0.0119^{**} \ (0.0055)$	-0.0119** (0.0055)	-0.0054 (0.0061)
$S_{0,t}1_{News}1_{Post}$	0.2856*** (0.1159)	0.1856^* (0.1122)	0.1149 (0.1342)	$AOI_t 1_{News} 1_{Post}$	-0.0157^{**} (0.0079)	-0.0023 (0.0137)	-0.0682 (0.0486)
$D_{0,t}$	-0.0028 (0.0023)	-0.0028 (0.0023)	-0.0012 (0.0024)	Sur_t1_{Post}	0.0252 (0.0238)	0.0158 (0.0192)	-0.0229 (0.0428)
$\mathrm{D}_{0,t}1_{News}$	0.0023 (0.0037)	0.0007 (0.0042)	-0.0009 (0.0047)	Adjusted R ²	0.2957	0.3380	0.5302
$\mathbf{D}_{0,t}1_{Post}$	0.0089** (0.0047)	0.0089** (0.0047)	0.0048 (0.0050)				
$\mathbf{D}_{0,t}1_{News}1_{Post}$	-0.0117** (0.0061)	0.0014 (0.0099)	0.0655^* (0.0385)				

quote. The *Most Important* sample has significant results only for depth at the best quote. I reiterate that rises in spread, rises in the depth at the second best quote, falls in the depth at the best quote, or rises in the absolute order imbalance, are all signs of illiquidity. The only significant coefficient that is contrary to this is the effect of depth at the best quote with the *Most Important* sample. Here, a rise in liquidity seems to lead to greater volatility. I conclude that inclusion of other announcement times does not materially change my findings. If anything, I obtain stronger results from this enlarged sample.

2.4.7 The Magnitude of Liquidity Variable Changes

I have shown that the sensitivity of volatility to liquidity measures changes in response to macroeconomic news. I now turn my attention to the level of my liquidity measures: are they large enough to make this sensitivity change economically significant? In this regard, I examine the difference between news days and no-news days during the 15-minute interval before and after 8:30. To have a closer look, I split each 15-minute interval before and after 8:30 into three 5-minute intervals.

Table 2.23 shows that the main liquidity measures with significant results are the bid-ask spread at the most competitive quote (S_0), absolute order flow (AOF), and absolute order imbalance (AOI). My sensitivity regressions (see Sections 2.3.2 and 2.4.6) show that volatility has a heightened sensitivity to S_0 and AOI following announcements (but not AOF), so I focus my attention on these variables. The difference between news days and no-news days for spread has a large and significant increase in the first 5-minute interval after the arrival of macroeconomic announcements. This difference becomes larger as I move from the full sample to the *Most Important* sample. The absolute order imbalance also shows an increase in difference between news days and no-news days. However,

TABLE 2.23: The Magnitude of Changes after Announcements.

Each number represents the average difference in the liquidity measure, at the indicated 5-minute window, between news days and no-news days. My Liquidity measures are: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the measures can be found in Section 2.3. There are 330 observations in full announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Interval	Full	RI	MI	Full	RI	MI	Full	RI	MI
		S_0			D_0			D_1	
-3	0.0291 (0.0297)	0.0437 (0.0317)	-0.0075 (0.0344)	-0.2146 (0.8399)	-0.3136 (0.8962)	-0.0041 (1.2749)	-0.3267 (1.0221)	-0.2380 (1.0760)	0.0827 (1.1179)
-2	-0.0228 (0.0342)	-0.0265 (0.0328)	0.0176 (0.0546)	-0.1194 (1.0828)	0.1037 (1.3666)	-1.6457 (0.9916)	-2.2088 (2.5917)	-3.0326 (2.5938)	-3.2965 (2.9457)
-1	-0.0141 (0.0259)	-0.0166 (0.0291)	-0.0501 (0.0350)	-0.0944 (0.5835)	$0.4205 \\ (0.6963)$	$1.3004 \\ (1.1246)$	-0.8729 (0.8213)	-0.5112 (0.9649)	0.1271 (1.8823)
+1	0.1694** (0.0534)	* 0.2204** (0.0789)	* 0.5633** (0.2563)	-1.0817 (1.1787)	-0.9523 (1.2520)	-0.9589 (1.4077)	-0.2982 (0.8469)	-0.5634 (0.8731)	0.9474 (1.8377)
+2	0.0046 (0.0251)	-0.0098 (0.0264)	-0.0208 (0.0329)	-1.4324 (1.4172)	-1.4651 (1.4659)	-2.7920 (1.4476)	$1.7437 \\ (1.4175)$	3.0385 (1.9906)	2.6560 (1.5554)
+3	0.0069 (0.0251)	0.0116 (0.0264)	0.0513 (0.0329)	$0.5272 \\ (0.7913)$	0.5447 (0.8621)	0.4780 (0.9089)	0.2548 (0.9592)	0.7724 (1.2072)	-0.4609 (1.5567)
Interval		AOF			AOI				
-3	17.9241** (5.8585)	* 4.4140 (4.2504)	-3.6583 (3.8985)	-5.1930** (0.6511)	$^*-0.5847$ (0.6878)	-0.2561 (0.8449)			
-2	20.0236** (9.4313)	* 2.4444 (5.8015)	-7.5893 (5.2297)	0.3894 (0.8274)	0.5826 (1.0945)	-0.5931 (0.7062)			
-1	26.9598** (4.5272)	* 6.9910** (3.0745)	* -0.7345 (3.0789)	-0.1183 (0.4995)	0.1731 (0.5434)	$1.5571 \\ (1.0266)$			
+1	27.3048** (4.9454)	* 9.6451** (2.5282)	* 0.3968 (2.3868)	-0.1183 (0.8454)	0.1731 (0.9551)	$1.5571 \\ (1.0839)$			
+2	44.1696** (7.0332)	* 20.0446** (4.1386)	* 16.6754** (6.6238)	-0.9489 (1.1922)	-0.6352 (1.2472)	-1.6418 (1.2236)			
+3	34.4161** (8.0706)	* 14.4881** (6.0167)	* 8.6071 (7.5121)	0.2948** (0.0297)	* 0.1545 (0.7638)	-0.2122 (0.8599)			

the large increase happens in the third 5-minute interval after the arrival of news. This liquidity measure does not follow a specific pattern compared to spread. Overall, my conclusion is that volatility not only shows a heightened sensitivity to spread following the arrival of news but also spread has (on average) a significant increase, which makes this sensitivity important. This pattern is not so clear and significant for other liquidity measures. This reinforces my conclusion from my main analyses that spread is the most important microstructure measure for capturing private information.

2.4.8 Risk Averse/Seeking Trader in the Market

In this part, I examine the market condition with different level of private information for a risk averse/seeking trader. I calculate the returns, the standard deviation of returns, and spread in each one-minute over the 15-minute interval after the 8:30 on news days and no-news days. The spread median is used as a threshold to determine low/high private information level. I evaluate the market condition for a risk averse/seeking trader using the *Full*, *Relatively Important*, and *Most Important* samples. Table 2.24 presents the results for the Canadian FX market in year 2005. Panel A, shows the Full sample with 122 observations on news days and 43 observations on no-news days. On average, returns and standard deviation of returns increase from 0.71% and 11.61% to 0.81% and 14.01% respectively on news days as the level of private information goes up. Similarly, on average, returns and standard deviation of returns increase from -0.11% and 7.44% to 0.13% and 7.53% respectively on no-news days as the level of private information increases. Results show that a higher level of private information (more than the spread median) provides higher return with higher standard deviation and a lower level of private information (less than the spread median) provides less return with less standard deviation.

Panel B presents the Relatively Important news with 81 and 42 observations on news days and no-news days respectively. On average, returns and standard deviation of returns increase from -0.13% and 12.65% to 0.74% and 13.91% respectively on news days as the level of private information goes up. Similarly, on average, returns and standard deviation of returns increase from -0.65% and 8.01% to 0.82% and 7.89% respectively on no-news days as the level of private information increases. Finally, Panel C presents the Most Important news with 54 and 42 observations on news days and no-news days respectively. On average, returns and standard deviation of returns increase from 0.09% and 14.14% to 1.74% and 14.92% respectively on news days as the level of private information goes up. Similarly, on average, returns and standard deviation of returns increase from -0.65% and 8.01% to 0.82% and 7.89% respectively on no-news days as the level of private information increases. I use the t-test to evaluate whether there is a significant difference between high and low sample on news days and no-news days as well as across groups. Results show no significant results for any of comparisons that might be due to small sample and having close returns on average.

Overall, a risk averse trader might trade on days with less level of private information and a risk seeking trader might trade on days with higher level of private information. Results hold the same across all three samples in Panels A, B and C on news days and no-news days.

2.4.9 Realised Variance in Economic Samples

In this section, I split all the macroeconomic announcements (i.e. Tables 2.1 and 2.2) into 4 different samples based on their economic function: output, prices, employment, and rate decision. The time of announcements varies within each sample and includes 8:30 and times other than it. Table 2.25 presents all pre-scheduled macroeconomic announcements in US. Simi-

TABLE 2.24: Market Condition for Risk Averse/Seeking Trader.

This table presents the return, its standard deviation, and spread for each one-minute over a 15-minute interval after 8:30 on news days and no-news days. The median of spread is used as a threshold to determine low/high level of private information in the market. Panels A, B, and C include 122, 81, and 54 observations on news days and 43, 42, and 42 no-news days respectively.

Sample	News Day	s Information	No-News Days Information		
oumpre	Low	High	Low	High	
Panel A: Full Sample					
Avg_Return	0.0071	0.0081	-0.0011	0.0013	
Min_Return	-0.0938	-0.1301	-0.0580	-0.0347	
Max_Return	0.1124	0.1137	0.0380	0.0405	
Avg_std	0.1161	0.1401	0.0744	0.0753	
Min_std	0.0000	0.0000	0.0000	0.0302	
Max_std	0.3105	0.3907	0.1287	0.1306	
Panel B: Relatively Important Sample					
Avg_Return	-0.0013	0.0074	-0.0065	0.0082	
Min_Return	-0.0938	0.0000	0.0322	-0.0297	
Max_Return	0.1124	0.3067	0.0380	0.0405	
Avg_std	0.1265	0.1391	0.0801	0.0789	
Min_std	0.0000	0.0000	0.0322	0.0302	
Max_std	0.3229	0.3067	0.1028	0.1306	
Panel C: Most Important Sample					
Avg_Return	0.0009	0.0147	-0.0065	0.0082	
Min_Return	-0.0752	-0.1301	-0.0580	-0.0297	
Max_Return	0.0757	0.1137	0.0380	0.0405	
Avg_std	0.1414	0.1492	0.0801	0.0789	
Min_std	0.0000	0.0000	0.0322	0.0302	
Max_std	0.3229	0.3067	0.1028	0.1306	

TABLE 2.25: Pre-scheduled Macroeconomic Announcements Survey in 4 Groups in US.

This table provides summary statistics for macroeconomic announcements according to different economic groups in the US in the year 2005.

Event	Source	Obs.	Time	Min	Max	Mean	Std.
Panel A: Output							
Bloomberg Consumer Confidence	В	10	17:00	-22	-8	-14.5	4.7668
Building Permits	CB	12	08:30	1985	2160K	2077.9K	49.8K
Business Inventories	CB	12	08:30	0.0010	0.0090	0.0038	0.0024
Capacity Utilization	FR	12	09:15	0.7890	0.8030	0.7953	0.0035
Chicago Purchasing Manager	CBB	12	10:00	52	62.500	58.617	3.289
Conf. Board Consumer Confidence	CB	12	10:00	88	106.30	99.050	5.706
Construction Spending MoM	CB	12	10:00	0.0030	0.0070	0.0049	0.0011
Consumer Credit	FR	12	15:00	4.1B	10B	6.4B	1.6B
Durable Goods Orders	CB	12	08:30	-0.0150	0.0150	0.0034	0.0111
Empire Manufacturing	FR	12	08:30	1	25	15.983	6.1573
Factory Orders	CB	12	10:00	-0.0230	0.0300	0.0058	0.0153
GDP Annualised QoQ	BEA	12	08:30	0.0330	0.0430	0.0368	0.0029
Housing Starts	CB	12	08:30	1903K	2090K	2013.6K	56.2K
ISM Manufacturing	ISM	12	10:00	51.400	58.500	55.383	2.4778
Leading Index	CB	12	10:00	-0.0050	0.0080	0.0003	0.0040
Monthly Budget Statement	T	12	14:00	-100B	60B	-26.9B	50.8B
NAHB Housing Market Index	NAHB	12	13:00	61	71	67.833	2.9797
New Home Sales	CB	12	10:00	1125K	1350K	1253.4K	77.5K
Personal Consumption	BEA	10	08:30	0.0300	0.0430	0.0359	0.0042
Personal Income	BEA	12	08:30	-0.0260	0.0340	0.0041	0.0128
Personal Spending	BEA	12	08:30	-0.0020	0.0100	0.0045	0.0036
Philadelphia Fed Business Outlook	PFR	12	12:00	10	25	14.733	4.6582
Retail Sales Advance MoM	CB	12	08:30	-0.0140	0.0210	0.0038	0.0093
Trade Balance	CB	12	09:30	-62.9B	-54B	-58.7B	2.6B
Wholesale Inventories MoM	СВ	12	10:00	0.0030	0.0090	0.0056	0.0017
Panel B: Prices							
CPI MoM	BLS	12	08:30	-0.0040	0.0090	0.0025	0.0033
Current Account Balance	BEA	4	08:30	-205B	-183B	-192.75B	9.2B
Import Price Index MoM	BLS	12	08:30	-0.0050	0.0140	0.0047	0.0068
Industrial Production Price MoM	FR	12	09:15	-0.004	0.01	0.0034	0.0031
PCE Core YoY	BEA	11	08:30	0.0150	0.0190	0.0174	0.0016
Panel C: Employment							
Change in Nonfarm Payrolls	BLS	12	08:30	-150K	225K	159.3K	101.1K
Continuing Claims	DL	50	08:30	2573K	2939K	2664.7K	93.2K
Employment Cost Index	BLS	4	08:30	0.0080	0.0100	0.0085	0.0010
Initial Jobless Claims	DL	52	08:30	309K	450K	329.8K	24.5K
Nonfarm Productivity	BLS	8	08:30	0.0150	0.0450	0.0238	0.0100
Unemployment Rate	BLS	12	08:30	0.0500	0.0540	0.0516	0.0015
Panel D: Decision Rate							
	FR	8	14:15	0.0250	0.0425	0.0338	0.0061

TABLE 2.26: Pre-scheduled Macroeconomic Announcements Survey in Canada.

This table provides summary statistics for macroeconomic announcements according to 4 groups in the Canada in the year 2005.

Event	Source	Obs.	Time	Min	Max	Mean	Std.
Panel A: Output							
Building Permits MoM	STCA	12	08:30	-0.0500	0.0400	-0.0082	0.0270
Housing Starts	STCA	12	08:30	215K	230K	224.5K	4.48K
Int'l Merchandise Trade	STCA	12	08:30	4.3B	6.9B	5.19B	0.787B
Manufacturing Sales MoM	STCA	12	08:30	-0.0060	0.0100	0.0038	0.0046
Quarterly GDP Annualised	STCA	4	08:30	0.0170	0.0360	0.0263	0.0078
Retail Sales MoM	STCA	12	08:30	-0.0050	0.0140	0.0031	0.0065
Wholesale Trade Sales MoM	STCA	12	08:30	0.0020	0.0070	0.0050	0.0015
Panel B: Prices							
CPI YoY	STCA	12	07:00	0.0180	0.0330	0.0233	0.0046
Capacity Utilization Rate	STCA	4	08:30	0.8560	0.8720	0.8625	0.0072
Current Account Balance	STCA	4	08:30	3.4B	9B	6.14B	2.51B
Industrial Product Price MoM	STCA	12	08:30	-0.0100	0.0090	0.0030	0.0047
New Housing Price Index MoM	STCA	12	08:30	0.0030	0.0050	0.0039	0.0007
Raw Materials Price Index MoM	STCA	12	08:30	-0.0200	0.0400	0.0148	0.0201
Panel C: Employment							
Labor Productivity QoQ	STCA	4	08:30	0.0010	0.0050	0.0030	0.0016
Unemployment Rate	STCA	12	07:00	0.0660	0.0730	0.0687	0.0019
Panel D: Decision Rate Bank of Canada Rate Decision	BC	8	09:00	0.0250	0.0325	0.0269	0.0029

larly, I split pre-scheduled macroeconomic announcements in 4 Groups in the Canada. Table 2.26 presents the list of each group. I run the main regression model in equation (2.9) for the RV model for each economic group and explain the results in the following.

The first sample includes output announcements, which is presented in Panel A, Tables 2.25 and 2.26. Table 2.27 presents results of realised variance using output sample. The announcements of this sample come to the market at different times and overall there are 433 observations for the regression model. Spread at the best quote and Depth at the second

TABLE 2.27: Realised Variance Using the Output Sample.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 433 observations in output announcements sample. The superscripts ***, ***, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Li	iquidity Measure	(Output Annous	ncements Sample	·)
variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI
Intercept	0.0824*** (0.0236)	0.0667*** (0.0174)	0.0604*** (0.0150)	0.0575*** (0.0126)	0.0632*** (0.0126)
1_{News}	0.0234 (0.0318)	0.0394* (0.0227)	0.0356 (0.0231)	0.0302* (0.0177)	0.0380** (0.0178)
1_{Post}	-0.0175 (0.0352)	-0.0134 (0.0274)	-0.0251 (0.0274)	-0.0002 (0.0198)	0.0009 (0.0228)
$1_{News}1_{Post}$	-0.0613 (0.0719)	0.0890^* (0.0535)	0.1425** (0.0620)	0.0783^* (0.0467)	$0.0705 \\ (0.0478)$
RV_{t-15}	0.3486*** (0.1006)	0.3481*** (0.1031)	0.3544*** (0.1055)	0.3245*** (0.1006)	0.3441*** (0.1033)
$\mathrm{RV}_{t-15}1_{News}$	-0.0931 (0.1455)	-0.0894 (0.1495)	-0.1079 (0.1460)	-0.1090 (0.1483)	-0.0885 (0.1469)
$\mathrm{RV}_{t-15}1_{Post}$	0.1409 (0.2436)	0.1339 (0.2453)	0.1606 (0.2475)	0.1316 (0.2911)	0.1403 (0.2451)
$\text{RV}_{t-15}1_{News}1_{Post}$	-0.0988 (0.3236)	0.0598 (0.3515)	0.0396 (0.3366)	0.0567 (0.3834)	0.0557 (0.3501)
X_t	-0.0592 (0.0784)	0.0004 (0.0018)	0.0009 (0.0010)	$0.0005 \\ (0.0004)$	0.0021 (0.0017)
$X_t 1_{News}$	0.0149 (0.1003)	-0.0024 (0.0025)	-0.0010 (0.0019)	-0.0002 (0.0005)	$-0.0050* \\ (0.0027)$
$X_t 1_{Post}$	0.0396 (0.1003)	0.0009 (0.0026)	0.0018 (0.0021)	-0.0003 (0.0006)	-0.0032 (0.0029)
$X_t 1_{News} 1_{Post}$	0.4007^* (0.2408)	-0.0018 (0.0035)	-0.0083^{**} (0.0039)	-0.0002 (0.0008)	0.0019 (0.0043)
$\operatorname{Sur}_t 1_{Post}$	0.0279 (0.0234)	0.0211 (0.0255)	0.0172 (0.0261)	0.0273 (0.0246)	$0.0220 \\ (0.0251)$
Adjusted R ²	0.2416	0.1354	0.1450	0.1301	0.1347

best quote as proxies for private information shock are significant at the 10% and 5% level respectively. Surprise as a proxy for public information shock is not significant. Hence, with the arrival of news in the output sample, spread at the best quote and depth at the second best quote are the main proxies that signal volatility increasing illiquidity in the market.

The second sample includes price announcements, which is presented in Panel B, Tables 2.25 and 2.26. Table 2.28 presents results of the regression model for this sample with 200 observations. depth at the best quote and absolute order flow as proxies of private information shocks are significant at the 10% and 1% level respectively. Surprise as a proxy for public information shock is not significant. Hence, with the arrival of macroeconomic announcements, the private information shocks cause illiquidity for the price sample and results in higher volatility.

The third sample includes employment announcements, which is presented in Panel C, Tables 2.25 and 2.26. Table 2.29 presents the results of regression model for this sample that includes 189 observations. For this sample, depth at the best quote and absolute order flow as proxies for private information shocks are significant at the 10% level. Surprise as the public information shock is significant for absolute order imbalance at the 10% level.

Finally, the last sample includes rate decision sample announcements, which is presented in Panel D, Tables 2.25 and 2.26. Table 2.30 reports the results for FOMC sample with 58 observations. Private information shocks are significant through depth at the second best quote, absolute order flow, and absolute order imbalance at the 1%, 5%, and 5% level respectively. Public information shocks are not significant for the decision rate sample. Hence, when FOMC announcements come to the market, liquidity dries up through private information and causes higher volatility that can be captured through liquidity variables.

TABLE 2.28: Realised Variance Using the Prices Sample.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 200 observations in Prices announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	I	iquidity Measur	e (Prices Announ	cements Sample)	
variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI
Intercept	0.0722*** (0.0212)	0.0601*** (0.0148)	0.0562*** (0.0124)	0.0575*** (0.0109)	0.0580*** (0.0113)
1_{News}	0.0467 (0.0394)	0.0743*** (0.0281)	0.0881** (0.0361)	0.0543** (0.0232)	0.0763^{***} (0.0225)
1_{Post}	-0.0405 (0.0268)	0.0020 (0.0216)	-0.0080 (0.0180)	-0.0156 (0.0154)	$0.0001 \\ (0.0177)$
$1_{News}1_{Post}$	0.0678 (0.1113)	0.2518** (0.1302)	0.2068 (0.1560)	0.2157^* (0.1155)	0.1802 (0.1152)
RV_{t-15}	0.3605*** (0.0480)	0.3621*** (0.0510)	0.3606*** (0.0527)	0.3635*** (0.0526)	0.3644*** (0.0511)
$\mathrm{RV}_{t-15}1_{News}$	-0.3336* (0.1883)	-0.3513^* (0.1925)	-0.3696^* (0.2054)	-0.3997^{**} (0.1837)	-0.3903^{**} (0.1871)
$\mathrm{RV}_{t-15}1_{Post}$	0.0136 (0.1473)	0.0132 (0.1470)	-0.0032 (0.1443)	-0.1631 (0.1574)	$0.0001 \\ (0.1471)$
$\text{RV}_{t-15}1_{News}1_{Post}$	$0.6500 \\ (0.6002)$	1.1617** (0.5841)	0.6871 (0.6846)	1.1982** (0.5631)	1.0895^* (0.5938)
X_t	-0.0512 (0.0561)	-0.0007 (0.0016)	0.0000 (0.0009)	-0.0001 (0.0002)	-0.0009 (0.0018)
$X_t 1_{News}$	0.0464 (0.0932)	-0.0024 (0.0031)	-0.0032 (0.0025)	0.0004* (0.0002)	-0.0056 (0.0053)
$X_t 1_{Post}$	0.0923 (0.0667)	-0.0025 (0.0025)	-0.0001 (0.0011)	0.0010** (0.0004)	-0.0043 (0.0028)
$X_t 1_{News} 1_{Post}$	0.2277 (0.2357)	-0.0211^* (0.0120)	-0.0014 (0.0096)	-0.0029^{***} (0.0009)	-0.0151 (0.0176)
Sur_t1_{Post}	-0.0017 (0.0442)	-0.0391 (0.0586)	-0.0315 (0.0625)	-0.0060 (0.0560)	-0.0329 (0.0585)
Adjusted R ²	0.3243	0.3104	0.2584	0.2907	0.2844

TABLE 2.29: Realised Variance Using the Employment Sample.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 189 observations in employment announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liqu	ıidity Measure (I	Employment Anr	nouncements San	ıple)	
VIII 103	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI	
Intercept	0.0722*** (0.0212)	0.0601*** (0.0148)	0.0562*** (0.0125)	0.0575*** (0.0109)	0.0580*** (0.0114)	
1_{News}	-0.0139 (0.0318)	0.0468 (0.0295)	0.0170 (0.0271)	0.0298 (0.0207)	0.0404^* (0.0243)	
1_{Post}	-0.0405 (0.0268)	0.0020 (0.0216)	-0.0080 (0.0180)	-0.0156 (0.0154)	0.0001 (0.0177)	
$1_{News}1_{Post}$	0.0992 (0.0788)	-0.0566 (0.0945)	-0.0029 (0.1120)	0.0746 (0.1016)	0.0341 (0.1065)	
RV_{t-15}	0.3605*** (0.0481)	0.3621*** (0.0511)	0.3606*** (0.0528)	0.3635*** (0.0527)	0.3644*** (0.0512)	
$\mathrm{RV}_{t-15}1_{News}$	-0.0537 (0.1860)	-0.0884 (0.1927)	-0.0065 (0.1944)	-0.0993 (0.2265)	-0.0826 (0.1945)	
$\mathrm{RV}_{t-15}1_{Post}$	0.0136 (0.1476)	0.0132 (0.1472)	-0.0032 (0.1446)	-0.1631 (0.1577)	0.0001 (0.1473)	
$\text{RV}_{t-15}1_{News}1_{Post}$	-0.1195 (0.5201)	-0.0194 (0.5512)	-0.0949 (0.6501)	0.3320 (0.6052)	$0.1169 \\ (0.5525)$	
X_t	-0.0512 (0.0562)	-0.0007 (0.0016)	$0.0000 \\ (0.0009)$	-0.0001 (0.0002)	-0.0009 (0.0018)	
$X_t 1_{News}$	0.1462** (0.0741)	-0.0019 (0.0030)	0.0017 (0.0024)	0.0002 (0.0003)	-0.0018 (0.0043)	
$X_t 1_{Post}$	0.0923 (0.0668)	-0.0025 (0.0025)	-0.0001 (0.0011)	0.0010^* (0.0004)	-0.0043 (0.0028)	
$X_t 1_{News} 1_{Post}$	0.0364 (0.0918)	0.0299* (0.0171)	0.0142 (0.0190)	-0.0011^* (0.0007)	0.0188 (0.0171)	
$\operatorname{Sur}_t 1_{Post}$	0.0128 (0.0290)	0.0617 (0.0432)	0.0758 (0.0482)	0.0734 (0.0502)	$0.0684* \\ (0.0425)$	
Adjusted R ²	0.4646	0.3131	0.2893	0.2561	0.2646	

TABLE 2.30: Realised Variance Using the Rate Decision Sample.

Results of regression of dependent variable (RV) realised variance on liquidity measures: S_0 is spread, D_0 is the depth at the best quote, D_1 is the depth at the second best quote, AOF is the absolute order flow, AOI is the absolute order imbalance. Descriptions of the calculation of the variables can be found in Section 2.3. There are 58 observations in rate decision announcements sample. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively. The significance is assessed using Newey-West standard errors.

Variables	Liqu	iidity Measure (I	Rate Decision An	nouncements Sa	mple)	
Variables	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI	
Intercept	-0.0013^{***} (0.0214)	0.0261*** (0.0096)	-0.0119^{***} (0.0179)	0.0306*** (0.0096)	0.0316*** (0.0096)	
1_{News}	0.0589 (0.0488)	0.0234 (0.0390)	0.1291*** (0.0483)	0.0243 (0.0450)	0.0347 (0.0364)	
1_{Post}	-0.0008 (0.0247)	-0.0407^{***} (0.0142)	-0.0060 (0.0221)	-0.0429^{***} (0.0146)	-0.0463^{***} (0.0137)	
$1_{News}1_{Post}$	0.5323 (0.1355)	0.3491* (0.1826)	0.2336 (0.1808)	0.0648 (0.1922)	0.3611** (0.1536)	
RV_{t-15}	0.3157 (0.2588)	0.2596 (0.2531)	0.1875 (0.1406)	0.3292 (0.2515)	0.2425 (0.2518)	
$\mathrm{RV}_{t-15}1_{News}$	-0.2967 (0.5792)	-1.1330 (1.3446)	-0.3398 (0.4533)	-0.3294 (0.7130)	-0.5327 (0.5899)	
$\mathrm{RV}_{t-15}1_{Post}$	0.9124 (0.3629)	0.9435*** (0.3724)	1.0246*** (0.3085)	0.8869** (0.3702)	0.9605*** (0.3716)	
$\text{RV}_{t-15}1_{News}1_{Post}$	$ \begin{array}{c} 1.1625 \\ (2.4138) \end{array} $	-0.3855 (1.9571)	-1.0737 (1.4010)	0.3206 (1.4569)	-1.0814 (1.5017)	
X_t	0.0852 (0.0583)	-0.0004 (0.0028)	0.0069^* (0.0035)	-0.0005 (0.0005)	-0.0042 (0.0036)	
$X_t 1_{News}$	0.0959 (0.0000)	0.0203 (0.0179)	-0.0096** (0.0041)	0.0021 (0.0022)	0.0570*** (0.0209)	
$X_t 1_{Post}$	-0.1335 (0.0677)	0.0005 (0.0029)	-0.0065^* (0.0036)	0.0004 (0.0006)	0.0042 (0.0037)	
$X_t 1_{News} 1_{Post}$	-0.6887 (0.4131)	-0.0150 (0.0185)	0.0124*** (0.0048)	0.0112** (0.0057)	-0.0503^{**} (0.0217)	
$\operatorname{Sur}_t 1_{Post}$	0.0001 (0.0000)	0.0001 (0.0000)	0.0001 (0.0000)	0.0001 (0.0000)	0.0001 (0.0000)	
Adjusted R ²	0.8060	0.7527	0.7902	0.8373	0.7607	

2.5 Summary and Conclusion

In this section, I present a summary and conclusion of all regression analysis. First, I demonstrate the significant results in a summary table and how they change with different subsamples. Then, I express an overall conclusion of my study.

2.5.1 Summary of All Regression Analysis

The main assumption of my study is that with the arrival of macroeconomic announcements the level of asymmetric information increases either through public information shocks or private information shocks. The liquidity in the market dries up due to uncertainty and different levels of information among traders, which results in a volatile market. Anyone with more information is prone to make profit whereas traders with less information are exposed to lose money. I set up regression model (2.9) (in Section 2.3.2) to analyse the impact of public and private information shocks on realised variance. Surprise is a proxy for public information shocks and a range of liquidity variables (Spread at the best quote (S_0), depth at the best quote (S_0), depth at the second best quote (S_0), as a proxy for private information shocks.

Table 2.31 presents the summary of significant results of all regressions. The main regression examines the impact of explanatory variables on realised variance through the *Full*, *Relatively Important*, and *Most Important* samples. Spread is the only significant variable in the Full sample at the 5% level and other samples do not reveal any significant results. The reason might be due to relatively or highly discussed information among traders. In my robustness tests, I remove the lag of Realised Variance from the regression model and results hold the same. When I remove the sur-

prise variable from the regression model the significance level of spread increases from the 5% to 1% level. However, having the impact of surprise variable with positive or negative values does not change the primary results in the Full sample.

I evaluate the impact of macroeconomic announcements in the US and Canadian market separately and results hold the same for the US but no significant results are found for the Canadian announcements. Absolute return and modified absolute return are two alternative measures of volatility in the market. Using absolute return in the Full sample shows that spread at the best quote and absolute order flow are significant variables at the 10% and 5% level respectively. Absolute return in the Relatively Important sample shows that spread at the best quote is significant at the 10% level. Using modified absolute return in the Full sample shows that depth at the second best quote and absolute order flow are significant at the 1% and 5% level. Modified absolute return in the relatively important sample shows that depth at the second best quote, and absolute order flow are significant at the 1% and 10% level. The Most Important sample does not reveal any significant results for any sample, which might be due to these announcements being highly discussed among traders. The private information signal reveals through spread when I use absolute return as an alternative measure of volatility whereas it appears through depth at the second best quote and absolute order flow when I use modified absolute return.

Testing the impact of all variables jointly shows that spread at the best quote and depth at the second best quote are significant at the 10% level. Comparing all the results shows that the private information shocks is the main source of volatility compared with public information shocks. The sign of private information shocks reveals through different liquidity variables in different tests and we can not present a particular candidate that holds the same throughout all regression analysis. Having said

Table 2.31: Summary of Significant Results.

This table reports the summary of significant results for main regression model and robustness tests. Liquidity variables as a sign of private information signal through different variables whereas surprise as a public information shock is not significant. Results of ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively.

Test	\mathbf{S}_0	\mathbf{D}_0	\mathbf{D}_1	AOF	AOI	Sur	
Panel A: Significan	t Results in I	Using Full, Rel	ative, and the M	lost Important	Sample		
A1: The Full Samp	le						
RV	**	_	_	_	_	_	
RV_{No-Lag}	**	_	_	_	_	_	
RV_{No-Sur}	***	_	_	_	_	_	
${\rm RV}_{Sur}^{Pos-Neg}$	**	_	_	_	_	_	
RV^{US}	**	_	_	_	-	_	
AR	*	_	_	**	_	_	
\mathbf{R}^{resi}	_	_	***	**	_	_	
${\rm RV}^{Joint\ Variables}$	* * *	_	_	_	-	_	
${\rm RV}^{All\ Times}$	* * *	**	_	_	**	_	
A2: The Relatively	Important S	ample					
AR	*	_	_	_	_	_	
\mathbf{R}^{resi}	_	_	***	*	-	_	
$\mathrm{RV}^{All\ Times}$	*	_	**	_	**	_	
A3: The Most Impo	ortant Sampl	e					
${\rm RV}^{Joint\ Variables}$	*	_	*	_	_	_	
${\rm RV}^{All\ Times}$	_	*	_	_	_	_	
Panel B: Significant Results in Using Economic Samples							
RV^{Output}	*	_	**	_	_	_	
${\sf RV}^{Prices}$	_	*	_	_	* * *	_	
$RV^{Employment}$	_	*	_	*	_	_	
${\rm RV}^{Decision\ Rate}$	_	_	***	**	**	_	

that, spread is more likely to be the candidate compared with other liquidity variables. In robustness tests, I classify the macroeconomic announcements according to their economic functions: output, prices, employment, and decision rate. Results show that private information shocks are the main driver of volatility compared with public information shocks.

2.5.2 Conclusion

In this chapter, I use the limit order book to evaluate the impact of prescheduled macroeconomic announcements on realised variance in the Canadian FX market. I assert that the limit order book is one of the main sources of information: specifically asymmetric information and private information. I model the impact of information shocks through two main sources: public information shocks and private information shocks.

I find that public information shocks do not play much of a role in the Canadian FX market. On the other hand, private information is a key factor that explains volatility changes in the market. In other words, having access to private information (directly or indirectly) is the main driving force of shocks in the market, and has a significant effect on volatility. I exploit liquidity measures as a proxy for private information shocks. My results show that spread at the best quote is the most significant variable among all liquidity measures, and provides consistent results through my different samples and measures. Other liquidity measures do not reveal such robust results. I examine the macroeconomic announcements through the *Full*, *Relatively Important*, and *Most Important* samples. My results show that the *Relatively Important* and *Most Important* samples have less significant results than the *Full* sample, and posit that this could be due to extensive public analysis about more important news events, resulting in less scope for exploiting private information.

I extend the results through my robustness checks. The first robust-

ness check considers absolute return and modified absolute return as two alternative proxies for realised variance. Although the results do not hold the same pattern for these different measures of volatility, they confirm my main result that private information shocks are more important than public information shocks. I give the caveat that, compared with the other two proxies, realised variance is able to take into account the impact of all quotes within the interval while the other proxies miss this effect. Hence, I ascribe the most importance to my realised variance results.

The second robustness check considers the effect of all liquidity measures jointly. This confirms that the effect of volatility changes comes through sensitivity to spread at the best quote, depth at the best quote, and absolute order imbalance liquidity measures. Finally, I examine the effect of news occurring at times other than at 8:30. My results confirm that private information shocks are the main driving force of volatility, and spread at the best quote is the main liquidity measure with significant and consistent results. I also evaluate the magnitude change of liquidity measures in smaller intervals of three 5-minute intervals before and after 8:30, between news days and no-news days. I analyse if they are large enough to make this sensitivity change economically significant. My results show that spread at the best quote and absolute order flow are the main liquidity measures with significant results, confirming spread's importance.

Chapter 3

Optimal Enterprise Use Switching Policy

3.1 Introduction

Real options valuation is an important area of research that is used in evaluating projects and investment decisions under uncertainty. I use this technique to evaluate a switching problem with two products 1 and 2. As product net incomes change, agents are incentivised to switch from one type of product to another. My models evaluate the switch frequency of relatively fast switching processes compared with prolonged ones. I conjecture that the latter involves higher uncertainty and time cost that is not captured precisely in the literature.

In this chapter, I examine the application of real options analyses to switching problems. I evaluate a situation where agents can switch between two different products in response to product net income changes. The term real options was coined by Myers (1977) and it refers to available options in an investment. There are a wide range of possible options in a project, namely: the option to defer, abandon, expand, contract, and switch. Real options analysis is widely used by scholars to evaluate investment decisions.¹

I contribute to the literature in several ways. One of the common assumptions of real options to evaluate industry entry and exit is irreversibility.² In this chapter, I relax this assumption to analyse real options to switching problems with reversibility. Reversibility is a spectrum that on one side there is costless reversibility where no real option exists. On the other side, there is costly complete irreversibility. In my models, agents have the option to switch from one product to another and reverse back to the original product. However, switches are costly, which suppresses the probability of switching. As a result, I am interested to know

¹ See, for example, McDonald and Siegel (1985a), Brennan and Schwartz (1985), Mc-Donald and Siegel (1985b), Majd and Pindyck (1987), Carr (2009), Milne and Whalley (2000), Novy-Marx (2007), Guthrie (2007).

² In this chapter, reversibility means being able to go back to the original product.

what explicit/implicit costs might occur in the switching process. The literature typically introduces a cash cost that captures the transition cost from one state to another. However, it is not clear to what extent does the cash cost capture all switching costs? Pure cost models may be appropriate when agents can switch relatively fast (e.g. livestock farming where farmers switch from farming sheep to cows and vice versa). However, in some activities, this is not necessarily true as there is also a waiting period before the predetermined amount of enterprise becomes productive again. For example, in horticulture, farmers face not only a cash cost but also a prolonged period of no production while new crops (i.e. saplings) reach maturity. In this regard, I extend the existing literature (e.g. Majd and Pindyck (1987) and Milne and Whalley (2000)) by adding a time cost (implicit/opportunity cost) as well as a cash cost (explicit cost). In other words, I evaluate the interaction between uncertainty, reversibility, timing, and switching costs at the same time in order to find the optimal policy. The importance of this approach is amplified when the transition is time consuming. Suppose an agent decides to switch to a lucrative product that entails a long switching process. The risk is that by the time the agent completes the transition process, the product net income may no longer be as lucrative, or the direction of switching may change. In this case, the agent not only loses the existing cash flows due to prolonged switching process but also ends up with a poor outcome, financially. Similarly, the reversibility option is exposed with the same risk. In the case where agents decide to reverse the process, they wait for a long time to go back to the original product. Hence, my advanced model (the cash-time model) has this feature to capture the time cost of switching precisely.

Among studies about real options, Song, Zhao, and Swinton (2011) use a real options framework to evaluate farmers' decisions to switch between traditional annual crops and perennial energy crops. Their results show that stochastic returns and government polices affect the optimal switching policy. The government can reduce the cost of converting land

from corn-soybean to switch-grass by paying subsidies to incentivise farmers to make a one way switch. However, the indirect impact of this policy provides reverse incentives for farmers to switch from switch-grass to corn-soybean due to the trivial cost of reversing the land back to switch-grass. Nevertheless, the direct impact of this policy outweighs the indirect impact. My model differs from their model by taking into consideration the time cost and by giving the agents the option to reverse their decisions during the switch. In this regard, I conjecture that my model outperforms their model by capturing the explicit cost (the cash cost) as well as the implicit cost (the time cost) and may provide a better explanation for why agents may not act when there is a switching option or why the triggers of switching may vary from conventional cost-benefit analyses.

As I mentioned earlier, a part of the real options literature has focused on irreversible investments and entry/exit strategies. For instance, Shah and Ando (2016) examine the choice between conserving or converting land in a case study in Indonesia. Farmers in this study have the option to conserve the land and earn carbon credits, or convert the land to palm oil and earn returns from it. They analyse the financial and ecological cost of temporary and permanent land conservation policy incentives, which leads to policy advice. Also, Cunha-e-sa and Franco (2016) evaluate an irreversible real options problem in land redevelopment. They evaluate the residential rent changes and its impact on rural areas to switch to residential. This study advises that boundaries such as development halt periods and minimum-lot-sizes can postpone land conversion and help to protect open spaces. An example of entry/exit problem in agriculture is the evaluation of wine grape vineyard investment and disinvestment in north-west Victoria, Australia, by Seyoum-Tegegn and C. Chan (2013). They evaluate the entry/exit triggers of this industry under price and yield uncertainty using a real options approach, which is ignored by conventional cost-benefit analysis. Exploiting real options techniques makes them able to explain why investors might be in hysteresis and inaction zones as an

optimal response, which results in lower exit and higher entry triggers. High vineyard establishment costs make investors reluctant to exit easily and wait for a better future. Similarly, Cyr, Hanagriffb, and Kwongc (2010) examine the wine industry of Texas in the US. Their results indicate that wine prices should increase by 30 to 40 percent to give incentive for investment in vineyard industry. They capture the hysteresis in investment in a real options framework. Tauer (2006) sets up an entry/exit real options problem to evaluate the dairy market in New York. He measures the optimal entry and exit milk prices where agents decide to invest or disinvest in this industry.

A part of real options studies evaluate switches with regime switching. Bollen (1999) evaluates real options to expand project and incorporates the product life cycle. His results show that the conventional method undervalue the option to contract and overvalue the option to expand. This results become more important in high technology industries. Bensoussan, Hoe, Yan, and Yin (2017) provide a mathematical solution for an irreversible investment that is examined in a duopoly games between leader and followers. They achieve an optimal investment solution under an uncertain cash flow regime and present implications in regime switching. There are more studies in this strand that are different from our approach in terms of methodology and underlying assumptions.

The remainder of this chapter proceeds as follows: In Section 3.2, I describe my methodology for solving a partial differential equation for two assets in order to find the optimal enterprise value with different products. In this regard, I derive the PDE for two competing products and then explain how to solve this PDE using finite differences and the alternating direction implicit algorithm. Section 3.3 presents the baseline parameter values and a range of scenarios to analyse. I simulate underlying product net income and generate statistics for optimal behaviour. Sections 3.3.3 and 3.3.4 evaluate my numerical results by analysing the opti-

mal exercise frontiers and then provides sensitivity analysis for the model parameters. Finally, in Section 3.4, I summarise my results and provide advice for policy makers, agents, and future research.

3.2 Methodology

In this section, I explain how to derive and solve a partial differential equation (PDE) for enterprises with two competing products 1 and 2. The PDE solution enables us to examine what product net income changes will cause agents to switch enterprise use. In this regard, I derive the PDE and employ its solution to construct three models for the switching problem, presenting the solution process in Section 3.2.1.

In this study, I assume a complete market. The importance of risk aversion depends on whether the market is complete or not. In a complete market, all products are tradable and agents maximise the value of their assets and work with risk neutral probabilities. In contrast, in an incomplete market, valuation will require equilibrium to be solved for, which in turn will depend on the utility function (and other instruments of) the agents concerned. I consider an agent (e.g. a farmer) who has access to a predetermined amount of enterprise (e.g. a parcel of land). At any point in time, this enterprise can be allocated to a variety of states, which will be discussed subsequently below. The enterprise value in all cases depends on the current value of the two potential products. The net income of two competing products, S_1 and S_2 , are assumed to follow Geometric Brownian Motions (GBMs). I define the processes by the following stochastic differential equations:

$$dS_1 = \mu_1 S_1 dt + \sigma_1 S_1 dz_1 dS_2 = \mu_2 S_2 dt + \sigma_2 S_2 dz_2,$$
(3.1)

where S_i is the product i net income, μ_i and σ_i are the drift and volatility of product i, dt is a small interval of time, and dz_i is an increment of a Brownian motion. The enterprise value is a function of product net income and time, $V_i(S_1, S_2, t)$. I use the Itô formula to derive the enterprise value process, as follows:

$$dV_{i} = \frac{\partial V_{i}}{\partial S_{1}} \left[\mu_{1} S_{1} dt + \sigma_{1} S_{1} dz_{1} \right] + \frac{\partial V_{i}}{\partial S_{2}} \left[\mu_{2} S_{2} dt + \sigma_{2} S_{2} dz_{2} \right] + \frac{\partial V_{i}}{\partial t} dt$$

$$+ \frac{1}{2} \left[\frac{\partial^{2} V_{i}}{\partial S_{1}^{2}} \left[\sigma_{1}^{2} S_{1}^{2} dt \right] + \frac{\partial^{2} V_{i}}{\partial S_{2}^{2}} \left[\sigma_{2} S_{2}^{2} dt \right] + 2 \frac{\partial^{2} V_{i}}{\partial S_{1} \partial S_{2}} \left[\rho \sigma_{1} \sigma_{2} S_{1} S_{2} dt \right] \right],$$

$$(3.2)$$

Rearranging (3.2) and taking the expectation yields:

$$E(dV_{i}) = \left[\mu_{1}S_{1}\frac{\partial V_{i}}{\partial S_{1}} + \mu_{2}S_{2}\frac{\partial V_{i}}{\partial S_{2}} + \frac{\partial V_{i}}{\partial t} + \frac{1}{2}\sigma_{1}^{2}S_{1}^{2}\frac{\partial^{2}V_{i}}{\partial S_{1}^{2}} + \frac{1}{2}\sigma_{2}^{2}S_{2}^{2}\frac{\partial^{2}V_{i}}{\partial S_{2}^{2}} + \rho\sigma_{1}\sigma_{2}S_{1}S_{2}\frac{\partial^{2}V_{i}}{\partial S_{1}\partial S_{2}}\right]dt,$$

$$(3.3)$$

In equilibrium, the expected enterprise return must be equal to the capital gains over the time increment, dt (i.e. $rV_0dt = E(dV_0)$). An active enterprise has an extra term (return) over time that should be added to the capital gain. The expected value obeys:

$$rV_i dt = E(dV_i) + CF_i(S_1, S_2) dt,$$
 (3.4)

where $E(dV_0)$ is the expected change of enterprise value, r is the (constant) fair expected return associated with enterprise ownership, V_0 is the inactive enterprise value that plays the role of transition state due to having the option to produce any of products, dt is time increment, and CF_i are cash flows in state i. Equating equations (3.3) and (3.4) and dividing by dt

gives the following PDE:

$$0 = \frac{\partial V_i}{\partial t} + \mu_1 S_1 \frac{\partial V_i}{\partial S_1} + \mu_2 S_2 \frac{\partial V_i}{\partial S_2} + \frac{1}{2} \sigma_1^2 S_1^2 \frac{\partial^2 V_i}{\partial S_1^2} + \frac{1}{2} \sigma_2^2 S_2^2 \frac{\partial^2 V_i}{\partial S_2^2}$$

$$+ \rho \sigma_1 S_1 \sigma_2 S_2 \frac{\partial^2 V_i}{\partial S_1 \partial S_2} - r V_i + C F_i$$

$$\equiv \frac{\partial V_i}{\partial t} + \mathcal{L} V_i + C F_i,$$
(3.5)

where \mathcal{L} is the operator:

$$\mathcal{L}V = \mu_1 S_1 \frac{\partial V}{\partial S_1} + \mu_2 S_2 \frac{\partial V}{\partial S_2} + \frac{1}{2} \sigma_1^2 S_1^2 \frac{\partial^2 V}{\partial S_1^2} + \frac{1}{2} \sigma_2^2 S_2^2 \frac{\partial^2 V}{\partial S_2^2}$$

$$+ \rho \sigma_1 S_1 \sigma_2 S_2 \frac{\partial^2 V}{\partial S_1 \partial S_2} - rV$$
(3.6)

I solve the PDE (3.5) (see Section 3.2.1) with different state structures to evaluate agents' behaviour in a real options switching framework. I break my model into three models: (1) the cash-cost model, (2) the cashtime model, and (3) the projection model. The cash-cost model considers only the cash cost of switching, whereas the cash-time model not only considers the cash cost but also the time cost (opportunity cost). Finally, the projection model presents an approximation of the cash-time model, which is relatively faster in computation. To start with the cash-cost model, it is worth to know that the switching process is not always time consuming and in some cases such as livestock farming (e.g. beef and sheep) the process is faster compared with horticulture. Hence, the agents switch from one state to another by purely paying the cash cost (i.e. without going through a long waiting process). In this regard, I assume a scenario where agents have the option to have product 1 in state 1 and product 2 in state 2. The enterprise value in each state comes through product net income. The enterprise optimisation equation for the cash-cost model has

the following form:

$$\mathcal{L}V_i + CF_i \le 0 \tag{3.7a}$$

$$V_i \ge V_i - K_{ii,cc} \tag{3.7b}$$

where \mathcal{L} is the operator (3.6), V_i is the enterprise value in state i, CF_i is the cash flow in state i, V_j is the enterprise value in state j, and $K_{ij,cc}$ is the cash cost to be paid and switch from state i to state j. However, in reality switches are not always instantaneous and there are switches (e.g. horticulture) that require a prolonged process. Note that the problem is elliptic since I do not have the $(\frac{\partial V_i}{\partial t})$ term. We set up the cash-time model by assuming two productive states 1 and 2, and a transition state in between that agents cross to transfer from product 1 to 2 and vice versa. The productive states have the following options:

$$\mathcal{L}V_i + CF_i \le 0 \tag{3.8a}$$

$$V_i \ge V_{i,u}(0) - K_{i,u},$$
 (3.8b)

where $V_{i,u}(0)$ is the enterprise value where removing product i has just started, and $K_{i,u}$ is the upfront cash cost of removing product i. Note that the problem is elliptic since I do not have the $(\frac{\partial V_i}{\partial t})$ term. Agents have incentives to switch between products based on their respective net income. Suppose agents who are producing product 1 realise that the competing product (2) is more lucrative. Hence, they have an incentive to switch from product 1 to product 2. However, in the interim, this results in having an enterprise with no net income, called the "transition" state. The transition state provides the value of removing. The removing process takes time, which could be relatively fast and inexpensive for some products, but could be lengthy and expensive if preparation is required. We model

the transition state as follows:

$$V_{i,u}(T_i^{Remove}) = V_{Tr} (3.9a)$$

$$\mathcal{L}V_{i,u} + \frac{\partial V_{i,u}}{\partial t} - k_{i,u} = 0, \tag{3.9b}$$

where $V_{i,u}(t)$ is the enterprise value t years into the removing process, T_i^{Remove} is the time taken to remove product i, V_{Tr} is the transition enterprise value, the term $\frac{\partial V_{i,u}}{\partial t}$ accounts for the fact that time now affects enterprise value and I solve a parabolic PDE, and $k_{i,u}$ is the ongoing cost of removing product i that is incurred throughout the removing process. When agents finish the removing process, they are in the transition state, whose value obeys:

$$\mathcal{L}V_{Tr} \le 0 \tag{3.10a}$$

$$V_{Tr} \ge max \left(V_{1,p}(0) - K_{1,p}, V_{2,p}(0) - K_{2,p} \right)$$
 (3.10b)

where $V_{i,p}(0)$ is the enterprise value where product i has just been initiated (but has not yet reached productivity), and $K_{i,p}$ is the upfront cash cost of initialising product i. Agents in the transition state have the option to continue the switching process and initialise another product (or go back to the old product). However, the initialisation process takes time and is potentially slow and expensive. The enterprise value where initialisation is taking place solves:

$$V_{i,p}(T_i^{Initiate}) = V_i (3.11a)$$

$$\mathcal{L}V_{i,p} + \frac{\partial V_{i,p}}{\partial t} - k_{i,p} \le 0 \tag{3.11b}$$

$$V_{i,p} \ge V_{i,u}(0) - K_{i,u},$$
 (3.11c)

where $T_i^{Initiate}$ is the time taken to initialise product i, and $k_{i,p}$ is the on-

going cash cost throughout the initialisation process. Note that, like the removal state, here I solve a parabolic PDE.

Finally, the cash-time model can be approximated by the projection model, which computes the time cost without capturing the removal and initialisation process formally. Agents in the projection model switch from one product to another or go back to the original product instantaneously. All the waiting cost (in terms of the forgone profits) is modeled as the extra monetary cost of switching to be paid by agents. The projection model approximates the time cost as follows:

$$K_{ij,tc} = \int_0^{T_{i,u}} k_{i,u} e^{-rt} dt + \int_{T_{i,u}}^{T_{i,u}+T_{j,p}} k_{j,p} e^{-rt} dt + \frac{S_i}{\mu_i - r} \left[e^{(\mu_i - r)(T_{i,u} + T_{j,p})} - 1 \right],$$
(3.12)

where $K_{ij,tc}$ is the estimated time cost of switching from product i to j, $T_{i,u}$ is the time to remove product i, $T_{j,p}$ is the time to initialise product j, and μ_i , r, and S_i are drift, discount rate, and cash flow of product i respectively. The enterprise value in the projection model follows the following form:

$$\mathcal{L}V_i \le 0 \tag{3.13a}$$

$$V_i \ge V_j - K_{ij,tc} - K_{ij,cc},\tag{3.13b}$$

where $K_{ij,cc}$ is the upfront cash cost of removing product i and initialising product j (i.e. $K_{i,u} + K_{j,p}$). We conjecture that the frequency of switching increases in the projection model compared with the cash-time model due to the faster switching process. To this end, the cash-cost model neither computes the time cost formally nor approximates it, and focuses on only the cash cost. In contrast, the cash-time model by considering the exact time cost and the projection model by approximating the time cost capture a higher cost of switching. Hence, I conjecture that the frequency of switching in the cash-cost model will be considerably higher than the latter models since it overlooks the time cost entirely.

3.2.1 Solution Method

In this part, I present my solution technique for the PDE (3.5). The implicit Finite Difference Method (FDM) provides a stable solution. However, when the system of equations has a dimension greater than one, I lose the computational advantage of having a tridiagonal matrix. The solution to overcome this difficulty and make my computation more efficient is to use the Alternating Direction Implicit (ADI) method, which is a numerical approach to solving PDEs. Before applying the ADI method, I transform product net income into natural logarithms $x_1 = \ln(S_1)$, $x_2 = \ln(S_2)$ and derive the corresponding partial derivatives as follows:

$$\frac{\partial V_{i}}{\partial S_{1}} = \frac{\partial V_{i}}{\partial x_{1}} \frac{\partial x_{1}}{\partial S_{1}} = \frac{\partial V_{i}}{\partial x_{1}} \frac{1}{S_{1}}$$

$$\frac{\partial V_{i}}{\partial S_{2}} = \frac{\partial V_{i}}{\partial x_{2}} \frac{\partial x_{2}}{\partial S_{2}} = \frac{\partial V_{i}}{\partial x_{2}} \frac{1}{S_{2}}$$

$$\frac{\partial^{2} V_{i}}{\partial S_{1}^{2}} = \frac{\partial}{\partial S_{1}} \frac{\partial V_{i}}{\partial S_{1}} = \frac{\partial}{\partial S_{1}} \left(\frac{\partial V_{i}}{\partial x_{1}} \frac{1}{S_{1}} \right) = \frac{1}{S_{1}} \frac{\partial^{2} V_{i}}{\partial x_{1}^{2}} \frac{1}{S_{1}} + \frac{\partial V_{i}}{\partial x_{1}} \left(-\frac{1}{S_{1}^{2}} \right)$$

$$= \frac{1}{S_{1}^{2}} \frac{\partial^{2} V_{i}}{\partial x_{1}^{2}} - \frac{1}{S_{1}^{2}} \frac{\partial V_{i}}{\partial x_{1}}$$

$$\frac{\partial^{2} V_{i}}{\partial S_{2}^{2}} = \frac{\partial}{\partial S_{2}} \frac{\partial V_{i}}{\partial S_{2}} = \frac{\partial}{\partial S_{2}} \left(\frac{\partial V_{i}}{\partial x_{2}} \frac{1}{S_{2}} \right) = \frac{1}{S_{2}} \frac{\partial^{2} V_{i}}{\partial x_{2}^{2}} \frac{1}{S_{2}} + \frac{\partial V_{i}}{\partial x_{2}} \left(-\frac{1}{S_{2}^{2}} \right)$$

$$= \frac{1}{S_{2}^{2}} \frac{\partial^{2} V_{i}}{\partial x_{2}^{2}} - \frac{1}{S_{2}^{2}} \frac{\partial V_{i}}{\partial x_{2}}$$

$$\frac{\partial^{2} V_{i}}{\partial S_{1} \partial S_{2}} = \frac{\partial}{\partial S_{2}} \frac{\partial V_{i}}{\partial S_{1}} = \frac{\partial}{\partial S_{2}} \left(\frac{\partial V_{i}}{\partial x_{1}} \frac{1}{S_{1}} \right) = \frac{1}{S_{1}} \frac{1}{S_{2}} \frac{\partial^{2} V_{i}}{\partial x_{1} \partial x_{2}},$$

$$\frac{\partial^{2} V_{i}}{\partial S_{1} \partial S_{2}} = \frac{\partial}{\partial S_{2}} \frac{\partial V_{i}}{\partial S_{1}} = \frac{\partial}{\partial S_{2}} \left(\frac{\partial V_{i}}{\partial x_{1}} \frac{1}{S_{1}} \right) = \frac{1}{S_{1}} \frac{1}{S_{2}} \frac{\partial^{2} V_{i}}{\partial x_{1} \partial x_{2}},$$

I substitute the derivatives from equation (3.14) into equation (3.5) and derive the following PDE.

$$\frac{\partial V_i}{\partial t} = \nu_1 \frac{\partial V_i}{\partial x_1} + \nu_2 \frac{\partial V_i}{\partial x_2} + \frac{1}{2} \sigma_1^2 \frac{\partial^2 V_i}{\partial x_1^2} + \frac{1}{2} \sigma_2^2 \frac{\partial^2 V_i}{\partial x_2^2} + \rho \sigma_1 \sigma_2 \frac{\partial^2 V_i}{\partial x_1 \partial x_2} - rV_i + e^{x_i},$$
(3.15) where ν_1 is $(\mu_1 - \frac{1}{2}\sigma_1^2)$ and ν_2 is $(\mu_2 - \frac{1}{2}\sigma_2^2)$.

The ADI method is a powerful way of solving multidimensional PDEs. However, the ADI is not initially developed to handle cross-partial terms. Therefore, in order to solve equation (3.15), I need to use a transformation to remove the $\frac{\partial^2 V_i}{\partial x_1 \partial x_2}$ term. Clewlow and Strickland (1998) propose to compute the eigenvectors and eigenvalues of the covariance matrix of x_1 and x_2 .

$$\begin{bmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{bmatrix} = \begin{bmatrix} e_{11} & e_{21} \\ e_{12} & e_{22} \end{bmatrix} \begin{bmatrix} \lambda_1 & 0 \\ 0 & \lambda_2 \end{bmatrix} \begin{bmatrix} e_{11} & e_{12} \\ e_{21} & e_{22} \end{bmatrix},$$

where on the left hand side, I have covariance matrix of x_1 and x_2 and on the right hand side, the eigenvalue decomposition. The eigenvectors give the linear combinations of x_1 and x_2 , which are uncorrelated:

$$\begin{bmatrix} y_1 \\ y_2 \end{bmatrix} = \begin{bmatrix} e_{11} & e_{12} \\ e_{21} & e_{22} \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} e_{11}x_1 + e_{12}x_2 \\ e_{21}x_1 + e_{22}x_2 \end{bmatrix}$$
(3.16a)

$$\begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} e_{11} & e_{12} \\ e_{21} & e_{22} \end{bmatrix} \begin{bmatrix} y_1 \\ y_2 \end{bmatrix} = \begin{bmatrix} e_{11}y_1 + e_{12}y_2 \\ e_{21}y_1 + e_{22}y_2 \end{bmatrix}$$
(3.16b)

$$\begin{bmatrix} \alpha_1 \\ \alpha_2 \end{bmatrix} = \begin{bmatrix} e_{11} & e_{12} \\ e_{21} & e_{22} \end{bmatrix} \begin{bmatrix} v_1 \\ v_2 \end{bmatrix} = \begin{bmatrix} e_{11}v_1 + e_{12}v_2 \\ e_{21}v_1 + e_{22}v_2 \end{bmatrix},$$
(3.16c)

Now I use equation 3.16 to derive the corresponding derivatives of equa-

tion 3.15 with respect to x_1 and x_2 as follows:

$$\begin{split} \frac{\partial V_{i}}{\partial x_{1}} &= \frac{\partial V_{i}}{\partial y_{1}} \cdot \frac{\partial y_{1}}{\partial x_{1}} + \frac{\partial V_{i}}{\partial y_{2}} \cdot \frac{\partial y_{2}}{\partial x_{1}} = \frac{\partial V_{i}}{\partial y_{1}} \cdot e_{11} + \frac{\partial V_{i}}{\partial y_{2}} \cdot e_{21} \\ \frac{\partial V_{i}}{\partial x_{2}} &= \frac{\partial V_{i}}{\partial y_{1}} \cdot \frac{\partial y_{1}}{\partial x_{2}} + \frac{\partial V_{i}}{\partial y_{2}} \cdot \frac{\partial y_{2}}{\partial x_{2}} = \frac{\partial V_{i}}{\partial y_{1}} \cdot e_{12} + \frac{\partial V_{i}}{\partial y_{2}} \cdot e_{22} \\ \frac{\partial^{2} V_{i}}{\partial x_{1}^{2}} &= e_{11} \left[\frac{\partial^{2} V_{i}}{\partial y_{1}^{2}} \cdot \frac{\partial y_{1}}{\partial x_{1}} + \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} \cdot \frac{\partial y_{2}}{\partial x_{1}} \right] + e_{21} \left[\frac{\partial^{2} V_{i}}{\partial y_{2}^{2}} \cdot \frac{\partial y_{2}}{\partial x_{1}} + \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} \cdot \frac{\partial y_{1}}{\partial x_{1}} \right] \\ &= e_{11}^{2} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1}^{2}} + 2 \cdot e_{11} \cdot e_{21} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} + e_{21}^{2} \cdot \frac{\partial^{2} V_{i}}{\partial y_{2}^{2}} \\ &\frac{\partial^{2} V_{i}}{\partial x_{2}^{2}} = e_{12} \left[\frac{\partial^{2} V_{i}}{\partial y_{1}^{2}} \cdot \frac{\partial y_{1}}{\partial x_{2}} + \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} \cdot \frac{\partial y_{2}}{\partial x_{2}} \right] + e_{22} \left[\frac{\partial^{2} V_{i}}{\partial y_{2}^{2}} \cdot \frac{\partial y_{2}}{\partial x_{2}} + \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} \cdot \frac{\partial y_{1}}{\partial x_{2}} \right] \\ &= e_{12}^{2} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1}^{2}} + 2 \cdot e_{12} \cdot e_{22} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} + e_{22}^{2} \cdot \frac{\partial^{2} V_{i}}{\partial y_{2}^{2}} \\ &\frac{\partial^{2} V_{i}}{\partial x_{2}^{2}} = e_{11} \left[\frac{\partial^{2} V_{i}}{\partial y_{1}^{2}} \cdot \frac{\partial y_{1}}{\partial x_{2}} + \frac{\partial V_{i}}{\partial y_{1} \partial y_{2}} \cdot \frac{\partial y_{2}}{\partial x_{2}} \right] + e_{21} \left[\frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} \cdot \frac{\partial y_{1}}{\partial x_{2}} + \frac{\partial^{2} V_{i}}{\partial y_{2}^{2}} \cdot \frac{\partial y_{2}}{\partial x_{2}} \right] \\ &= e_{11} \cdot e_{12} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1}^{2}} \cdot \frac{\partial y_{1}}{\partial x_{2}} + \frac{\partial V_{i}}{\partial y_{1} \partial y_{2}} \cdot \frac{\partial y_{2}}{\partial x_{2}} \right] + e_{21} \left[\frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} \cdot \frac{\partial y_{1}}{\partial x_{2}} + \frac{\partial^{2} V_{i}}{\partial y_{2}^{2}} \cdot \frac{\partial y_{2}}{\partial x_{2}} \right] \\ &= e_{11} \cdot e_{12} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1}^{2}} + e_{11} \cdot e_{22} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} + e_{21} \cdot e_{12} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} + e_{21} \cdot e_{22} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} + e_{21} \cdot e_{22} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} + e_{22} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}} + e_{22} \cdot \frac{\partial^{2} V_{i}}{\partial y_{1} \partial y_{2}}$$

Substituting the corresponding derivatives from (3.17) into (3.15) and simplifying the terms gives:

$$\frac{\partial V_i}{\partial t} = \alpha_1 \frac{\partial V_i}{\partial y_1} + \alpha_2 \frac{\partial V_i}{\partial y_2} + \frac{1}{2} \lambda_1 \frac{\partial^2 V_i}{\partial y_1^2} + \frac{1}{2} \lambda_2 \frac{\partial^2 V_i}{\partial y_2^2} - rV_i + e^{(e_{i1}y_1 + e_{i2}y_2)}, \quad (3.18)$$

Generally, in numerical analysis, I replace continuous PDEs with discrete approximations, which reduce continuous differential equations into equivalent linear equation systems. I discretise my partial differential equation according to the corresponding dimensions and approximate the solution at isolated points rather than developing the solution everywhere. Hence, I define $\underline{y_i}$ and $\overline{y_i}$ as the lower and upper bounds with N_i number of steps in dimension y_i . In general, the greater the number of points, the more accurate the solution. According to the type of finite difference in approximation, there will be errors called *truncation errors*. I can construct finite difference approximations of different accuracy by defining a linear combination of function values and expanding the values by Taylor series. As an example, suppose I want to find the second order approximation for the first derivative with a forward finite difference using three steps, I have the following set up:

$$\frac{\partial V}{\partial y_1} = \frac{aV(y_1, y_2) + bV(y_1 + \Delta y_1, y_2) + cV(y_1 + 2\Delta y_1, y_2)}{\Delta y_1} + O(\Delta y_1)^2,$$
(3.19)

where $O(\Delta y_1)^2$ is the truncation error or high order approximation that I neglect in my approximation. I need to find the coefficients a, b, and c. This is a straight forward process if I use the Taylor series to expand the values in the y_1 dimension as follows:

$$V(y_{1} + \Delta y_{1}) = V(y_{1}) + \Delta y_{1} \frac{\partial V}{\partial y_{1}} + \frac{(\Delta y_{1})^{2}}{2!} \frac{\partial^{2} V}{\partial y_{1}^{2}} + \frac{(\Delta y_{1})^{3}}{3!} \frac{\partial^{3} V}{\partial y_{1}^{3}} + O(\Delta y_{1})^{4}$$

$$V(y_1 + 2\Delta y_1) = V(y_1) + 2\Delta y_1 \frac{\partial V}{\partial y_1} + 4\frac{(\Delta y_1)^2}{2!} \frac{\partial^2 V}{\partial y_1^2} + 8\frac{(\Delta y_1)^3}{3!} \frac{\partial^3 V}{\partial y_1^3} + O(\Delta y_1)^4,$$

I substitute the corresponding approximations of y_1 with forward steps into (3.19). As a result, the following system of equations should be solved:

$$\begin{bmatrix} 1 & 1 & 1 \\ 0 & 1 & 2 \\ 0 & 1 & 4 \end{bmatrix} \begin{bmatrix} a \\ b \\ c \end{bmatrix} = \begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix},$$

Solving this system of equations gives a=-1.5, b=2, and c=-0.5. Hence, the second order approximation of the first derivative with three

steps forward is as follows:

$$\frac{\partial V}{\partial y_1} = \frac{-3V(y_1, y_2) + 4V(y_1 + \Delta y_1, y_2) - V(y_1 + 2\Delta y_1, y_2)}{2\Delta y_1} + O(\Delta y_1)^2,$$
(3.20)

Similarly, for all the other finite difference approximations, I construct the function based on the order of accuracy, order of derivative, and the scheme type (forward, backward, or central). To this end, I employ the Peaceman and Rachford (1955) ADI method to solve my PDE in equation (3.18) for the first and second derivatives with second order accuracy. I discretise the spatial terms of the operator as the discrete operator $\mathcal{L}_{y_i}^{mid}$ when considering a point on the interior of the grid (i.e. $\underline{y}_i < y_i < \overline{y_i}$):

$$\mathcal{L}_{y_i}^{mid}V = \alpha_i \left[\frac{V_{y_i + \Delta y_i, y_j} - V_{y_i - \Delta y_i, y_j}}{2\Delta y_i} \right] + \frac{1}{2}\lambda_i \left[\frac{V_{y_i + \Delta y_i, y_j} - 2V_{y_i, y_j} + V_{y_i - \Delta y_i, y_j}}{(\Delta y_i)^2} \right] - rV_{y_i, y_j}$$

$$(3.21)$$

where if i=1 then j=2 and vice versa. Similarly, I define operators $\mathcal{L}_{y_i}^{\mathcal{LB}}$ and $\mathcal{L}_{y_i}^{\mathcal{UB}}$ for the equations system as follows:

$$\mathcal{L}_{y_i}^{\mathcal{LB}}V = \alpha_i \left[\frac{-3V_{y_i,y_j} + 4V_{y_i + \Delta y_i,y_j} - V_{y_i + 2\Delta y_i,y_j}}{2\Delta y_i} \right]$$

$$+ \frac{1}{2}\lambda_i \left[\frac{2V_{y_i,y_j} - 5V_{y_i + \Delta y_i,y_j} + 4V_{y_i + 2\Delta y_i,y_j} - V_{y_i + 3\Delta y_i,y_j}}{(\Delta y_i)^2} \right] - rV_{y_i,y_j}$$

$$\mathcal{L}_{y_{i}}^{\mathcal{UB}}V = \alpha_{i} \left[\frac{3V_{y_{i},y_{j}} - 4V_{y_{i}-\Delta y_{i},y_{j}} + V_{y_{i}-2\Delta y_{i},y_{j}}}{2\Delta y_{i}} \right]$$

$$+ \frac{1}{2}\lambda_{i} \left[\frac{2V_{y_{i},y_{j}} - 5V_{y_{i}-\Delta y_{i},y_{j}} + 4V_{y_{i}-2\Delta y_{i},y_{j}} - V_{y_{i}-3\Delta y_{i},y_{j}}}{(\Delta y_{i})^{2}} \right] - rV_{y_{i},y_{j}}$$
(3.22)

where if i=1 then j=2 and vice versa. The system of equations using the

middle equations and boundaries is built up as follows:

$$\mathcal{L}_{y_i} = \begin{cases}
\mathcal{L}_{y_i}^{mid}, & \text{if} \qquad \underline{y}_i < y_i < \overline{y}_i \\
\mathcal{L}_{y_i}^{\mathcal{LB}}, & \text{if} \qquad y_i = \underline{y}_i \\
\mathcal{L}_{y_i}^{\mathcal{UB}}, & \text{if} \qquad y_i = \overline{y}_i
\end{cases} ,$$
(3.23)

To solve theses systems of equations, I use the ADI method, which is widely used to solve the PDEs in numerical analysis. This method was first used by Peaceman and Rachford (1955) and Douglas and Rachford (1955). The ADI method has the advantage of being unconditionally stable, as well as being second order accurate. I can also reduce the systems of equations that must be solved to being tridiagonal, resulting in fast computation of solutions. The ADI method works by solving the PDE using two partial time steps instead of a full time step. In the first partial time step, I solve the system of equation by having y_1 explicit and y_2 implicit (solving in rows), and then switch the direction by having y_1 implicit and y_2 explicit (solving in columns). I solve the following systems of equations at each time step:

$$0 = \left\lceil \frac{V_{y_1, y_2}^t - V_{y_1, y_2}^{t - \frac{\Delta t}{2}}}{\frac{\Delta t}{2}} \right\rceil + \mathcal{L}_{y_1} V^t + \mathcal{L}_{y_2} V^{t - \frac{\Delta t}{2}} + CF(y_1, y_2)$$
(3.24)

$$0 = \left[\frac{V_{y_1, y_2}^{t - \frac{\Delta t}{2}} - V_{y_1, y_2}^{t - \Delta t}}{\frac{\Delta t}{2}} \right] + \mathcal{L}_{y_1} V^{t - \frac{\Delta t}{2}} + \mathcal{L}_{y_2} V^{t - \Delta t} + CF(y_1, y_2)$$
(3.25)

Note that (3.24) generate a system of equations for each level of y_1 (N_1 systems of equations) while (3.25) generates a system of equations for each level of y_2 (N_2 systems of equations). Hence, a full time step requires solving N_1 ($N_2 \times N_2$) tridiagonal systems and N_2 ($N_1 \times N_1$) systems. This requires $O(N_1 \times N_2)$ operations to compute.

3.3 Numerical Illustration

3.3.1 Simulation and Bootstrap

So far, I have established three models to evaluate agent behaviour using a real options switching model under uncertainty and reversibility. These three models vary in terms of level of complexity and how they compute the cost of switching. The cash-time model is able to evaluate situations where the switching process is time consuming. A prolonged switching process with uncertainty means that agents are not sure if product net income remain in their favour by the time they finish the switching process. If the product net income are not according to what they plan, they may miss the current cash flows and end up in an undesirable situation. At the same time, having the option to reverse the switching process part way through is also uncertain due to the time taken to re-establish the old product. The cash-time model has the advantage of being able to capture all these impacts in terms of the time cost along with the cash cost of switching. The projection model is an approximation of the cash-time model, and captures the time cost partially compared with the cash-time model. In the projection model, switching occurs very fast and agents do not need to wait for a long time but the waiting cost is captured partially in terms of a lump sum cash cost to be paid when switching occurs. When the time of switching increases, the projection model fails to capture the exact cost of switching and underestimates the cost due to this approximation. Finally, the cash-cost model is the simplest model, it captures only the cash cost of switching, and neglects the time cost. Now, the question is: to what extent these three models are different from each other? In order to answer this question, I employ simulation and bootstrapping techniques to simulate a large number of agents and evaluate their behaviours and polices in each model. Then, I test if the results of these models are qualitatively/quantitatively different from each other. I examine agents' behaviour in different

scenarios and test if these three models are statistically different from each other.

I generate a simulation of 65000 paths over 200 years for two competing products.¹ This net income simulation helps examine agents' switching behaviour and evaluate when a threshold is crossed to trigger a switch, when they hit the barriers, and how many times they change the state of activity. I use equation (3.26) to generate the product net income processes.

$$\Delta \log S_1 = \left(\mu_1 - \frac{\sigma_1^2}{2}\right) \Delta t + \sigma_1 \epsilon_1$$

$$\Delta \log S_2 = \left(\mu_2 - \frac{\sigma_2^2}{2}\right) \Delta t + \sigma_2 \left(\rho_{12} \epsilon_1 + \sqrt{1 - \rho_{12}^2} \epsilon_2\right)$$
(3.26)

where $\Delta \log S_i$ are the changes in log net income of each product, μ_i are the drifts or expected cash flow growths of each product, σ_i are the volatilities of each product, Δt is the time step, ρ is the correlation between the two products, and ϵ_i is a normal random innovation. The ϵ_1 and ϵ_2 are independent and are normally distributed with a mean and variance of $N\left(0,\sqrt{\Delta t}\right)$.

Another useful technique that I exploit through my analysis is bootstrapping to test how the three models are significantly different from each other. I solve my optimisation problem according to baseline parameters and examine the behaviour of agents in terms of switching frequency in each model. Then I change my baseline parameters individually and compare the switching frequency across my three models as well as between before and after changes (see Sections 3.3.3 and 3.3.4). In this regard, I find the switching pattern of agents in different scenarios. I generate 10000 random samples with replacement based on my main results in each sce-

¹ The choice of 65000 paths and 200 years is ad hoc and can be any number. The life of 200 years might be a long time and a small number e.g. 40 years looks more realistic. The life of 40 years generates less switching and small probabilities that require scaling for a bigger numbers. This can be approximated by dividing the results by 5.

nario, which has the same size of 65000 observations. This resampling enables us to test whether the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place are quantitatively/qualitatively different across my three models in each scenario.¹

3.3.2 Baseline Parameters and Scenarios

I test my theory by proposing some baseline parameters. In this regard, Table 3.1 presents a set of reasonable numbers for my optimisation of enterprise allocation. In Section 3.3.4, I examine the sensitivity analysis and the effect of changes in each variable on switching. I evaluate how changes in input parameters may affect switch frequency across the three models and compared with baseline. The time cost changes according to the type of product. For example, in livestock farming (e.g. beef and sheep) there is almost no time cost, and I expect all three models to provide similar results, while with seasonal products (e.g. beetroot and carrot) there is a time cost of one year, and in horticulture (e.g. apples and grapes) the time cost increases to 5 or 6 years, which is a crucial time cost and I expect that the three models to lead to qualitatively different results. My modeling approach allows me to investigate these different cases under one general framework and to carry out counterfactual analyses.

According to my baseline parameters in Table 3.1, it costs \$1000 and takes 1 year to remove each product, whereas it costs \$10000 and takes 3 years to initialise Product 1 has a 10% volatility with 1% drift while product 2 has 12% volatility with 1% drift. These two products have correlation of 0.3 in a market with 5% discount rate. In Section 3.3.3, I use these parameters to solve my models for the value of the enterprise, extract the optimal

 $^{^1}$ I choose P $(n \ge 1)$ and P $(n \ge 2)$ due to a very few observation and small probability for numbers greater than 2.

TABLE 3.1: Benchmark and Sensitivity Analysis Scenarios

Note: Panel A displays the baseline parameters for two competing products in my analyses. The variables of interest are: cost, time, volatility, drift, discount rate, and correlation. In Panel B, I present the sensitivity analysis in four scenarios: Scenario 1 changes in cost and time, Scenario 2 changes in volatility and drift, Scenario 3 changes in discount rate, and Scenario 4 changes in correlation. Variables are scaled based on thousands (1000) per hectare and time variables are per annum. The sign + indicates that an increase in the variable would be expected to cause arise switching; - indicates that an increase in the variable causes the switching to decrease, and? indicates that the effect is not clear.

Parameters	Symbol	Product 1	Product 2	Effect
Panel A: Baseline Variables				
Cost (removal, initialisation)	K_{ij}	(1, 10)	(1, 10)	
Time (removal, initialisation)	T_{ij}	(1, 3)	(1, 3)	
Volatility	σ_i	0.10	0.12	
Drift	μ_i	0.01	0.01	
Discount rate	r	0.05		
Correlation	ho	0.3		
Panel B: Sensitivity Analysis				
Scenario 1: Change in Cost and Time	2			
1.A. Cost (removal, initialisation)	K_{ij}	(10,100)	(10,100)	-
1.B. Time (removal, initialisation)	T_{ij}	(3,8)	(3,8)	-
Scenario 2: Change in Volatility and	Drift			
2.A. Volatility	σ_i	0.16	0.18	?
2.B. Drift	μ_i	0.04	0.04	-
Scenario 3: Change in Discount Rate				
3.A Discount rate	r	0.12		-
Scenario 4: Change in Correlation				
4.A. Correlation	0	+ 0.8		?
T.71. Correlation	ρ	- 0.8		·

exercise frontier, and finally examine the frequency of switching in three models. Panel B in Table 3.1 shows four alternative scenarios that I examine in Section 3.3.4 in detail. In each scenario, I change one parameter at a time and evaluate its impact on the agents' optimal switching behaviour. Therefore, I evaluate different scenarios by having simulated GBM paths of product net income from equation (3.26). I begin my analysis by evaluating the behaviour of agents according to the baseline parameters from Table 3.1. Agents in my simulation start in state 1, and, in response to changes in product net income they switch their state of activity. These results provide a benchmark for further analysis when the underlying parameters change.

3.3.3 Optimal Exercise Frontiers

In this section, for a given set of parameters, I solve the valuation problem to find how much the enterprise is worth, calculate the optimal strategy, and finally, run a Monte Carlo simulation to see how often agents switch between products over an extended period. In this regard, I calibrate my proposed theory with a set of reasonable numbers from Table 3.1 to quantify enterprise redevelopment in three models: the cash-cost model, the projection model, and the cash-time.

Table 3.2 shows the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place, based on 65000 simulation paths over 200 years. I evaluate the switching behaviour between two competing products in three different models. In this regard, there will be some combinations where there is no switch in either direction. Also, there are some simulations where there is a switch from product 1 to product 2 but not a switch from product 2 to product 1. Finally, there are some simulations where multiple switches take place. The magnitude of switches in our results may indicate infre-

TABLE 3.2: Benchmark Switching Analysis

Note: This table presents the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place, based on 65000 simulation paths over 200 years. The choice of 65000 and 200 is ad hoc and it should be enough to evaluate the behaviour of agents. I bootstrap 10000 random samples with the same size of 65000 using sampling with replacement. I use this technique to test if the statistics of the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place (in the projection model vs the cash-cost model and the cash-time model vs the projection model) are quantitatively different. The superscripts ***, ***, and * indicate significance at the 1%, 5%, and 10% level respectively.

Variable	Cash-Cost Model	Projection Model	Cash-Time Model
Baseline Pa	rameters Analysis		
Mean	1.6049	1.0072***	0.9621***
$\mathbf{P}\ (n\geqslant 1)$	0.8233	0.7551***	0.7425***
$\mathrm{P}\left(n\geqslant2\right)$	0.3914	0.2011***	0.1804***

quent switches, due to the costly process of switching. However, the interest of this section is to emphasise the difference between the three models of switching no matter if they are different with frequent or infrequent switches. An important contribution of evaluating underlying factors in different scenarios is to introduce more relevant model of valuation according to different problems. Also, this differentiation among the three models helps to avoid underestimating or overestimating the switching of a problem.

Table 3.2 presents significance level for tests for difference in the reported statistics between the cash-cost and projection models as well as between the projection and cash-time models. In all cases, I reject the null of no difference at the 1% significance level. However, the difference between the cash-cost and projection models are quantitatively larger than the difference between the projection and cash-time models. The difference arises due to the measure of cost effect in each model. The cash-cost model takes into account only the cash cost of switching, whereas the pro-

jection model considers cash cost along with the expected monetary cost of time (in terms of forgone profits) and the cash-time model considers the cash cost along with the exact time cost. In other words, the cash-cost model captures only the cash cost whereas the projection and cash-time models capture not only the cash cost but also the time cost, (albeit, with different accuracy). In Section 3.3.4, I evaluate the changes in my baseline parameters, the effect on switch frequency, and the underlying reasons for my results.

The difference between the three models arises from the complexity of calculating switching cost as the net income of each product changes over the course of time. According to the set up and constraints, each model will have a different set of optimal exercise frontiers (OEFs). The OEFs trace the locus of product net income levels that will trigger an agent to switch products. The OEFs enable the models to provide normative advice on how enterprise should be used (i.e. when products should be changed). Figure 3.1 compares the OEFs for the three models according to baseline parameters in Table 3.1. Note that the OEFs for the projection model and the cash-time model are not qualitatively very different in Figure 3.1 due to the baseline parameters. Once I change the time taken to switch (see Section 3.3.4), I see the projection and cash-time model OEFs start to diverge. In other words, formal treatment of time, when time is a larger part of the cost, is very important.

Figure 3.1 displays the optimal exercise frontiers for two competing products in three models: the cash-cost model with the dotted lines, the projection model with the dash-dotted lines, and the cash-time model with the solid lines. The horizontal axis shows the product 1 net income and the vertical axis shows the product 2 net income. The area above the lower solid line in the cash-time model indicates no switch from product 2, whereas the area below the upper solid line shows no switch from product 1, and the area between two boundaries (solid lines) is an inactive

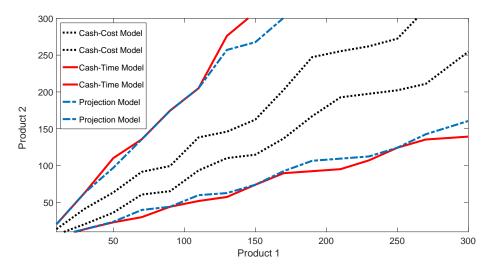


FIGURE 3.1: Optimal Exercise Frontier for Baseline Parameters.

Note: This figure displays the optimal exercise frontiers for two competing products in three models: the cash-cost model with the dotted lines, the projection model with the dash-dotted lines, and the cash-time model with the solid lines. The horizontal axis shows the product 1 net income and the vertical axis shows the product 2 net income. The area above the lower solid line in the cash-time model indicates no switch from product 2, whereas the area below the upper solid line shows no switch from product 1, and the area between two boundaries (solid lines) is an inactive region that agents will not start transition. Similarly, the same area can be defined for the cash-cost and projection models according to their boundaries, which are the dotted lines and the dash-dotted lines respectively. Note that the cash-cost model has the lowest gap among all three models due to capturing only the cash cost. The projection model is a good approximation of the cash-time model in this case because they are under the baseline parameters. The cash-time model starts to have a higher gap than the projection model when the time of switching increases and agents should go through a prolonged switching processes under uncertainty and reversibility.

region with no switch and agents will not start transition. Similarly, the same area can be defined for the cash-cost and projection models according to their boundaries, which are the dotted lines and the dash-dotted lines respectively.

The cash-cost model is oversimplified and agents would be expected to switch more often because they only need to pay a cash cost to switch (loss of product net income is ignored). Meanwhile, the projection model with the dash-dotted lines provides less chance for agents to switch due to accounting for the loss of product net income during the switch. In other words, agents should not only pay the cash cost in the cash-cost model but also pay a cost that proxies for the time cost, and is related to the current product values (see equation (3.12)). This opportunity cost is cheap for

products that are not worth much, but very expensive for products that are lucrative. Finally, the cash-time model, with the solid lines, fully and formally takes into account the time cost and has exercise frontiers with a wider gap. The cash-time model, compared with the other two models, makes it even harder to switch due to the extra time cost compared with the cash-cost model (and full time cost compared with the projection model that partially incorporates the time cost). Agents in the cash-time model have the option to abort the switching process and go back to the original product. However, they have to pay a cash cost and wait for a while to resume their initial activity. What makes the prolonged switching process more uncertain and perilous is that by the time agent complete the transition, cash flows may have changed substantially. Comparing all three models show that the time cost is proportional to product values whereas the cash cost does not scale according to product net income. Having said that, the OEFs in all three models are close to each other when product net income are low but further apart when the product net income are high. It is important to note that any changes in underlying parameters will cause the OEFs to get closer or move apart. For example, as the transition time increases, the gap between the projection and cash-time models increases more (see Figure 3.3), which is not significant with my baseline parameters in Figure 3.1. The projection model is much simpler and computationally cheaper for situations where the switching process is not so prolonged. Hence, the projection model leads us into implementing all types of analyses much easier. Conversely, when the transition time gets longer and we are not in the baseline scenario, the cash-time model is the most comprehensive model for my analyses and the results will be significantly different from the cash-time model.

The jaggedness of the OEFs is due to the discreteness in my finite difference grids. In other words, I discretise the PDE problem and solve it on specific product net income combinations, which are not continuous. Therefore, as I increase the number of steps and make my discrete points

closer, I will achieve smoother lines. The caveat is that the higher number of net income steps will increase the computational cost of solving the problem. Having said that, in some cases, for example, when I have negative correlation I will achieve smooth lines rather than jagged lines. The main reason for the jaggedness is the use of the transformation (see equation (3.16)). The negative correlation affects the covariance matrix and results in smooth lines due to using a different transformation. Figure 3.2 is an example of this situation using the parameters for my negative correlation sensitivity analysis in Section 3.3.4 (see Table 3.1). Figure 3.2 shows

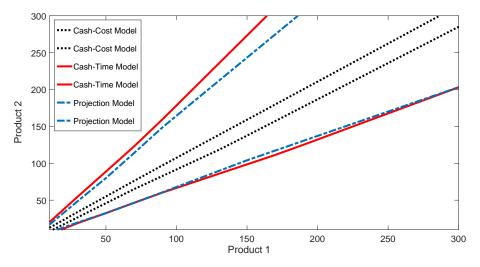


FIGURE 3.2: Optimal Exercise Frontier for Baseline Parameters but Negative Correlation.

Note: This figure displays the optimal exercise frontiers for two competing products in three models: the cash-cost model with the dotted lines, the projection model with the dash-dotted lines, and the cash-time model with the solid lines. The horizontal axis shows the product 1 net income and the vertical axis shows the product 2 net income. I note that the cash-cost model has the lowest gap among all three models due to capturing only the cash cost. The projection model is a good approximation of the cash-time model in this case because they are under the baseline parameters. The cash-time model starts to have a higher gap than the projection model when the time of switching increases and agents should go through a prolonged switching processes under uncertainty and reversibility. This figure how the jaggedness of figures disappear when I have a negative correlation. The reason refers to using eigenvalues in my transformation where the negative correlation affects the covariance matrix and results in smooth lines due to different transformation.

OEFs with smooth lines for three different models: the cash-cost model with the dotted lines, the projection model with the dash-dotted lines, and finally the cash-time model with solid lines.

3.3.4 Sensitivity Analysis

Cost and Time Sensitivity Analysis

In this section, I analyse the likelihood of switching when the baseline parameters change, as shown in Table 3.1. In this regard, I evaluate the impact of an increase in cost and time in two separate scenarios. In Scenario 1.A, I change the total cost of removing from \$1000 per ha to \$10000 per ha and the total cost of initialisation from \$10000 per ha to \$100000 per ha. In Scenario 1.B, I increase the time of removing from 1 year to 10 years and the time of initialisation from 3 years to 8 years. Then, in each scenario, I compare the switching frequency across my three models. Table 3.3 demonstrates the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place across my three models due to an increase in cost (Scenario 1.A) and time (Scenario 1.B) separately.

Table 3.3 reports the baseline parameter results for benchmark, Scenario 1.A (dollar cost increases), and Scenario 1.B (time cost increases). When I increase the cost of switching, I expect all three models to be affected by this effect, because all three models take into account the cash cost. I note that the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place, all decrease compared with my benchmark results. In other words, an increase in the cash cost makes switching expensive and pushes the OEFs further out, which in turn results in less switching. The difference between the cash-cost and projection models as well as the difference between the projection and cash-time models are significant at the 1% level, but the magnitude of the differences are not very large. However, in Scenario 1.B, where I increase the time cost (duration of switch) I note that the difference between model pairs starts to widen a lot more than when I change dollar costs.

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TABLE 3.3: Cost and Time Switching Analysis

Note: This table presents the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place based on 65000 simulation paths over 200 years. The choice of 65000 and 200 is ad hoc and it should be enough to evaluate the behaviour of agents. I use bootstrap method to generate 10000 random samples with the same size of 65000 and replacement. I use this technique to test if the statistics of the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place in the projection model vs the cash-cost model and the cash-time model vs the projection model are quantitatively/qualitatively different. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively.

Variable	Cash-Cost Model	Projection Model	Cash-Time Model						
Baseline Pa	Baseline Parameters Analysis								
Mean	1.6049	1.0072***	0.9621***						
$\mathrm{P}\left(n\geqslant1\right)$	0.8233	0.7551***	0.7425***						
$\mathbf{P}\ (n\geqslant 2)$	0.3914	0.2011***	0.1804***						
Scenario 1.A: Change in Cost									
Mean	0.6342	0.5674***	0.4765***						
$\mathbf{P}\ (n\geqslant 1)$	0.5484	0.5225***	0.4524***						
$\mathbf{P}\ (n\geqslant 2)$	0.0702	0.0412***	0.0227***						
Scenario 1.1	B: Change in Time								
Mean	1.6049	0.6480***	0.5330***						
$\mathbf{P}\ (n\geqslant 1)$	0.7383	0.5895***	0.5100***						
$\mathbf{P}\ (n\geqslant 2)$	0.2506	0.0551***	0.0225***						

Scenario 1.B reports that an increase in switching time decreases the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place in all three models (compared to my results in the benchmark). The difference across my three models are statistically significantly different at the 1% level. However, I expect that an increase in time effect provides a larger difference between my competing models when time is a much bigger part of the cost. In this case, the projection model is different from the cash-cost model due to larger time cost and the cash-time model is different from the projection model due to its more formal treatment of time, which becomes more important when the conversion process is longer. The projection model implies that an increase in the time cost leads to a higher approximation of opportunity cost that agents should pay, along with the cash cost. The cash-time model implicitly captures the opportunity cost by letting agents be in a "Transition" state where no products are produced. The longer the time, the bigger the difference gets between the projection and cash-time models. I demonstrate the optimal policies under Scenario 1.B for my three models in Figure 3.3.

Figure 3.3 shows the OEFs for three models. The cash-cost model is displayed with the dotted lines, the projection model with dash-dotted lines, and finally the cash-time model with solid lines. Figure 3.3 indicates that when the switching process entails more time, the optimal polices for the projection model and the cash-time model start to diverge and become significantly different from each other, but it has little effect on the cash-cost model (compared with Figure 3.1). The reason is that the cash-cost model captures only the cash cost and time of switching is not involved whereas the projection and cash-time models capture both, and the time cost becomes more crucial as the time taken for switching increases. On the one hand, agents could wait for a long time to complete the process and miss the current cash flows. On the other hand, uncertainty of product net income makes agents reluctant to exercise their options because

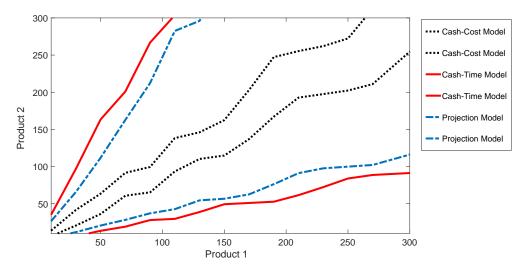


FIGURE 3.3: Optimal Exercise Frontier when Time Changes (Scenario 1.B).

Note: This figure shows the optimal exercise frontiers for three models. The cash-cost model is displayed with the dotted lines, the projection model with dash-dotted lines, and finally the cash-time model with solid lines. The horizontal axis displays the product 1 net income and the vertical axis displays product 2 net income. The cash-cost model has the lowest gap of optimal polices due to capturing only the cash cost. The projection model and the cash-time model have the higher gap compared with the cash-cost model due to capturing the cash cost and the time cost. However, the projection model is an approximation of the time cost whereas the cash-time model is the exact time cost. I note that in prolonged switching processes the cash-time model is the competent model.

they are not sure what product net income they will end up at the end of process and even if they reverse the process part way through it will not be quick to go back to the existing activity. Hence, the projection model approximates the impact of prolonged cost of switching whereas the cashtime model computes the exact time cost and results in wider optimal polices compared with the projection model. Overall, an increase in cost and time both have a negative impact on the mean number of switches and the probability of switches but the magnitude of changes varies according to the scenario and type of model.

Volatility and Drift Sensitivity Analysis

In this section, I change volatility and drift values for the net income process. Therefore, I evaluate the impact of each variable separately. In Sce-

TABLE 3.4: Volatility and Drift Switching Analysis

Note 1: This table presents the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place based on 65000 simulation paths over 200 years. The choice of 65000 and 200 is ad hoc and it should be enough to evaluate the behaviour of agents. I use bootstrap method to generate 10000 random samples with the same size of 65000 and replacement. I use this technique to test if the statistics of the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place in the projection model vs the cash-cost model and the cash-time model vs the projection model are quantitatively/qualitatively different. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively.

Variable	Cash-Cost Model	Projection Model	Cash-Time Model						
Baseline Pa	Baseline Parameters Analysis								
Mean	1.6049	1.0072***	0.9621***						
$\mathbf{P}\ (n\geqslant 1)$	0.8233	0.7551***	0.7425***						
$\mathrm{P}\left(n\geqslant2\right)$	0.3914	0.2011***	0.1804***						
Scenario 2.A	A: Change in Volatility								
Mean	1.5342	0.9465***	0.9416						
$P(n \geqslant 1)$	0.8107	0.6938***	0.6938						
$\mathrm{P}\left(n\geqslant2\right)$	0.3597	0.1920***	0.1941						
Scenario 2.E	3: Change in Drift								
Mean	2.4251	0.9016***	0.7171***						
$\mathbf{P}\left(n\geqslant1\right)$	0.8606	0.6862***	0.5949***						
$P(n\geqslant 2)$	0.5320	0.1729***	0.1067***						

nario 2.A, I increase the volatility of product 1 from 10% to 16% and product 2 from 12% to 18% and examine the results. In Scenario 2.B, I set back everything and increase the drift of product 1 from 1% to 4% and product 2 from 1% to 4% and see what is the impact of drift this time. Table 3.4 presents the results for changes and the corresponding impacts on the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place.

Table 3.4 shows the baseline parameter results as the benchmark, Scenario 2.A (increased volatility), and Scenario 2.B (increased drift). Scenario 2.A reports that an increase in volatility has a negative impact on the three measures of interest. In other words, my results in Scenario 2.A are quantitatively less than baseline, except the probability of two or more switches taking place in the cash-time model, which increases. The results across the three models show that the difference between the cash-cost and projection models is large and significant at the 1% level whereas the difference between the projection and cash-time models is almost zero and not significant. In this case, my results suggest that the projection model is a good approximation of the cash-time model. The projection model captures the impact of an increase in volatility by measuring the expected changes of the product net income as an extra cost that decreases switching compared with the cash-cost model. In contrast, the cash-time model measures the volatility as uncertainty in product net income that has a negative impact on switching. Agents in the cash-time model are concerned that product net income might move against their favour while they are in the process of removing and initialisation or even when they abort the switching process part way through to go back to their existing activity.

A higher volatility provides more time value to the options, and results in agents being more reluctant to exercise their options. In other words, agents are able to profit in the cash-time model because they get more value from uncertainty but the value of the option to switch products is so high that agents become reluctant to exercise their options. For the time being, I focus on the enterprise value of the cash-time model. To illustrate this point, I display the enterprise value when agents produce product 1 in Figure 3.4.

Figure 3.4 demonstrates the enterprise value with product 1 in the cash-time model. The enterprise value increases in both dimensions be-

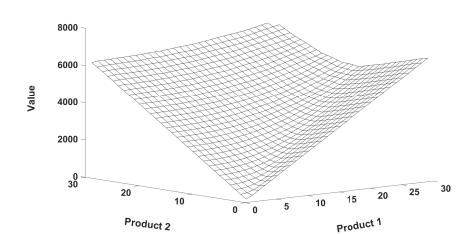


FIGURE 3.4: Enterprise Value with Product 1 in the cash-time model when volatility increases. **Note**: This figure shows the enterprise value in the cash-time model, producing product 1 when volatility increases in the baseline parameters. The enterprise value is on the z axis and the product net income are on the x and y axes. I show how agents switch from producing one type of product to another when the product net income change over the course of time.

cause either the product that the agent is producing is worth more or the option to switch to the competing product is producing in value. I observe that the enterprise value is curved in the middle due to the time value of the option to switch. The enterprise value is convex, which means there is option value. To better illustrate this convexity, I plot the secondary diagonal (lower left entry to the upper right entry) of enterprise values producing products 1 and 2 before (baseline) and after changes in volatility, which is displayed in Figure 3.5.

Figure 3.5 shows a cross-section of enterprise values producing products 1 and 2. The vertical axis displays the enterprise values and the horizontal axis displays the secondary diagonal of matrix values for each enterprise. The values under the baseline parameters are displayed with the dotted curve and under Scenario 2.A with the solid curves. Also, enterprises producing product 1 are displayed with the circle marker and enterprises producing product 2 with the star marker. All the values are evaluated using the cash-time model, because this model is able to evalu-

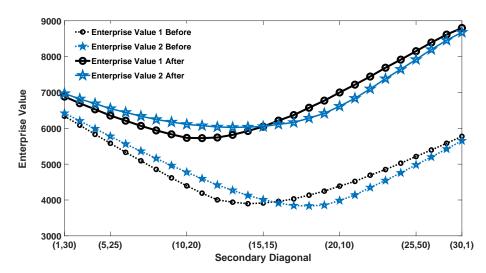


FIGURE 3.5: Enterprise Values Before (baseline) and After Change in Volatility in the cash-time model. **Note**: This figure shows the cross section of enterprise values producing products 1 and 2. The horizontal axis shows the secondary diagonal of matrix values and the vertical axis show the corresponding values of each coordinates. The values under the baseline parameters are displayed with the dotted curve and under Scenario 2.A with the solid curves. Also, enterprises producing product 1 are displayed with the circle marker and enterprises producing product 2 with the star marker.

ate prolonged real options to switch problems under uncertainty and reversibility and capture the exact switching cost. An increase in volatility shifts the curve up and to the left, and offers more option value. Higher volatility results in more uncertainty, but at the same time adds convexity to the curve that provides value to the switching option. In Figure 3.5, I observe that the right part of curve gains more option value than the left part when the volatility goes up. In other words, one of the products is really in the money and gives a higher value than the competing product and by exercising the switching option I can obtain a higher value.

Scenario 2.B in Table 3.4 reports the results for a change in drift. When the drift goes up, the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place in the cash-cost model increases (compared with the baseline) whereas they decrease in the projection model and the cash-time model. One explanation is that when drift goes up, agents will gain from the product returns due to the growth rate that convince them to switch.

In contrast, the drift (growth rate) is not enough to cover both the cash cost and the time cost in the projection and cash-time models. My results across my three models display that the difference between all my measures, comparing the cash-cost model and the projection model as well as comparing the projection model and the cash-time model, are significantly different at the 1% level. However, the actual difference between the cash-cost and projection models is larger than the difference between the projection and cash-time models. Overall, comparing the impact of drift and volatility may indicate that switches are more sensitive to drift than volatility. As I explained, the reason is that volatility has an option value effect (see Figure 3.5) that reduces the impact of an increase in volatility whereas a change in drift has a fundamental effect on values through discounted dividend model (DDM) that discounts the values by the difference between discount rate and drift $(CF_i/(\mu_i - r))$.

Discount Rate Sensitivity Analysis

In this section, I examine the impact of an increase in the discount rate on the frequency of switching. Therefore, I increase the discount rate from 5% to 12% and compare the results with my benchmark and across my three models. Table 3.5 presents the results for changes in switching frequency for agents when I increase the discount rate.

Table 3.5 shows that when the discount rate goes up, the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place, decreases in all three models. Changes in the discount rate have a two opposing effects and the outcome depends on which effect dominates. On the one hand, when the discount rate goes up, the time value of the option (regardless of the option type) goes down, and agents become impatient and prefer to pull the trigger and switch to the more profitable product. In other words, the value of future cash flows may not be worth as much as before and agents

TABLE 3.5: Discount Rate Switching Analysis

Note 1: This table presents the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place based on 65000 simulation paths over 200 years. The choice of 65000 and 200 is ad hoc and it should be enough to evaluate the behaviour of agents. I use bootstrap method to generate 10000 random samples with the same size of 65000 and replacement. I use this technique to test if the statistics of the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place in the projection model vs the cash-cost model and the cash-time model vs the projection model are quantitatively/qualitatively different. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively.

Variable	Cash-Cost Model	Projection Model	Cash-Time Model					
Baseline Parameters Analysis								
Mean	1.6049	1.0072***	0.9621***					
$P(n \ge 1)$	0.8233	0.7551***	0.7425***					
$P(n \geqslant 2)$	0.3914	0.2011***	0.1804***					
Scenario 3:	Change in Discount Ra	ite						
Mean	1.5366	0.9328***	0.6283***					
$P(n \geqslant 1)$	0.8124	0.7350***	0.5739***					
$P(n \geqslant 2)$	0.3700	0.1661***	0.0510***					

prefer to switch as soon as they get an opportunity that is profitable rather than waiting to see if they can achieve a higher profit. In the extreme case where the discount rate equals infinity, the time value of the option disappears completely and provides a myopic situation where agents prefer to exercise any option in hand at present without any regard for the future. On the other hand, when the discount rate goes up, agents' expectations of future returns tend to increase over the course of time. Hence, when they have any option to switch they are willing to wait and see if it gets even better. My results suggest that the latter effect outweighs the former, which means that an increase in discount rate has a negative impact on switch frequency, and agents prefer to wait more than to switch into an alternative product quickly.

Comparing my results across the three models shows that the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place decreases from the cash-cost model to the projection model, and from the projection model to the cash-time model. The difference between pair models is significant at the 1% level. I observe that the mean number of switches and the probability of two or more switches taking place has a larger difference between the cash-cost and projection models than between the projection and cash-time models, whereas the probability of one or more switches taking place has a larger difference between the projection and cash-time models than between the cash-cost and projection models. In other words, an increase in discount rates causes the agents to switch multiple times in the projection model than the cash-cost model and they prefer to have no switches or one switch in the cash-time model relative to the projection model.

Correlation Sensitivity Analysis

In this section, I examine the changes in correlation between the two competing products, net income. I evaluate the switch frequency when cor-

TABLE 3.6: Correlation Switching Analysis

Note 1: This table presents the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place based on 65000 simulation paths over 200 years. The choice of 65000 and 200 is ad hoc and it should be enough to evaluate the behaviour of agents. I use bootstrap method to generate 10000 random samples with the same size of 65000 and replacement. I use this technique to test if the statistics of the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place in the projection model vs the cash-cost model and the cash-time model vs the projection model are quantitatively/qualitatively different. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively.

Variable	Cash-Cost Model	Projection Model	Cash-Time Model							
Baseline Pa	Baseline Parameters Analysis									
Mean	1.6049	1.0072***	0.9621***							
$P(n \geqslant 1)$	0.8233	0.7551***	0.7425***							
$P(n \geqslant 2)$	0.3914	0.2011***	0.1804***							
Scenario 4.A	A: Change in Correlatio	on (Positive)								
Mean	1.0568	0.5076***	0.4378***							
$P(n \geqslant 1)$	0.7921	0.4990***	0.4334***							
$P(n \geqslant 2)$	0.1863	0.0085***	0.0044***							
Scenario 4.I	3: Change in Correlatio	on (Negative)								
Mean	5.5017	2.6163***	2.5173***							
$P(n \geqslant 1)$	0.9468	0.9052***	0.9060							
$P(n \geqslant 2)$	0.7962	0.6281***	0.6150***							

relation between product net income changes from +0.3 to +0.8 or -0.8. I report my results for changes in correlation in Table 3.6.

Table 3.6 reports the results for my benchmark, Scenario 4.A (a change in correlation in the positive direction), and Scenario 4.B (a change in correlation in the negative direction). Scenario 4.A shows that all three measures decrease when correlation goes up in the positive direction. I postulate that gaps in product net income persist. This means, the cash flows

from products move together in the same direction, which makes it more likely that if product A is more lucrative at time t it remains more lucrative after that. Evaluating the results across my three models shows that the difference between the cash-cost and projection models is larger than the difference between the projection and cash-time models, and all my results are significant at the 1% level.

Scenario 4.B demonstrates that the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place increase when the correlation is negative. In this case, the cash flows from products move in the opposite direction and when one product becomes less profitable, and the competing product becomes more lucrative. My results across the three models demonstrate that the difference between the cash-cost and projection models is larger than the difference between the projection and cash-time models. My results for the difference across pair models are significant at the 1% level except for the probability of one or more switches taking place between the projection and cash-time models.

Diff-in-Diff and Relative Value Changes Sensitivity Analysis

So far, I evaluated the significant changes between three different models (the cash-cost model, the projection model, and the cash-time model) by increasing baseline parameters one at a time. As a follow up, I examine if there is a significant change before and after change compared with the baseline parameters. Table 3.7 presents the summary results for all changes. Results show that the mean and the probability of one or more for all scenarios except the change in time for the cash-cost model are significantly different from the baseline model. The change in time will not affect the mean of the cash-cost model. The probability of two or more for the projection model are significant when there is a change in drift, discount rate and correlation. The probability of two or more for the cash-

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time model are significant when there is a change in cost, discount rate, and negative correlation.

In the following, I evaluate the sensitivity analysis when the relative values are changed compared with the baseline parameters. Table 3.8 reports the baseline results for benchmark, Scenario 5.A (dollar cost increases), Scenario 5.B (time cost increases), Scenario 5.C (volatility increases), and Scenario 5.D (drift increases). All values are increased for product 2 so we can examine the change in relative values. Therefore, the cost of removing product 2 increases from \$1000 to \$10000 and its cost to initialise increases from \$10000 to \$30000. The time of removing product 2 increases from 1 year to 3 years and its time to initialise increases from 3 years to 6 years. The volatility and drift of product 2 increases from 10 and 0.01 to 14 and 0.04 respectively.

The results for relative changes in Table 3.8 shows that our main results from the previous analysis holds and the projection model is significantly different from the cash-cost model, while the cash-time model is significantly different form the projection model. When the cost, time, volatility, and drift of product 2 increase, it is more costly to switch from one product to another and the mean number of switches, probability of one or more, and probability of two or more switches drop noticeably. Overall, as we increase the cost, time, volatility, and drift of one product alone, the switching process converges to zero and the level of being harder to switch increases as we move from the cash-cost to projection model and finally to the cash-time model.

¹ I test the difference across the three models when the parameters for product 2 change and for product 1 do not change.

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TABLE 3.7: Changes in the Switching Behaviour When the Baseline Parameters Changes in Each Scenario.

Note 1: This table presents the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place based on 65000 simulation paths over 200 years. I examine the diff-in-diff significance level between the baseline parameters and each scenario. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively.

Variable	Cash-Cost Model	Projection Model	Cash-Time Model	
Baseline Pa	rameters Analysis			
Mean	1.6049	1.0072	0.9621	
$P(n \ge 1)$	0.8233	0.7551	0.7425	
$P(n \geqslant 2)$	0.3914	0.2011	0.1804	
Scenario 1.	A: Change in Cost			
Mean	0.6342***	0.5674***	0.4765***	
$P(n \ge 1)$	0.5484***	0.5225***	0.4524***	
$P(n \geqslant 2)$	0.0702	0.0412	0.0227***	
Scenario 1.1	B: Change in Time			
Mean	1.6049	0.6480***	0.5330***	
$P(n \ge 1)$	0.7383	0.5895***	0.5100***	
$P(n \geqslant 2)$	0.2506	0.0551	0.0225	
Scenario 2.	A: Change in Volatility			
Mean	1.5342***	0.9465***	0.9416***	
$P(n \ge 1)$	0.8107***	0.6938***	0.6938***	
$P(n \geqslant 2)$	0.3597	0.1920	0.1941	
Scenario 2.1	B: Change in Drift			
Mean	2.4251***	0.9016***	0.7171***	
$P(n \geqslant 1)$	0.8606***	0.6862***	0.5949***	
$P(n \geqslant 2)$	0.5320***	0.1729***	0.1067	
Scenario 3:	Change in Discount Ra	te		
Mean	1.5366***	0.9328***	0.6283***	
$P(n \geqslant 1)$	0.8124***	0.7350***	0.5739***	
$P(n \geqslant 2)$	0.3700***	0.1661***	0.0510***	
Scenario 4.	A: Change in Correlatio	n (Positive)		
Mean	1.0568***	0.5076***	0.4378***	
$P(n \geqslant 1)$	0.7921***	0.4990***	0.4334***	
$P(n \geqslant 2)$	0.1863***	0.0085***	0.0044	
Scenario 4.1	B: Change in Correlation	•		
Mean	5.5017***	2.6163***	2.5173***	
$\mathrm{P}\left(n\geqslant1\right)$	0.9468***	0.9052***	0.9060***	
$P(n \geqslant 2)$	0.7962***	0.6281***	0.6150***	

TABLE 3.8: Changes in Switching Behaviour Across Three Models When Product 2 Parameters change.

Note 1: This table presents the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place based on 65000 simulation paths over 200 years. The choice of 65000 and 200 is ad hoc and it should be enough to evaluate the behaviour of agents. I use bootstrap method to generate 10000 random samples with the same size of 65000 and replacement. I use this technique to test if the statistics of the mean number of switches, the probability of one or more switches taking place, and the probability of two or more switches taking place in the projection model vs the cash-cost model and the cash-time model vs the projection model are quantitatively/qualitatively different. The superscripts ***, **, and * indicate significance at the 1%, 5%, and 10% level respectively.

Variable	Cash-Cost Model	Projection Model	Cash-Time Model					
Baseline Pa	rameters Analysis							
Mean	1.6049	1.0072***	0.9621***					
$P(n \geqslant 1)$	0.8233	0.7551***	0.7425***					
$P(n \geqslant 2)$	0.3914	0.2011***	0.1804***					
Scenario 5.A	A: Change in Cost							
Mean	0.0959	0.0041***	0.0000***					
$P(n \ge 1)$	0.0766	0.0020***	0.0000***					
$P(n \geqslant 2)$	0.0030	0.0001***	0.0000					
Scenario 5.I	3: Change in Time							
Mean	1.6049	0.2352***	0.0178***					
$P(n \geqslant 1)$	0.0593	0.0011***	0.0000***					
$P(n \ge 2)$	0.0127	0.0000***	0.0000					
Scenario 5.0	C: Change in Volatility							
Mean	0.7769	0.6487***	0.1148***					
$P(n \geqslant 1)$	0.2715	0.2706***	0.0009***					
$P(n \geqslant 2)$	0.2414	0.1910***	0.0004***					
Scenario 5.I	Scenario 5.D: Change in Drift							
Mean	0.5780	0.5164***	0.0581***					
$P(n \ge 1)$	0.0692	0.0992***	0.0000***					
$P(n \ge 2)$	0.0259	0.0134***	0.0000***					

3.4 Conclusion

This chapter develops a framework for analysing the behaviour of agents who have the option to produce different products. Agents change between different products as the net income of competing products changes over time. I set up a partial differential equation model and then solve it with the Alternating Direction Implicit approach. Switching products entails a waiting cost and agents miss their cash flows during this period. The projection and cash-time models provide an optimal solution for options to switch by capturing not only the cash cost but also the time cost.

The cash-cost model is what I see in traditional literature (e.g. Song, Zhao, and Swinton (2011), Livingston, Roberts, and Y. Zhang (2014), Hauer, Luckert, Yemshanov, and Unterschultz (2016), and Hossiso and Ripplinger (2017)) that evaluates a situation where agents pay a cash cost and switch to a new state of activity. The cash-time model enables agents to move to a transition state initially, which allows them to switch back to the original product, if the cash flows change in its favour. In this model, I examine the time effect very carefully and find the exact time cost. Agents in this model must wait for a while in "Transition" state before they become active in a new state of activity. However, agents have the option to abort the switching process and go back to the original product. The projection model is an approximation for the cash-time model where agents have to pay the cash cost but they take into account the opportunity cost and missing the cash flows from existing products. Therefore, agents consider the cash cost and a monetary proxy for the time cost (in terms of foregone profits), for switching to a new state of activity.

My study shows that the cash-cost model that is used throughout the literature overestimates switching frequency. The projection model presents better results compared with the cash-cost model, but underperforms compared with the cash-time model due to approximating the time cost. The cash-time model computes the exact time cost and its results are significantly different from the projection model when the time of waiting is substantial. The sensitivity results from the cash-time model are significantly different from the projection model at the 1% level in all scenarios except in the case when volatility is high. My results indicate that increases in cost, time, volatility, drift, and discount rate have negative impacts on the switch frequency. Correlation has a negative impact on switch frequency if it is positive but a positive impact if it is negative.

The results for this study are important and useful in theory and practice. In theory, I contribute to the literature on valuation, especially the real options approach, by presenting a novel formulation of real options to switch. Researchers can use the results from my study for their analysis or to develop the idea for further exploration. In practice, economics agents, policy makers, analysts, and advisors in industries with similar natures may find my methodology useful.

Chapter 4

Orchards Dataset Analysis

4.1 Data Statistics

The focus of Chapter 4 is to evaluate apple orchards in New Zealand, specifically in Hawke's Bay region. This region is quite well suited to orchards, and growing apples is a lucrative, attractive, and competitive activity. Ministry of Primary Industry reports show that the Hawke's Bay orchards have the highest production and profit for the pipfruit industry in New Zealand (MPI, 2017). In the following, I evaluate orchards and the corresponding statistics.

The apple dataset includes 59197 (orchard-time) observations with 1098 specific identification codes over time period 2010-15. Each identification code is allocated to a specific parcel of land owned by an individual orchardist. For example, orchardist A might have 25 identification codes that are different either in terms of land size or the year of examination. It is worth to note that orchardists might have a portfolio of apple varieties that affects on their decisions to switch when the apple prices change over time. The minimum and maximum land sizes are 0.01 and 22.58 hectare respectively. Apple types are classified into 20 variety groups where similar apple varieties are merged into one variety. Hence, the apple variety groups shrink from 20 to 14 groups. Table 4.1 presents the amount of total land in hectares that is allocated to each variety group.

Table 4.1 shows that less land is allocated to Nashi, Beauty, and Cox with 0.24%, 1.25%, and 2.08% respectively and more land to apple groups Gala, Braeburn, Jazz, and Fuji with 26.98%, 17.21%, 11.12%, and 10.43% respectively. This specification is one of the criteria that help to choose the right apple type to consider in the switching model in Chapter 5 model.

Table 4.2 shows how much land is allocated to apple varieties in each

¹ For example, High Colour Gala and Gala as Gala, High Colour Braeburn and Braeburn as Braeburn, High Colour Fuji and Fuji as Fuji, and finally High Colour Cripps Pink, Cripps Pink, and Pink Lady as Cripps.

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TABLE 4.1: Allocated Land to apple varieties in New Zealand

Note: This table shows the allocated land in hectares to 14 apple varieties in New Zealand. Less land is allocated to Nashi, Beauty, and Cox with 0.24%, 1.25%, and 2.08% respectively and more land to apple groups Gala, Braeburn, and Jazz with 26.98%, 17.21%, and 11.12% respectively. There are 20 apple varieties that shrink to 14 varieties due to combining varieties within the same class.

Variety	Total Land	Percentage
Nashi	114.68	0.24%
Beauty	587.18	1.25%
Cox	973.53	2.08%
Granny Smiths	1370.47	2.92%
Envy	1460.48	3.12%
Rose	2030.54	4.33%
Pears	2195.07	4.68%
Cripps Pink	2287.26	4.88%
Queen	2514.35	5.36%
Other	2522.30	5.38%
Fuji	4890.94	10.43%
Jazz	5211.63	11.12%
Braeburn	8068.21	17.21%
Gala	12649.43	26.98%

TABLE 4.2: Allocated Land to apple varieties in different Regions in New Zealand

Note: This table shows the amount of allocated land in hectares to apples in different regions in New Zealand. There are 13 regions around the country that are focused on growing apple varieties. Bay of Plenty, Manawatu, and Auckland regions have the lowest land (with 0.06%, 0.11%, and 0.16% respectively) and Hawke's Bay, Nelson, and Otago regions have the largest amount of allocated land (with 61.29%, 28.76%, and 4.68% respectively) for growing apple varieties out of 100% total land allocated for growing apple in New Zealand.

Variety	Total Land	Percentage
Bay of Plenty	29.97	0.06%
Manawatu/Wanganui	51.98	0.11%
Auckland	73.47	0.16%
Horowhenua	111.67	0.24%
Marlborough	118.58	0.25%
South Canterbury	137.56	0.29%
Canterbury	240.96	0.51%
Wairarapa	430.77	0.92%
Gisborne	560.93	1.20%
Waikato	714.64	1.52%
Otago	2194.81	4.68%
Nelson	13482.24	28.76%
Hawke's Bay	28728.49	61.29%

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region. Bay of Plenty, Manawatu, and Auckland regions have the lowest land (with 0.06%, 0.11%, and 0.16% respectively) and Hawke's Bay, Nelson, and Otago regions have the largest amount of allocated land (with 61.29%, 28.76%, and 4.68% respectively) for growing apple varieties out of 100% total land allocated for growing apple in New Zealand. The apple dataset also gives information about the year that trees are planted. This is important, since it allows us to measure the apple trees' age.

The number of observations in the dataset decreases from 59197 (1098 IDs) to 50504 (1027 IDs) observations after filtering all regions by deleting observations with zero land area, missing year planted, and finally merging similar apple varieties into one variety. The orchardists with blocks of land plant trees with different apple types. Some orchardists might split a large block of land to small blocks or combine small blocks to a large block of land. In this regard, it is hard to recognise if this is exactly the same block of land that is under examination. To overcome this difficulty, I create a new identification code for orchards by combining parcel identification code, planting year, apple type, and the area of land that is allocated each year to track orchardists' activity through time. Therefore the new identification code is a parcel-time code and it becomes clearer which identification code is traceable over time by controlling key factors over time. The new identification code breaks down the total 50504 (1027 IDs) observations into 38597 (11334 IDs) traceable observations and 11907 (4902 IDs) non-traceable observations. The dataset with traceable observations provides a robust examination of orchardists' behaviour. I create a dummy variable for apple varieties to examine how the apple type changes for each specific orchardist in traceable dataset. If the apple type for the new identification code remains the same, the dummy variable equals zero, and otherwise it equals one. Hence, there are, 1,950 (440 IDs) observations that switch from one type of apple to another. Table 4.3 maps all the switches to discover the switch direction and orchardists' preferences towards each apple variety. Results show that there are 420

TABLE 4.3: The Number of Switches Among Apple Varieties in All Regions.

Note: This table shows the number of switches among apple varieties in all regions in New Zealand. Orchardists switch from one type of apple to another since the profitability of an apple variety changes over time. The total number of switches From and To each apple variety are reported in the last column and row respectively.

ion									То							
Direction	Variety	Braeburn	Queen	Gala	Fuji	Beauty	Envy	Cripps	Others	Jazz	Pears	Rose	Cox	Granny	Nashi	Total
	Braeburn	0	32	14	18	0	1	3	5	3	3	4	0	2	0	85
	Queen	1	0	3	0	0	0	2	1	0	2	3	0	0	0	12
	Gala	9	12	0	6	0	6	4	3	9	0	6	0	3	0	58
	Fuji	5	27	21	0	0	11	3	4	4	0	4	1	1	0	81
_	Beauty	0	3	4	2	0	4	0	1	0	0	1	0	0	0	15
From	Envy	0	0	4	1	0	0	1	0	2	0	0	0	0	0	8
	Cripps	1	6	3	3	0	1	0	1	2	0	0	0	0	0	17
	Other	6	7	27	7	0	3	3	0	0	6	2	1	1	1	64
	Jazz	2	7	10	9	0	4	1	1	0	1	0	0	1	0	36
	Pears	0	0	1	0	0	0	0	3	0	0	2	0	0	0	6
	Rose	1	2	12	0	0	0	0	2	0	2	0	0	0	0	19
	Cox	4	3	1	2	0	2	0	1	0	0	1	0	0	0	14
	Granny	0	0	4	0	0	0	0	0	0	0	1	0	0	0	5
	Nashi	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
	Total	29	99	104	48	0	32	17	22	20	14	24	2	8	1	420

IDs that switch once from one type of apple to another and 20 IDs that switch multiple times. Table 4.3 shows that in total, less orchardists switch *from* Nashi, Granny Smiths, and Pears with 0, 5, and 6 number of switches respectively and more *from* Braeburn, Fuji, and Other with 85, 81, and 64 number of switches respectively. In contrast, less orchardists switch to Beauty, Nashi, and Cox with 0, 1, and 2 number of switches respectively and more *to* Gala, Queen, and Fuji with 104, 99, and 48 number

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TABLE 4.4: Allocated Land Based on Switches in All Regions

Note: This table shows the amount of allocated land in hectares and the number of switches across different regions in New Zealand. Switches occur in 8 different regions: Hawke's Bay, Nelson, and Otago with 359.00, 71, and 15 number of switches respectively. The amount of allocated land and the number of switches follow the same pattern, the higher allocated land, the higher number of switches.

Note: This table is different from Table 4.2, because it includes only the land size with switches.

Allocated Land and Switches Across Regions								
Region	Total Land	Percent	Switch	Percent				
Marlborough	0.92	0.06	1.00	0.22				
Horowhenua	2.76	0.18	1.00	0.22				
Waikato	4.67	0.30	2.00	0.43				
Wairarapa	17.86	1.14	8.00	1.73				
Gisborne	33.89	2.16	5.00	1.08				
Otago	44.74	2.85	15.00	3.25				
Nelson	391.44	24.95	71.00	15.37				
Hawke's Bay	1072.43	68.37	359.00	77.72				

of switches respectively. Among all combinations, results show that orchardists switch mainly from Braeburn to Queen, from Fuji to Queen, and from Other to Gala with 32, 27, and 27 number of switches.

I examine the number of switches according to the regions and corresponding allocated land. Table 4.4 shows that approximately 70% of allocated land and 80% of the switches belongs to Hawke's Bay. Nelson with approximately 25% of allocated land and 15% number of switches is the second important region. The rest of regions may be ignored given the allocated land along with the low number of switches. The number of switches in Table 4.4 includes both single and multiple switches.

The dataset evaluation emphasizes the importance of Hawke's Bay region to be the target regions due to having the highest amount of allocated land to apple varieties and higher number of switches. The number

TABLE 4.5: Allocated Land to apple varieties in Hawke's Bay

Note: This table shows the allocated land in hectares to apple in Hawke's Bay. Nashi, Cox, and Beauty with 0.02%, 0.22%, and 1.60% respectively, have the least allocated land, and Gala, Braeburns, and Fuji with 30.21%, 16.32%, and 13.16% respectively, have the largest amount of allocated land to apple varieties.

Allocated Land to Apple Varieties						
Obs	Variety Group	Total Land	Percent			
1	Nashi	5.24	0.02			
2	Cox	63.91	0.22			
3	Beauty	460.36	1.60			
4	Envy	614.57	2.14			
5	Pears	730.55	2.54			
6	Granny Smiths	978.38	3.41			
7	Other	1045.35	3.64			
8	Cripps Pink	1504.04	5.24			
9	Rose	1773.75	6.17			
10	Jazz	2121.08	7.38			
11	Queen	2285.23	7.95			
12	Fuji	3780.25	13.16			
13	Braeburn	4687.77	16.32			
14	Gala	8678.05	30.21			

of observations for Hawke's Bay region is 34237 (640 IDs) after implementing the same filtering process that I had for all regions. Table 4.5 presents the amount of allocated land to each apple variety in Hawke's Bay region.

Table 4.5 shows that Nashi, Cox, and Beauty with 0.02%, 0.22%, and 1.60% respectively, have the least allocated land, and Gala, Braeburns, and Fuji with 30.21%, 16.32%, and 13.16% respectively, have the largest amount of allocated land among competitive apple varieties. I follow the same filtering process to calculate the number of switches whose land parcels are

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traceable and map them among apple varieties. The first row and column in Table 4.6 show the switching direction and the last row and column show the total number of switches *to* and *from* respectively. Results in Table 4.6 show that there are 326 single switches from one type of apple to another and 16 multiple switches from one type of apple to another and then to another apple type that raise the total number of switches to 342.

Table 4.6 shows that orchardists in Hawke's Bay, switch from one type of apple to another. The total number of switches *from* and *to* each apple variety are reported in the last column and row respectively. In terms of total switches, less orchardists switch *from* Cox, Pears, and Granny Smiths with 3, 3, and 4 switches respectively and more *from* Braeburn, Fuji, and Gala with 81, 59, and 47 switches respectively. In contrast, less orchardists switch *to* Beauty, Nashi, and Granny Smiths with 0, 1, and 5 number of switches respectively and more *to* Queen, Gala, and Fuji with 94, 79, and 41 switches respectively. Orchardists switch mainly from Braeburn to Queen, from Fuji to Queen, and from Other to Gala with 32, 26, and 17 switches respectively. The switching pattern among apple varieties in Hawke's Bay is the same as the switching pattern in all regions in Table 4.3. This implies that Hawke's Bay are representative of the industry more generally.

Table 4.7 reports the price path of apple varieties from 2007-16. There is a longer historical apple price series compared with the orchardists' dataset.¹ Apple prices are reported in dollar per Tray Carton Equivalent (TCE) and includes 12 types of apples. The missing types (Envy and Nashi) are deleted due to having missing values.²

¹ The table of apple prices is good to find the parameters and the more the better. Unfortunately we have a short historical data. To evaluate the behaviour of orchardist we have only the data for 2010-2015.

² A TCE is defined as 18.6 kg packed weight and equates to 18.0 kg sale weight.

TABLE 4.6: The Number of Switches Among Apple Varieties in Hawke's Bay

Note: This table shows the number of switches among apple varieties in Hawke's Bay. Orchardists switch from one type of apple to another since the profitability of an apple variety changes over time. The total number of switches From and To each apple variety are reported in the last column and row respectively. In terms of total switches, less orchardists switch from Cox, Pears, and Granny Smiths with 3, 3, and 4 number of switches respectively and more from Braeburn, Fuji, and Gala with 81, 59, and 47 number of switches respectively. In contrast, less orchardists switch to Beauty, Nashi, and Granny Smiths with 0, 1, and 5 number of switches respectively and more to Queen, Gala, and Fuji with 94, 79, and 41 number of switches respectively. Orchardists switch mainly from Braeburn to Queen, from Fuji to Queen, and from Other to Gala with 32, 26, and 17 number of switches respectively.

ion							То)							
Direction	Variety	Braeburn	Queen	Gala	Fuji	Beauty	Envy	Cripps	Others	Jazz	Pears	Rose	Granny	Nashi	Total
	Braeburn	0	32	14	18	0	1	2	4	3	1	5	1	0	81
	Queen	1	0	3	0	0	0	2	1	0	2	3	0	0	12
	Gala	8	11	0	5	0	5	4	1	6	0	5	2	0	47
	Fuji	3	26	16	0	0	2	2	3	4	0	2	1	0	59
Щ	Beauty	0	3	3	2	0	4	0	1	0	0	0	0	0	13
From	Envy	0	0	2	1	0	0	1	0	1	0	0	0	0	5
	Cripps	1	7	3	3	0	0	0	0	2	0	0	0	0	16
	Other	0	6	17	3	0	2	2	0	0	4	2	0	1	37
	Jazz	1	7	8	8	0	4	0	0	0	1	0	1	0	30
	Pears	0	0	0	0	0	0	0	1	0	0	2	0	0	3
	Rose	1	2	10	0	0	0	0	2	0	1	0	0	0	16
	Cox	2	0	0	1	0	0	0	0	0	0	0	0	0	3
	Granny	0	0	3	0	0	0	0	0	0	0	1	0	0	4
	Total	17	94	79	41	0	18	13	13	16	9	20	5	1	326

TABLE 4.7: Apple Prices Over Time Period 2006-16.

Note: This table shows the apple prices for 12 apple varieties over time period 2006-16. Prices are adjusted for inflation using Consumer Price Index (CPI) for fruits and vegetables group. All prices are reported in \$/tray carton equivalent that increase after year 2012. However, some apple varieties (e.g. Queen, Rose, and Pears) experience higher prices command with some apple varieties.

Variety	2006	2007	2008	2009	2010	2011	2012	2013	2014	2015	2016
Braeburn	18.74	19.78	20.79	20.68	20.13	17.46	19.06	22.29	24.72	25.40	27.86
Cox	19.19	20.02	21.54	19.70	18.95	17.10	17.16	21.90	26.49	30.01	33.11
Cripps Pink	19.96	22.14	21.65	23.57	22.78	19.76	21.56	26.40	29.80	31.84	34.49
Fuji	21.87	22.51	23.35	26.20	24.67	22.15	24.62	27.07	32.73	33.65	40.14
Granny Smith	19.92	21.42	22.73	25.44	25.59	23.06	24.86	30.17	29.34	32.17	32.38
Jazz	17.93	18.94	18.98	19.33	20.63	19.42	22.41	29.34	32.03	31.24	32.90
Beauty	20.97	23.79	24.03	35.78	29.53	27.88	28.30	30.80	32.20	33.11	38.00
Queen	20.81	24.24	26.74	36.14	31.97	29.86	32.42	36.28	40.78	47.92	42.73
Rose	22.56	23.64	25.96	30.39	31.54	29.54	33.54	37.25	39.63	44.04	37.88
Royal Gala	19.78	21.43	20.94	24.09	23.74	21.20	23.36	26.00	28.12	26.79	33.11
Other	21.48	20.19	24.09	29.42	24.22	21.39	29.39	36.82	40.52	43.36	53.09
Pears	24.24	26.34	26.55	29.31	31.38	31.76	32.11	33.80	36.59	37.74	37.85

Table 4.7 shows that the apple prices tend to increase over time. Queen, Rose, and Pears experience the highest prices whereas Braeburn and Cox experience the lowest prices among competitive apple varieties. Figure 4.1 displays the apple prices path from 2006 to 2016.

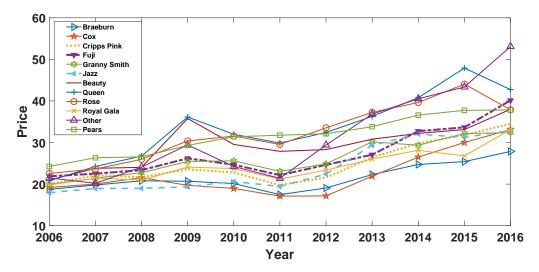


FIGURE 4.1: Apple Prices Path Over Time Period 2006-16.

Note: This figure shows the apple prices path over time period 2006-16. The horizontal axis shows the years and the vertical axis shows the apple prices, which are adjusted for inflation by CPI. Queen, Rose, and Pears experience the highest prices whereas Braeburn and Cox experience the lowest prices among competitive apple varieties.

Among all apple varieties, Braeburn and Queen are two candidates to be examined. Figure 4.2 shows the ratio of Braeburn price to Queen price over time period 2006-16. The ratio of these two apple prices has been volatile with decreasing trend over this time period. In the following, I will describe how Braeburn and Queen apple types are the best candidates for the set up model.

Apple types Queen, Beauty, and Other are more volatile over time whereas Jazz, Pears, and Braeburn are less volatile. According to the Ministry of Primary Industry (MPI) Model for Hawkes Bay, orchardists pick 3104 tray carton apples in each hectare, which gives a reference to scale the apple prices per hectare.¹ To avoid any biases and random choice of ap-

¹ MPI reports two pipfruit models: Hawkes Bay pipfruit model and Nelson pipfruit

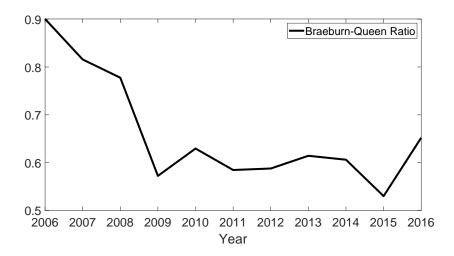


FIGURE 4.2: Ratio of Braeburn Price to Queen Price Over Time Period 2006-16. **Note**: This figure shows the ratio of Braeburn apple price to Queen price over time period 2006-16. The horizontal axis shows the years and the vertical axis shows the apple prices ratio.

ple variety for my model, I exploit a classification technique to narrow the choice from 15 different apple varieties into two variety groups. Hence, I use the Principal Component Analysis and K-means clustering as two unsupervised clustering methods. The Principal Component Analysis tries to find the trends and patterns in the dataset by reducing dimensions into a limited number of principal components. K-means (like Principal Component Analysis) clusters by putting all observations into K cluster by optimising the mean between each cluster. The apple prices are based on dollar per hectare (\$/ha), which is the product of prices (\$/tce) with 3104 trays for each hectare. Also, I normalise the apple prices by subtracting their mean and dividing by their standard deviation to avoid biases in clustering. The clustering technique with the aim of having two clusters specifies that Braeburn, Cripps Pink, Granny Smiths, Jazz, Queen, Rose, and pears belong to Cluster-1 whereas Fuji, Beauty, Royal Gala, and other belong to Cluster-2. My model is set up for a switching option with two

model. These two regions are the main pipfruit growing in New Zealand (see Table 4.4). Each model is based on the valuation of production and profitability for 40 hectares planted orchards in each region. The average 3104 TCE/ha is for any type of apple.

products and requires estimation of drifts, volatilities, and correlation parameters. I estimate drift and volatility of log revenues under the assumption of a GBM as follows:

$$\mu_{i} = \frac{\widehat{\alpha}_{i}}{\Delta t} + \frac{1}{2} \left(\frac{\widehat{\phi}_{i}}{\sqrt{\Delta t}} \right)^{2}$$

$$\sigma_{i} = \frac{\widehat{\phi}_{i}}{\sqrt{\Delta t}},$$
(4.1)

where μ_i is drift, $\widehat{\alpha_i}$ is the mean of changes in log revenues, Δt is time increment, and $\widehat{\phi_i}$ is the standard deviation of changes in log revenues. The first method is to take the average of apple prices in each cluster that accounts for the corresponding apple types 1 and 2. Using the average log prices in Table 4.8 and plugging them into equation (4.1) gives drifts of 5% and 6%, volatilities of 7% and 11% for apple clusters 1 and 2 respectively, and 0.56 correlation between the two apple clusters. The result of considering the average log prices in each cluster is that the high volatility apple types get affected by less volatile types. Hence, the model would be expected to reveal less switching when using drift, volatility, and correlation parameters.

TABLE 4.8: Clustering Apple Varieties.

Note: This table shows the clustering of apple varieties over time period 2006-16. I take the average prices per hectare by multiplying apple prices to the amount of tray carton apple in each hectare (e.i. 3104 based on MPI Model). The apple prices in this table are scaled by 1000. The difference of log apple prices are calculated for each cluster to be used for finding drifts, volatilities, and correlation.

Cluster	2006	2007	2008	2009	2010	2011	2012	2013	2006 2007 2008 2009 2010 2011 2012 2013 2014	2015	2016
Average Prices for Each Cluster	rices fo	r Each C	Cluster								
Cluster-1 63.92 69.40 72.46 81.97 81.60 75.76 82.46 95.57 103.27 111.01 109.13	63.92	69.40	72.46	81.97	81.60	75.76	82.46	95.57	103.27	111.01	109.13
Cluster-2 65.26 68.23 71.70 89.61 79.28 71.88 82.00 93.67 103.65 106.24	65.26	68.23	71.70	89.61	79.28	71.88	82.00	93.67	103.65	106.24	127.53
Log Average Prices for Each Cluster	ge Price	es for E	ach Clu	ster							
Cluster-1 4.16 4.24 4.28 4.41 4.40 4.33 4.41	4.16	4.24	4.28	4.41	4.40	4.33	4.41	4.56	4.64	4.71	4.69
Cluster-2 4.18 4.22 4.27 4.50 4.37 4.28 4.41	4.18	4.22	4.27	4.50	4.37	4.28	4.41	4.54	4.64	4.67	4.85
Log Average Prices Difference for Each Cluster	ge Price	es Diffe	rence fo	r Each	Cluster						
Cluster-1 0.00 0.08 0.04 0.12 0.00 -0.07 0.08	0.00	0.08	0.04	0.12	0.00	-0.07	0.08	0.15	0.08	0.07	-0.02
Cluster-2 0.00 0.04 0.05 0.22 -0.12 -0.10 0.13 0.13	0.00	0.04	0.02	0.22	-0.12	-0.10	0.13	0.13	0.10	0.02	0.18

In the second approach, I use the *maximum* price among all apple prices in cluster-1 as the first price series and the minimum price among all apple prices in cluster-2 as the second price series over time. Hence, by having the two apple price series I calculate the drift, volatility, and correlation variables of model. Similarly, I find the variables by using the minimum price among apple prices in cluster-1 and maximum price among all apple prices in cluster-2 over time. All the apple prices are in terms of log price. In this approach, the results refer to a group of apple types instead of a particular type. An alternative approach is to select a particular apple type with the highest price versus a particular apple type with the lowest price. This approach removes the ambiguity of having a group of apple type as a whole. Queen, Gala (with high prices), and Braeburn (with low prices) are good candidates for this approach. Queen has 8% drift and 13% volatility, Gala has 6% drift and 10% volatility, and finally Braeburn has 4% drift and 8% volatility. The correlation for the Queen-Braeburn pair is 0.26, whereas the correlation for the Gala-Braeburn pair is 0.68. Queen and Braeburn with different volatilities are better candidates than the Gala and Braeburn pair with the same volatilities. Hence, the focus of study narrows down into two types of apple with more clarity. In the next chapter, I set up my model for two options to switch and this simplification helps to select only two suitable candidates for examination.

The final test on the apple dataset is to look at the age of the apple trees. In this regard, I select the traceable dataset in all regions and check if the planted year changes with the change in apple variety. If the planted year changes with the change in apple variety, it implies that orchardists replace the existing apple tree with another variety. In contrast, if the planted year remains the same but the type of apple changes, it implies grafting another apple type on the existing rootstock. There are 1330 (238 IDs) observations that account for replacing and 620 (103 IDs) observations that account for grafting. Table 4.9 shows how orchardists change the apple type either by uprooting and planting young apple trees, or by grafting

TABLE 4.9: Age Groups of Switches in All Region.

This table presents the orchardists who switch by uprooting and planting new apple trees or by grafting on the existing rootstocks. The number of switches and the corresponding percentage are reported in different columns. The strategy of switching is categorised according to the age of apple trees. The uproot-plant strategy is more likely in age categories 20-25 and 8-20 years with 81.45% and 75.41% respectively. In contrast, the grafting strategy is more likely in age categories 0-8, 8-20, and 20-25 years with 58.14%, 24.59%, and 18.55% respectively.

Age Group	Number of Ol	bservations	Percentage of C	bservations
3 1	Uproot-Plant	Graft	Uproot-Plant	Graft
0–8	36	50	41.86%	58.14%
8–20	92	30	75.41%	24.59%
20–25	101	23	81.45%	18.55%
> 30	9	0	100.00%	00.00%
Total	238	103	69.79%	30.21%

on existing trees as a function of tree's age, with old trees getting uprooted. The number of switches and the corresponding percentage are reported in different columns. The strategy of switching is categorised according to the age of apple trees. The age is measured by finding the difference between the current year and planted year of apple trees. The uproot-plant strategy is more likely in age categories 20-25 and 8-20 years with 81.45% and 75.41% respectively. In contrast, the grafting strategy is more likely in age categories 0-8, 8-20, and 20-25 years with 58.14%, 24.59%, and 18.55% respectively.

As a follow up, I examine if the uproot-plant and the graft strategies are statistically significantly different according to the apple trees' age. Hence, I implement a two sample t-test with unequal variance to see if the uproot-plant and the graft samples (two switching strategies) are significantly different when the tree's age changes. My intuition is that the uproot-plant strategy is employed for old apple trees, whereas the graft strategy is ideal for young trees. Table 4.10 shows the results of the t-test with a null hypothesis of having the same mean between two samples

TABLE 4.10: t-test Results for Switches Based on Trees' Age.

This table presents the t-test between the apple trees' age in two samples: uproot-plant and graft. There are 238 orchardists who employ the uproot-plant strategy, where the trees' age have the mean 19.1218 with 7.5578 standard deviation. In contrast, there are 103 orchardists who employ the graft strategy, where the trees' age have the mean 12.2788 with 8.6546 standard deviation. I use a t-test with unequal variance to examine my null hypothesis, whether the difference in means is 0. According to the t-test results, the t-value is greater than the critical value and I reject the null hypothesis. In other words, these two switching strategies are employed at different ages and each strategy is ideal for a specific tree's age. Note that apple trees have ages from 1 year old to 33 years old and more.

Parameters	Uproot-Plant Age	Graft Age
Observations	238	103
Ages Mean	19.12	12.28
Ages Variance	57.12	74.90
Ages Standard Deviation	7.56	8.66
Degree of freedom	175	
t-Statistics	6.98	
$P(T \le t)$ one-tail	0.00	
t-Critical one-tail	1.65	
$P(T \le t)$ two-tail	0.00	
t-Critical two-tail	1.97	

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with unequal variance. I reject the null hypothesis in favour of the alternative that the uproot-plant sample mean is significantly different from the graft sample mean at the 1% level. I evaluate the uproot-plant strategy versus the graft strategy in my model in chapter 5 and compare the results with what orchardists adopt in practice.

Another key parameter is the discount rate, which is the main foundation of valuation in finance. In the real options literature, there is a range of different approaches of adopting the discount rate: (1) to assume an ad hoc discount rate and then changing it in a sensitivity analysis framework (see Majd and Pindyck (1987), Song, Zhao, and Swinton (2011), Schmit, Luo, and Conrad (2011), Shah and Ando (2016), M. Zhang, P. Zhou, and D. Zhou (2016), Hossiso and Ripplinger (2017)), (2) to use a proxy such as central bank discount rate with/without modification (see Odeninga, Muhoff, and Balmannb (2005), Ihli, Maart-Noelck, and Musshoff (2013)), (3) to approximate a long term historical average bond returns as a measure of opportunity cost, (4) and finally to use a required return either capital asset pricing model (CAPM) or dividend growth model (DGM) (see Purvis, Boggess, Moss, and Holt (1995), and Hauer, Luckert, Yemshanov, and Unterschultz (2016)). The main driving forces to specify the discount rate are the assumptions of risk preference and complete markets (Bulan, Mayer, and Somerville, 2009). The two well-known models for this purpose are the DGM and the CAPM. The DGM states that the required return of a security is the sum of its dividend yield and capital gain yield, $R = \frac{D_1}{P_0} + g$. Table 4.11 presents the parameters for estimating the discount rate. I estimate D_1 as the price of renting land (\$/ha), P_0 as the average 10-year price of land (\$/ha), and finally g is the average drift across both apple types (Gala and Braeburn). Note that leasing land for livestock farming is much easier compared to agriculture, due to the fast and easy switch process. A land for running an orchard needs at least 5-6 years time to become productive and it would not be viable to lease empty land for a shorter period of time. So, the contracts for this purpose do not follow an infor-

TABLE 4.11: The DGM Approach to Estimate Discount Rate.

This table presents the parameters to estimate discount rate based on the DGM model for Chapter 5. Here, D_1 is the price of renting land, P_0 is the average 10-year price of land (\$/ha), and finally g is the average drift of both apple types.

Parameters	Symbol	Value
10-years Average sales (\$/ha)	D_1	167600
Horticulture Rent (\$/ha)	P_0	1114
Average Drifts	g	0.06

mal procedure of negotiation but a more formal one with tight rules and regulations through agencies to avoid any conflict. To this end, I estimates 6.87% discount rate according to the DGM.

The second approach, the CAPM model, states that the required return is the summation of the risk free rate and the risk premium from investing in the market as follows:

$$r_i = r_f + \beta \left(r_m - r_f \right) \tag{4.2a}$$

$$\beta = \frac{\rho_{i,m}\sigma_i\sigma_m}{\sigma_m^2},\tag{4.2b}$$

where r_i is the required return on apple i, r_f is the risk free rate, r_m is the expected return on market, β is the market risk, $\rho_{i,m}$ is the correlation between apple i's returns and the market return, σ_i is apple i's volatility, and σ_m is the market volatility. Table 4.12 shows the estimated parameter to use for the CAPM model. I use numbers from Frijns and Tourani-Rad (2016) to measure the market risk (σ_m) and risk premium ($r_m - r_f$) in the New Zealand market. I also measure the risk free rate (r_f) from reserve bank treasury bills and the correlation through the time series of stock market return and apple revenues ($\rho_{i,m}$). I estimate α 1.99% discount rate using the CAPM model, which is lower than the first method. However, the discount rate from DGM approach fits better in my model due to being

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TABLE 4.12: The CAPM Approach to Estimate Discount Rate.

This table presents the parameters to estimate discount rate based on the CAPM model for Chapter 5. I use numbers from Frijns and Tourani-Rad (2016) to measure the market risk (σ_m) and risk premium ($r_m - r_f$) in the New Zealand market. I also measure the risk free rate (r_f) from reserve bank treasury bills and the correlation through the time series of stock market return and apple revenues ($\rho_{m,i}$).

Parameters	Symbol	Value
Risk Premium (Geometric Mean-Frijns and Tourani-Rad (2016))	$r_m - r_f$	0.03
Correlation (Stock, Avg Crop Prices)	$ ho_{m,i}$	-0.22
Risk Free (Reserve Bank)	r_f	0.02
Average Crops Variances	σ_i	0.11
Market Volatility (Frijns and Tourani-Rad (2016))	σ_m	0.20
Estimated Volatility in Use	σ_i/σ_m	0.20

more compatible with main parameters: drift, volatility, and correlation.

Chapter 5

Building A Complex Switching Option Model And Its Application In Horticulture

5.1 Introduction

There has been a considerable increase in the number of papers on real options over the past decades. Hence, my study stands on the shoulder of large body of literature. Reviewing the literature indicates that this theory is analysed well. However, there is a gap in the literature to see whether practitioners adopt real options valuations in their decisions. Basically, to what extent does the actual behaviour (what practitioners do) deviate from optimal values (suggested by theory)? My survey on empirical real options papers shows that the use of a case study is to extract the variables of interest and employ them to solve a valuation problem (I describe some of them in the following discussion). The unanswered question is to what extent people follow real options valuation?

Survey papers confirm that there is such a gap in the literature (see, Cetinkaya and Thiele (2014), Brounen, Jong, and Koedijk (2004), Brunzell, Liljeblom, and Vaihekoski (2013) Savolainen (2016), Kozlova (2017), Lambrecht (2017), Trigeorgis and Tsekrekos (2018)). This study aims to fill this gap and contribute to the literature by setting up a switching options problem for orchards and test if orchardists follow real options strategies in their exercise decisions. These orchardists are active in apple orchards where there is an uncertainty about apple revenue changes in future.² Orchardists allocate their orchards to a specific type of apple or to a combi-

¹ Slade (2001), Purvis, Boggess, Moss, and Holt (1995), Hyde, Stokes, and Engel (2003), Tauer (2006), Tubetov, Musshoff, and Kellner (2012), Shah and Ando (2016), Winter-Nelson and Amegbeto (1998), Cunha-e-sa and Franco (2016), Fackler and Livingston (2002), Kennedy (1986), Williams and Wright (1991), Deaton and Laroque (1992), Anderson and Brorsen (2005), Hagedorn, Irwin, Good, and Colino (2005), H. S. Kim and Brorsen (2012), Song, Zhao, and Swinton (2011), Dixit (1992), Tyner (2008), James, Swinton, and Thelen (2003), Song, Zhao, and Swinton (2011), Livingston, Roberts, and Y. Zhang (2014), Pederson and Zou (2008), Schmit, Luo, and Conrad (2011) and Schmit, Luo, and Tauer (2009), Hauer, Luckert, Yemshanov, and Unterschultz (2016), and Hossiso and Ripplinger (2017).

² In this stud, revenue refers to net revenue (profit), otherwise I will clearly state the term gross.

nation of apple varieties in an analogous fashion to a portfolio. As time goes by, and apple revenues change, they have an incentive to switch their orchards to the more lucrative apple type. I evaluate a situation where the orchardists have the option to change between Braeburn and Queen apple trees by two competing switch strategies: the uproot-plant or the graft strategy.¹

Grafting is the process of donating a scion, which is a shoot or a bud from another tree, to a rootstock of another tree (the receiver). The uprootplant strategy is relatively more expensive with higher upfront cash cost and a slow process, but results in a rejuvenated orchard, whereas the graft strategy is less expensive and faster but results in carrying on with the existing trees. In this regard, my study not only captures the value of waiting for an investment under uncertainty but also evaluates alternative switch options to optimise the orchardists' land use polices. In other words, the orchardists in each stage of a tree's lifetime have 3 options: to continue with the existing apple trees, to switch to a new type of apple by uprooting the existing apple trees and planting the more lucrative apple type, or to switch by grafting on existing rootstock. In addition, I relax the assumption of irreversibility and allow my orchardists to have the option to reverse their decisions and go back to the old apple type.² The *hysteresis* effect is revealed through orchardists waiting to exercise the uproot-plant or graft strategies.³ For instance, there might be a situation where the orchardists can exercise their option and switch to a more lucrative apple type by uprooting and planting new trees, or by grafting, but they choose to continue with the existing apple trees until the revenues hit the hurdle

¹ The choice of two apple types is explained in Chapter 4.

² I explained in chapter 3 that irreversibility can be discussed in a spectrum from complete reversibility to complete irreversibility depending on the cost of switching. Reversibility in this chapter means to go back to the original apple type but by paying cash cost and time cost.

³ Hysteresis here means that the switching and implementing decisions will not happen immediately and occurs over time. Therefore, farmers think twice before they switch from one type of apple to another.

rate, which is higher than the Marshallian rule. Hence, I contribute to the literature not only by examining whether practitioners follow real options valuations in practice and how far they are from theoretical values but also by capturing the exact cash cost and time cost of prolonged switching options under uncertainty and reversibility.

In the Section 3.2, I evaluate a situation where representative agents have the option to switch from one product to another, by proposing three models: the cash cost, time cost, and projection models. The assumptions are: (1) products have an infinite lifetime and (2) there is only one switching strategy (removing the existing product and initiating the other product). In contrast, in this chapter, I assume: (1) apple trees have a finite productive lifetime similar to what I see in reality and (2) the model is more complex due to having two competing switch strategies: the uproot-plant and graft strategies. However, I conjecture that the uproot-plant strategy is the optimal policy when orchards are old whereas the graft strategy is optimal when orchards are young. To begin with, I present a short survey of recent papers on empirical real options to clarify how case studies have been utilised in the literature.

Some real options studies are implemented in the mining sector or examine managerial flexibility. For instance, Sina and Guzman (2019) evaluate an open pit mine with two processing method using real options. This model optimises the mine value by assuming commodity price uncertainty and unlimited extraction capacity. They use a numerical example to test their model, and results show that having two processing methods instead of one provide a higher mine value. Rebiasz, Gawel, and Skalna (2017) evaluate the managerial flexibility for steel industry investments. In this study, the case study of a product switch is valued using Monte Carlo simulation. Their results indicate that having the option to diversify products generates a significant value for metallurgical projects. Kim, Ha, and Hyoungkwan (2017) examine managerial flexibility to find the effective

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adaptation strategy in infrastructure using a chooser option. They implement their study on a case study in Seoul and find that having the chooser option increases the project value, which emphasises the strength of real options rather than comparing optimal and actual behaviours. Climate change and renewable energy are other applications of real options. X. Wang and Du (2016) study the carbon capture and storage investment decision for a power plant in China. They use their case study to estimate the corresponding model parameters and examine the decision making using both net present value and real options approaches. This study follows a scenario analysis and emphasises on the strength of real options in policy making. M. Zhang, P. Zhou, and D. Zhou (2016) examine investment in renewable energy using a real options approach. They use an application to solar photovoltaic power generation in China to capture their parameters. The authors show that policies such as subsidies, technological progress, and market stability are more important than price volatilities and investment costs.

In the agriculture sector, Ewald, Ouyang, and Siu (2016) employ the real options approach to optimise the harvesting problem (lease vs ownership) for fish farms in Norway. The variables of interest are captured from a newly-established market in salmon futures. The authors show that the existence of salmon futures market and managerial flexibility play key roles for risk averse farmers. In corporate finance, Davis and Cairns (2017) examine a lumpy investment in capital budgeting using irreversible and reversible real options approach. They use the parameters from Brennan and Schwartz (1985) to study completely irreversible investments versus completely reversible investments. A recent paper in investment and disinvestment is Tang, H. Zhou, Chen, K. Wang, and Cao (2017). They study the optimal overseas investment of China using real options under uncertainty and irreversibility. They employ a case study to present the optimal investment through scenario analysis. Results show a higher risk for earlier start-up projects and evaluate the factors that affect on results. My

study is different from the above mentioned literature by proposing three models based on real option approach and comparing results from theory with my case study in orchards.

The remainder of this chapter proceeds as follows: Section 5.2 describes the horticulture sector in New Zealand and some characteristics of apple orchards. I explain two strategies that are adopted by orchardists to switch from one type of apple to another. Section 5.3 presents the methodology that I follow to analyse the orchardists investment decisions in apple orchards. Section 5.4 proposes empirical results and examines how the optimal land use policy proposed by model is different from what farmers do in practice. Finally 5.5 draws conclusions and discusses the possible solutions and suggestions to the orchardists and authorities to make optimal investments.

5.2 Apple Orchard Characteristics

5.2.1 Horticulture Sector

According to New Zealand Horticulture (2017), the product of the horticulture sector exceeds \$8.8 billion in 2017. The dominant products are wine, kiwifruit, and apples. The global demand for New Zealand apples is growing dramatically due to their high quality. As a consequence, apple growers are improving their level of technology and improving practices to keep up with this demand. However, these improvements have made the switching between varieties of apples more frequent than before.

The apple industry in New Zealand developed its commercial role in international markets during the 1910's by establishing the New Zealand Fruitgrowers' Federation (NZFF) and the Fruit Export Control Board (FECB). In recent years apple growers have become concentrated in areas that pro-

vide higher returns due to comparative advantage compared with other regions. In this regard, the Hawke's Bay and Nelson regions account for almost 5115 ha (60%) and 2496 ha (30%) of total land (8845 ha) that is allocated for growing apples (Horticulture, 2017). The Hawke's Bay is a region with fertile soils, high sunshine, and great climate for agriculture (ANZ Agri Focus, 2015). The current situation makes this region an ideal case study to examine the orchardists' behaviour in terms of exercising switching options.

There are a wide range of factors that drive apple growers to switch to new apple types. For example, the demand or market for specific types of apple, available government financing and etc. However, it is not always possible to invest and redevelop any further in large scale orchards due to financial and time constraints as well as missing the cash flow stream. Therefore, orchardists look for an adjustment such as grafting or modifying training and pruning techniques for existing trees (MPI, 2017). Originally, orchardists used to plant trees that grew slowly but eventually became big and tall due to wide spacing. As time changed, new standards have come through and competition pushes orchardists to choose a type of rootstock that allows them to grow trees much more closely and naturally have them not so tall with no shade, higher density of trees, and higher productivity. This also saves on labour costs due to the trees being much easier to spray, trim, and pick without need for ladders. The strategy of having high density orchards is one of the common management systems that help growers to profit from high early production. Orchardists can speed up the time to become productive by using a dwarf or semi-dwarf rootstock.

5.2.2 The Uproot-Plant Versus The Graft

Orchardists have the option to uproot old apple trees quickly and cheaply, then plant new trees very slowly and expensively due to the high upfront cash cost of planting and the time required for the trees to reach maturity. The alternative way is to graft branches onto an existing tree. This is less expensive and faster but continues with the existing trees. In the grafting process, the new scions grow up on old trees and save time as they use the original root systems. The yield during the early year of a tree's lifetime depends on a wide range of factors. However, the relationship between tree density and cumulative yield follows the law of diminishing returns. In the first stage of planting new trees, there is an upward sloping relationship then it follows a constant relationship. In other words, adding more density will not increase the yield to compensate the extra unit cost. Moreover, some other variables such as water availability, soil type, climate, weather, elevation, direction of slope, irrigation system, and diseases will affect on productivity. Assuming all the conditions to be constant, the cash flows throughout the trees' lifetime depend on the age and how far they are on to the grafting process (if there is any graft). Table 5.1 demonstrates the cash flows of the two competing strategies throughout a tree's lifetime.

Table 5.1 shows that the productivity of an apple tree grows gradually until it reaches full productivity. While the apple trees approach to the end of their lifetime, they lose their productivity gradually. However, in the interim, an orchardist might use the uproot-plant or graft strategies. The uproot-plant strategy is slower in reaching full productivity than the graft strategy. The uproot-plant strategy begins with no productivity in the first year and is expected to grow at 25% in the following years, with the productivity increasing by 25% per year until the productivity reaches to 100%. Similarly, the graft strategy follows the same process but the productivity begins with 25% in the first year and reaches full productivity

TABLE 5.1: Apple Trees Cash Flows Throughout Their Lifetime Using the Uproot-Plant and the Graft Strategy.

This table shows the productivity of an apple tree throughout its lifetime according to 3 options. The productivity grows gradually according to the age of the apple tree until it reaches full productivity, when orchardist keep the existing trees. Note that the productivity decreases gradually as apple trees approach to the end of their productive lifetime. If orchardists uproot the existing trees, they lose the productivity for the first year and then according to the tree's age it increases gradually again. In the grafting option production is a fraction of tree's productivity at that age and depends how far the process is on to the grafting process.

Option				Pro	ductiv	ity in	Lifeti	me		
Keep (g(t))	0%	25%	50%	75% 100%	100%	100%	100%	100%	100%	 0%
Uproot					0%	25%	50%	75%	100%	 0%
Graft (h(t))					25%	50%	75%	100%	100%	 0%

one year earlier. The orchardists can apply graft at any age of productive apple trees. Hence, the trees' production depends on the trees' age and how far they are on to graft process. Apple trees in both strategies begin to lose productivity as they approach to the end of their lifetime until they die. One would expect the orchardists prefer to employ the graft strategy when trees are young and uproot them when they are old. Table 5.2 shows a summary of initial costs to establish an orchard. The cost varies based on the root stock, soil quality, water availability and etc.

5.3 Methodology

In this chapter, I look at a real options switching problem where the orchardists have blocks of land and have the option to grow different types of apples. I examine the orchardists' behaviour in the Hawke's Bay region, where they switch from one type of apple to another according to apple revenue changes. I focus on Braeburn and Queen types due to their higher switch frequency and more volatile revenues in the dataset. The

TABLE 5.2: Initial Cost to Establish an Orchard.

This table presents a summary of initial costs to establish an orchard. These costs can change depending on the root stock category, soil quality, water availability and etc.

Item	\$/Tree	\$/ha
Trees	\$40-\$50	\$16000-\$20000
Planting Costs (Labour & equipment)	\$15-\$18	\$6000-\$7000
Fertiliser	\$2.50-\$5.00	\$1000-\$2000
Mulch	\$7.30	\$2900
Tree shelter	\$15	\$6000
Orchard shelter	\$3.10	\$1250
Total	\$83-\$98	\$33150-\$39150
Site Specific Costs		
Irrigation	\$13-\$20	\$5000-\$8000
Drainage	\$5-\$10	&2000-&4000
Contouring	\$13-\$25	\$5000-\$10000
Total	\$113-\$153	\$45150-\$61150

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orchardists plant Braeburn apple trees, which generates revenue S_B , or Queens, with revenue S_Q . However, with the passage of time and changes in competing apple revenues, they find themselves in a situation where they have the option to switch to a more lucrative apple type.

I assume a complete market with two agents. The role of risk aversion depends on whether the market is complete. In a complete market, all products are tradable and agents maximise the value of their assets and work with risk neutral probabilities. In contrast, in an incomplete market, all products are not tradable and one should work with utility function, marginal utility, consumption, endowments and etc. Also, I consider orchardists who have access to a parcel of land. At any point in time, this parcel of land can be allocated in a variety of states, which will be discussed subsequently below. The land value in all cases depends on the current value of two apple types. The underlying assumption here is that apple revenues follow geometric Brownian Motions (GBMs) as follows:

$$dS_B = \mu_B S_B dt + \sigma_B S_B dz_B$$

$$dS_Q = \mu_Q S_Q dt + \sigma_Q S_Q dz_Q,$$
(5.1)

where dS_B and dS_Q represent apple revenue changes, μ_B and μ_Q are apple drifts, S_B and S_Q are apple revenues, σ_B and σ_Q are volatilities of apple prices, dt is the small interval of time, and finally dz_B and dz_Q are increments of Brownian motions (that have correlation ρ_{BQ}). The land value in my problem is derived based on the apple revenue processes (5.1) over time. In this chapter, I embellish my model from the third chapter, where

land value is governed by the operator \mathcal{L} .

$$0 = \frac{\partial V_i}{\partial t} + \mu_B S_B \frac{\partial V_i}{\partial S_B} + \mu_Q S_Q \frac{\partial V_i}{\partial S_Q} + \frac{1}{2} \sigma_B^2 S_B^2 \frac{\partial^2 V_i}{\partial S_B^2} + \frac{1}{2} \sigma_Q^2 S_Q^2 \frac{\partial^2 V_i}{\partial S_Q^2}$$

$$+ \rho \sigma_B S_B \sigma_Q S_Q \frac{\partial^2 V_i}{\partial S_B \partial S_Q} - r V_i + C F_i$$

$$\equiv \frac{\partial V_i}{\partial t} + \mathcal{L} V_i + C F_i,$$
(5.2)

where V_i is the land value in state i, \mathcal{L} is the operator:

$$\mathcal{LV} = \mu_B S_B \frac{\partial V}{\partial S_B} + \mu_Q S_Q \frac{\partial V}{\partial S_Q} + \frac{1}{2} \sigma_B^2 S_B^2 \frac{\partial^2 V}{\partial S_B^2} + \frac{1}{2} \sigma_Q^2 S_Q^2 \frac{\partial^2 V}{\partial S_Q^2} + \rho \sigma_B S_B \sigma_Q S_Q \frac{\partial^2 V}{\partial S_B \partial S_Q} - rV,$$

 ρ is the correlation between apple types, r is the discount rate, and the CF_i is an extra term compared with a conventional option pricing formula, that captures the revenues generated in state i. I employ the Peaceman and Rachford (1955) alternating direction implicit (ADI) method to solve the PDE (5.2) and find the land value in different states of activity (see 3.24 and 3.25). In the third chapter (3.5), the cash flows depend only on the states. In contrast, in this chapter, the cash flows are more complex and not only depend on the states but also on the age of apple trees and the switching strategy. I define the cash flows as follows:

$$CF_{i}(t, y_{B}, y_{Q}) = g(t)e^{(e_{iB}y_{B} + e_{iQ}y_{Q})}$$

$$CF_{i}^{graft}(t, b, y_{B}, y_{Q}) = g(b+t)h(b)e^{(e_{iB}y_{B} + e_{iQ}y_{Q})},$$
(5.3)

where g(t) is the fraction of a fully productive cash flow that a t years old tree generates, $CF_i^{graft}(t,b,y_B,y_Q)$ is the cash flow of a tree that is b+t years old and b years into a graft, h(t) is the fraction due to grafting, and $e_{iB}y_B+t$

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 $e_{iQ}y_Q$ is a term derived from transformed apple prices times eigenvalues.

In this chapter, I set up a dynamic optimisation model to find the optimal land value. The model examines the orchardists' behaviour in an uncertain environment where they allocate land for growing one of two competing apple types, Braeburn or Queen. However, there are situations where the orchardists realise having apple trees of the other type is more lucrative than their current type. In the model, there are 6 states for the land: two productive states, two unproductive states, and two grafting states. Orchardists in productive states B and Q, grow apple types Braeburn or Queen. The unproductive states consist of the *Transition* state where the land has no trees on it and the Deceased state where the land has unproductive (dying) trees. Finally, the grafting states entail grafting Braeburn or Queen branches on to existing trees. In the following, I describe each state and the corresponding strategy for the orchardists. I begin with unproductive states (Transition and Deceased) as they are the beginning and the end of lifetime of a tree and account for terminal values later in my discussion.

Transition State is an unproductive state where the orchardists have land with no apple trees but are ready to plant either Braeburn or Queen apple trees. The transition land value solves:

$$\mathcal{L}V_{Tr} \le 0$$

$$V_{Tr} \ge V_{B,p}(0) - K_{B,p}$$

$$V_{Tr} \ge V_{Q,p}(0) - K_{Q,p},$$
(5.4)

where V_{Tr} is the value of the land in the transition state, $V_{B,p}(0)$ and $V_{Q,p}(0)$ are the values of the land where Braeburn and Queen apple types have just been planted respectively (but have not yet reached productivity), $K_{B,p}$ and $K_{Q,p}$ are the upfront cash costs of planting apple types Braeburn and Queen respectively. In the transition state, the orchardists' options are to plant Braeburn apple trees or to plant Queen apple trees. The planting is

relatively expensive but is instantaneous, and results in young (unproductive) apple trees with a long life ahead of them.

Deceased State is the other unproductive state where orchardists have land with unproductive old apple trees (age above T^{Die}) that generate no cash flows. In this state, the only option available is to uproot the trees and convert the land into the transition state. The land value in this state obeys:

$$\mathcal{L}V_D \le 0$$

$$V_D > V_{Tr} - K_u,$$
(5.5)

where V_D is the deceased land value, and K_u is the upfront cash cost of uprooting that is the same for both tree types.

Productive States are situations where orchardists have productive apple trees growing Braeburn or Queen apples. Orchardists receive cash flows from Braeburn in the first productive state and Queen in the second productive state. Productive states are represented as a continuum of states accounting for the tree's age. Therefore, the land values are conditional on tree's type and age. I compute the land use value at each age of apple trees with the deceased land value (V_D) accounting for the terminal value (I solve my dynamic model by backward induction). Orchards in this state have two options, one to uproot the trees, and the second to begin grafting. The land value solves:

$$V_{i,p}(T^{Die}) = V_D$$

$$\mathcal{L}V_{i,p}(t) + CF_i(t, y_B, y_Q) + \frac{\partial V_{i,p}}{\partial t} \le 0$$

$$V_{i,p}(t) \ge V_{Tr} - K_{i,u}$$

$$V_{i,p}(t) \ge V_{j,q}(0, t) - K_{j,q},$$

$$(5.6)$$

where $V_i(t)$ is the planted land value at age t, $K_{i,u}$ is the upfront cash cost

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of uprooting in state i, $V_{j,g}(0,t)$ is the value of the land with a fresh graft on a t year old tree, and $K_{j,g}$ is the upfront cash cost of grafting with type j apples.

The orchardists have two alternative strategies to switch from one type of apple to another: uprooting and planting new trees, or by grafting. They uproot their trees by paying the upfront cash cost of uprooting and moving to the transition state, where land is ready for planting either apple trees (see equation (5.4)). Alternatively, an orchardist can graft new branches on to existing rootstock. Technically, this process is relatively less expensive and faster than the uproot-plant strategy but continues with the existing trees. If the orchardists have lands with Braeburn apple trees, and realise Queens are more lucrative, they can graft Queen branches on to the rootstock of their existing apple trees. The orchardists pay the upfront cash cost of grafting and receive reduced cash flows for T^{graft} years until the apple trees reach full productivity. By the time the grafting process is completed, the orchardists have apple trees of the alternative type whose age is T^{graft} years older. The value of the land in the beginning of the graft on to a tree with b years age is $(V_{i,q}(0,b))$, where land with grafted trees obeys:

$$V_{i,g}(T^{graft}, b) = V_{i,p}(b + T^{graft})$$

$$\mathcal{L}V_{i,g}(t, b) + CF_i^{graft}(t, b, y_B, y_Q) + \frac{\partial V_{i,g}(t, b)}{\partial t} \le 0$$

$$V_{i,g}(t, b) \ge V_{Tr} - K_u$$

$$V_{i,g}(t, b) \ge V_{i,g}(0, b + t),$$

$$(5.7)$$

where b is the tree's age at the start of grafting, t is the time spent grafting, $V_{i,g}(t,b)$ is the value of the land with grafted trees, $V_{i,p}(b+T^{graft})$ is the value of the land with $b+T^{graft}$ year old trees that accounts for the terminal value of grafting, and $V_{j,g}(0,b_t)$ is the value of the land if the orchardist started grafting the other apple type when the tree was age b+t.

Note that I use the ADI method to solve the system of equations in (5.4) and (5.5) (transition and deceased states), which are elliptic problems whereas, (5.6) and (5.7) (productive states), are parabolic problems. I solve (5.4) and (5.5) by including a $\frac{\partial V}{\partial t}$ term but iterating until this becomes 0.

Intuitively, the orchardists switch from one type of apple to another only if the competing apple type is sufficiently lucrative to cover the cash cost and the time cost of switching. Both the uproot-plant and graft strategies entail a long period of time before reaching fully productive apple trees. The uproot-plant strategy is more expensive and slower to reach full productive apple trees, but has the advantage of rejuvenating the orchard. In contrast, the graft strategy is relatively less expensive and faster to reach full production due to keeping the rootstock of existing apple trees but continues with the existing trees. However, neither the uprootplant nor the graft strategy is able to suppress the level of uncertainty and the time cost of switching. The level of uncertainty for options to switch is higher in horticulture compared with other sectors due to a prolonged switching process. When orchardists switch from one type of apple to another, they lose the existing cash flows and by the time they complete the process, apple revenues might no longer be in their favour. This model is able to capture the exact time cost of this process. Orchardists in the model have the option to reverse the switching process at any stage of the process. However, reversing the switching process and going back to the existing apple type (either by grafting or uprooting) is again time consuming.

5.4 Empirical Results

In this section, I report the results of my switching real options model. The model outputs entail the land value for each apple type (Braeburn and Queen) at each age according to keeping the existing apple trees (Tval),

switching to the competing apple type by the uproot-plant strategy (Uval), switching by the graft strategy (Gval), and the optimal land value (highest value across available options). My results demonstrate the optimal exercise frontiers and what policy (Tval, Uval, and Gval) is ideal when apple revenues change. I compare optimal land value from my switching real options problem with what value orchardists achieve in the orchards and examine how their choice is different from the model. Employing the conditional logit regression model helps us to test the behaviour of orchardists, conditional on the money at stake along with their size.

Table 5.3 presents the set up parameters for apple types Braeburn and Queen. It costs \$10000 per hectare and 3 months to uproot and costs \$100000 per hectare to plant each apple type. In contrast, grafting costs \$35000 in 4 years, which is faster and less costly. Braeburn and Queen have drifts of 4% and 6% with volatilities of 8% and 13% respectively. The correlation between two apple types is 0.26 in an economy with a 7% discount rate.

Each apple tree in my model generates three output results: Tval, Uval, and Gval. All three outputs are generated over the age of apple trees from 1 year to 34 years. Hence, there are 204 different output for each orchard-time observations with two apple types each with 34 years and 3 different situations (Tval, Uval, Gval). Also, the results include *transition* and *deceased* states. Table 5.4 demonstrates the summary of outputs model. On average, apple trees in the sample are in their middle age and full productivity. Comparing the available land use values (Tval, Uval, Gval) shows that on average, keeping the existing trees has the highest land value with less deviation from its mean compared with switching to competing apple type (either the uproot-plant or the graft). Results show that the strategy chosen by orchardists is different from the optimal policy of the model. However, it is not possible to reject the choice of orchardists

¹ In this study, I assume that all apple varieties require the same cost and time to grow and have the same productivity.

TABLE 5.3: Queen and Braeburn Apple Types Parameters.

Note: This table shows the parameters for apple types Queen and Braeburn in my analyses. The variables of interest are: cost, time, volatility, drift, discount rate, and correlation. All values are scaled based on thousands dollar per hectare in years.

Parameters	Symbol	Braeburn	Queen
Cost (uproot, plant)	$K_{u,p}$	(10, 100)	(10, 100)
Time (uproot, plant)	$T_{u,p}$	(0.25, 5)	(0.25, 5)
Cost (Graft)	K_g	35	35
Time (Graft)	T_g	4	4
Volatility	σ	8%	13%
Drift	μ	4%	6%
Discount Rate	r	7%	
Correlation	ho	0.26	

due to having assumptions in the model set up, and the psychological environment. Also, orchardists may diversify their orchards across different apple varieties that affects on their decisions. The model implies that minimum and maximum percentage difference dollar value is 0% and 8.4% per hectare respectively and on average, 0.5% per hectare. The amount of percentage difference dollar value for orchardists varies based on the land area and the price of apple type in that year. Orchardists pick 3104 tray carton apples in each hectare and Braeburn apple gross revenues vary from \$54186.28 to \$86485.62 whereas Queen apple gross revenues vary from \$64596.38 to \$148754.60 per hectare over years 2006-16 respectively (see Chapter 4). Hence, according to the current model set up, on average an orchardist with a hectare land area loses 0.5% of the corresponding apple net revenue. Assuming the same cost of \$50 for both apple types, on average the loss varies varies from \$220.93 to \$382.43 for Braeburn and from \$272.98 to \$693.77 for Queen.

Figure 5.1, shows the optimal land use policy for Braeburn and Queen

TABLE 5.4: Sample Results of Dynamic Optimisation.

This tables presents the results of land use optimisation. I report key variables: age of apple trees (Age), land use for keeping the existing apple trees type 1 or 2 (Tval), land use for the competing apple type by using the uproot-plant strategy (Uval), land use for the competing apple type by using the graft strategy (Gval), optimal value amongst all 3 options (optimal), the percentage difference dollar value between the optimal land use and what orchardists choose (Diff), and finally the land area (Area). All values are scaled based on thousands dollar and hectares.

Variable	N Obs	Mean	Std Deviation	Min	Max
Age	5188	17.24	7.95	1.00	33.00
Keep the Existing Trees Value (Tval)	5188	2283.50	347.47	1555.62	2634.67
Uproot-Plant Value (Uval)	5188	2109.36	318.57	1771.15	2505.03
Grvaft Value (Gval)	5188	2187.60	243.85	987.95	2653.04
Optimal	5188	2301.80	333.83	1271.15	2634.67
(Optimal-Actual)/Actual	5188	0.005	0.015	0.000	0.084
Area	5188	1.02	1.25	0.02	22.58

at age 5, 15, 25, and 33 years old. In this figure, orchardists grow Braeburn apple type (x axis) and Queen is the competing type (y axis), which is more lucrative. The axes are dollar value prices of apple types, scaled on 0 to 30 dollars. Orchardists receive the Braeburn revenues over time by keeping the existing trees. However, when the competing apple type (Queen) is more lucrative, they have the incentive to switch either by the uproot-plant or the graft strategy (y axis). The former is an expensive and slow process but results in a rejuvenated orchard while the latter is less expensive and faster but continues with the existing trees. Having said that, orchardists exhibit inertia, keeping their existing apple type and not exercising their options except when the competing apple type is really in the money. Orchardists have a hurdle rate in their mind and they exercise their option when revenues hit the exercise frontier. Figure 5.1 shows that when orchardists have orchards with 5-year old apple trees they are more likely to keep the existing apple tree (Braeburn) but when the Queen type is more lucrative, they exercise their option and switch by grafting. As apple trees become older and they are 15-years old, the grafting becomes

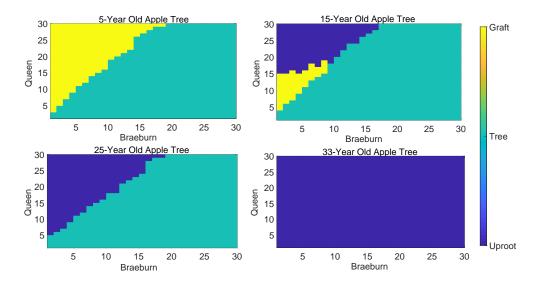


FIGURE 5.1: Optimal Land Use Policy. This figure shows orchardists who grow Braeburn apples and optimise land use policy as the apple net revenues change over time. Orchardists look at the net revenues of both apple types and decide to keep the existing apple trees or switch to the competing apple type. In this figure, as the Braeburn net revenues grow (corresponding land value on x axis) orchardists prefer to keep the existing apple trees. However, looking at Queens net revenues (corresponding land value on y axis) shows that when net revenues hit the hurdle they are willing to switch to the competing apple type. For a 5-year old apple tree the switch is through the graft strategy, but as the trees become older (25-year) and Queen is more lucrative, the uproot-plant becomes a better alternative way to switch having the advantage of rejuvenating the orchard. As apple trees near their unproductive age, grafting is not an ideal switching strategy any more and uprooting is the optimal policy. By the time trees die (33-years), the optimal policy is to uproot them and plant new trees. Note that the land value of apple types on each axis are scaled on 30 dollars value.

the ideal strategy for switching. Note that both switch strategies are available throughout the life of apple trees and the choice of switch strategy is driven by two key factors: the profitability of alternative apple type, and the age of apple trees. Keeping the existing apple type (Braeburn) is more likely throughout the productive life of apple trees. Finally, apple trees at age 33-years old become unproductive and die. Therefore, the optimal policy is to uproot unproductive apple trees and have a rejuvenated orchard. Overall, results indicate that orchardists rarely switch.

I use a conditional logit model to examine orchardists' choices while the choice properties vary across orchardists in my sample data. The conditional logit model is similar to logistic regression but individuals are subject to different situation before they select their choices. In this regard, I offer them the choice of keeping the existing apple trees, uprooting and planting young apple trees, and grafting on the existing apple trees, while using the size of land as a condition. Equation (5.8) presents the conditional logit model set up.

$$U_{choice} = \beta_1 Val_{choice} + \beta_2 D_{Tree} + \beta_3 D_{Tree} * D_{size}$$
 (5.8)

where U_{choice} is the orchardists choice value, Val_{choice} is what policy they adopt whether to keep the existing trees or employ a switching strategy by uproot-plant or graft, D_{Tree} is a dummy for keeping the exiting apple trees that is 1 if it is tree and 0 otherwise, and finally D_{size} is the dummy variable for the land size that is one if it is above 2.02 hectare and 0 otherwise.

Table 5.5 shows the results of logit model. I examine the attributes of orchardists' choice by having orchardists' choice as the response variable and land values, dummy variable for keeping existing trees, and the interaction of keeping existing trees with the large orchardist dummy variable as explanatory variables. The orchardists' choices as the response variable are 1, 2, or 3, which are keeping the existing trees, switching to the competing apple type using the uproot-plant strategy, and switching by employing the graft strategy respectively. The dummy variable for the situations where orchardists keep the existing trees is 1 and 0 otherwise. To test how this choice might be different between small orchardists versus large orchardists, I define a dummy variable based on the land size. If orchardists' land area falls in the top 10% of sizes (above 2.02 ha) dummy is 1 and 0 otherwise. I also include the interaction of dummy for keeping the existing apple trees and large orchardists in my regression to test if they are biased in favour of any choice. My results show that orchardists are in favour of keeping the existing apple trees rather than switching to the competing apple type, using the uproot-plant or graft strategies. Results are significant at the 1% level. However, comparing my results for small orchardists versus large orchardists do not reveal any significant re-

TABLE 5.5: Conditional Logistic Regression Results.

This table presents the results of orchardists' choice of land use when the attributes of choice vary across orchardists. I run a conditional logit regression model and report the estimate values, standard errors and corresponding P-values. The first variable entails the corresponding values of keeping the existing apple trees, switching by the uproot-plant strategy, and switching by the graft strategy. The second variable is the dummy variable that is 1 if orchardists keep the existing trees and 0 otherwise. Finally, the third variable is the interaction of the dummy variable for keeping the existing apple trees and the large size land dummy. The land area above 2.02 ha (top 10% area) accounts for large orchardists. Hence, the dummy is 1 if the land area is above 2.02 ha and 0 otherwise

Parameters	Estimate	P-Value
Land Values	0.0037	0.2303
Tval Dummy	5.2578***	<.0001
Tval Dummy*Large Orchardist Dummy	17.5087	0.9960

sults. Therefore, the bias towards inertia is not more significant for small orchardists compared with large orchardists.

5.4.1 Sensitivity Analysis

An Increase in uproot-plant Cash Cost decreases the probability of switches by this strategy due to higher switching costs. I increase the cash cost of uprooting from \$10000 to \$15000 and planting from \$100000 to \$150000 respectively. Table 5.6 presents the results after increasing the uproot-plant cash cost. Results show that on average keeping the existing trees has the highest value compared with the uproot-plant and graft strategies. The reason is the higher cash cost to pay and the time cost to wait. On average the value of graft strategy is more than the uproot-plant strategy due to increasing the cash cost of switching. After an increase in cash cost, on average, the percentage difference dollar value between the optimal and actual values decreases from 0.5% to 0.2% that might be due to having less switches after increasing the cash cost. In this case, by assuming the same cost of \$50 for both apple types, orchardists lose between \$58.37

TABLE 5.6: Uproot-Plant Cash Cost Change Sensitivity.

This tables presents the results of land use optimisation when the uproot cash cost increases from \$10000 to \$15000 and the plant cash cost from \$100000 to \$150000 respectively. I report key variables: land use for keeping the existing Braeburn or Queen apple trees (Tval), land use for the competing apple type by using the uproot-plant strategy (Uval), land use for the competing apple type by using the graft strategy (Gval), optimal value amongst all 3 options (optimal), and finally the percentage difference dollar value between the optimal land use and what orchardists choose (Diff). All values are scaled based on thousands dollar and hectares.

Variable	N Obs	Mean	Std Deviation	Min	Max
Keep the Existing Trees Value (Tval)	5188	911.54	158.87	461.68	1268.94
Uproot-Plant Value (Uval)	5188	713.62	162.38	567.13	1013.19
Grvaft Value (Gval)	5188	822.62	206.33	167.93	1229.95
Optimal	5188	932.62	158.73	567.13	1268.94
(Optimal-Actual)/Actual	5188	0.002	0.07	0.00	0.50

and \$122.92 per hectare for Braeburn and between \$179.19 and \$247.51 per hectare for Queen over years 2006-16.

Increasing time cost is another key variable that affects on orchardists switches. I increase the time cost of uprooting from 3 months to 6 months and planting time from 5 years to 7 years. Table 5.7 presents the results after increasing the uproot-plant time cost. On average, keeping the existing trees has the highest value compared with the switching strategies (Uval and Gval). On average the value of graft strategy is more than the uproot-plant strategy due to increasing the time cost of switching. On average the percentage difference dollar value between the optimal and actual values decreases from 0.5% to 0.15% (compared with the baseline scenario) which could be due to predicting less switches and keeping the existing apple trees after increasing the time cost. In this case, by assuming the same cost of \$50 for both apple types, orchardists lose between \$31.28 and \$79.69 per hectare for Braeburn and between \$46.89 and \$173.13 per hectare for Queen over years 2006-16.

A decrease in drift is the next sensitivity scenario. Table 5.8 presents

TABLE 5.7: Uproot-Plant Time Cost Change Sensitivity.

This tables presents the results of land use optimisation when the uprooting time cost increases from 3 months to 6 months and planting time cost form 5 years to 7 years respectively. I report key variables: land use for keeping the existing Braeburn or Queen apple trees (Tval), land use for the competing apple type by using the uproot-plant strategy (Uval), land use for the competing apple type by using the graft strategy (Gval), optimal value amongst all 3 options (optimal), and finally the percentage difference dollar value between the optimal land use and what orchardists choose (Diff). All values are scaled based on thousands dollar and hectares.

Variable	N Obs	Mean	Std Deviation	Min	Max
Keep the Existing Trees Value (Tval)	5188	842.44	156.38	370.53	1144.25
Uproot-Plant Value (Uval)	5188	610.73	134.28	489.70	939.27
Grvaft Value (Gval)	5188	695.45	171.47	103.52	1151.02
Optimal	5188	860.62	142.72	489.70	1151.02
(Optimal-Actual)/Actual	5188	0.0018	5 0.088	0.00	0.63

the results when Braeburn drift decreases from 4% to 2% and Queen drift decreases from 6% to 4%. In this scenario, on average, the value of keeping the existing apple trees is more than employing the uproot-plant or graft strategies. On average the graft value is more than the uproot-plant value. A decrease in drift (growth rate) decreases the probability of switches between apple types due to a negative impact on return of apple types. Cash cost and time cost of switching along with less return (growth rate) encourage orchardists to keep the existing apple trees. On average, the percentage difference dollar value between the optimal and actual values decreases from 0.5% to 0.3% (compared with the baseline scenario). In this scenario, by assuming the same cost of \$50 for both apple types, on average, orchardists lose between \$112.56 and \$209.38 per hectare for Braeburn and between \$143.79 and \$396.26 per hectare for Queen over years 2006-16.

Volatility sensitivity evaluates a situation where the baseline volatility parameters increase. Table 5.9 presents the results when Braeburn volatility increases from 8% to 12% and Queen volatility increases from

TABLE 5.8: Drift Change Sensitivity.

This tables presents the results of land use optimisation when drift decreases from 4% and 6% in Table 5.4 to 2% and 4% for Braeburn and Queen apple types respectively. I report key variables: land use for keeping the existing Braeburn and Queen apple types (Tval), land use for the competing apple type by using the uproot-plant strategy (Uval), land use for the competing apple type by using the graft strategy (Gval), optimal value amongst all 3 options (optimal), and finally the percentage difference dollar value between the optimal land use and what orchardists choose (Diff). All values are scaled based on thousands dollar and hectares.

Variable	N Obs	Mean	Std Deviation	Min	Max
Keep the Existing Trees Value (Tval)	5188	1007.24	171.61	611.90	1405.56
Uproot-Plant Value (Uval)	5188	918.45	187.74	750.25	1264.99
Grvaft Value (Gval)	5188	962.37	232.63	247.84	1434.16
Optimal	5188	1036.75	184.81	750.25	1434.16
(Optimal-Actual)/Actual	5188	0.003	0.066	0.00	0.27

13% to 17%. Results show that on average keeping the existing apple trees has a higher value than the uproot-plant and graft strategies. However, the graft strategy has a higher value than the uproot-plant strategy. An increases in volatility provides more time value to the options and results in orchardists being more reluctant to exercise their switching option. In this scenario, by assuming the same cost of \$50 for both apple types, on average, orchardists lose between \$96.30 and \$183.44 per hectare for Braeburn and between \$124.41 and \$351.64 per hectare for Queen over years 2006-16.

An Increase in the discount rate will increase the orchardists' expectation of future returns over the course of time. Hence, when they find a switching option, they are willing to wait and see if it gets even better. Table 5.10 presents the results after increasing the discount rate from 7% to 11%. On average keeping the existing apple trees value is more than the value achieved by the uproot-plant or graft strategies. In this scenario, by assuming the same cost of \$50 for both apple types, on average, orchardists lose between \$80.05 and \$157.50 per hectare for Braeburn and

TABLE 5.9: Volatility Change Sensitivity.

This tables presents the results of land use optimisation when volatility increases from 8% and 13% in Table 5.4 to 12% and 17% for Braeburn and Queen apple types respectively. I report key variables: land use for keeping the existing Braeburn and Queen apple trees (Tval), land use for the competing apple type by using the uproot-plant strategy (Uval), land use for the competing apple type by using the graft strategy (Gval), optimal value amongst all 3 options (optimal), and finally the percentage difference dollar value between the optimal land use and what orchardists choose (Diff). All values are scaled based on thousands dollar and hectares.

Variable	N Obs	Mean	Std Deviation	Min	Max
Keep the Existing Trees Value (Tval)	5188	994.74	142.64	614.60	1329.44
Uproot-Plant Value (Uval)	5188	877.10	152.51	746.45	1163.07
Grvaft Value (Gval)	5188	935.67	198.98	299.90	1327.78
Optimal	5188	1019.29	145.31	746.45	1329.44
(Optimal-Actual)/Actual	5188	0.0027	7 0.06	0.00	0.28

TABLE 5.10: Discount Rate Change Sensitivity.

This tables presents the results of land use optimisation when discount rate increases from 7% in Table 5.4 to 11%. I report key variables: land use for keeping the existing Braeburn and Queen apple trees (Tval), land use for the competing apple type by using the uproot-plant strategy (Uval), land use for the competing apple type by using the graft strategy (Gval), optimal value amongst all 3 options (optimal), and finally the percentage difference dollar value between the optimal land use and what orchardists choose (Diff). All values are scaled based on thousands dollar and hectares.

Variable	N Obs	Mean	Std Deviation	Min	Max
Keep the Existing Trees Value (Tval)	5188	723.04	102.11	251.82	817.58
Uproot-Plant Value (Uval)	5188	537.71	46.25	480.11	605.09
Grvaft Value (Gval)	5188	625.55	110.20	93.18	736.04
Optimal	5188	735.10	77.33	480.11	817.58
(Optimal-Actual)/Actual	5188	0.0028	0.10	0.00	1.40

between \$105.03 and \$307.01 per hectare for Queen over years 2006-16.

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5.5 Conclusion

In this chapter, I construct a real options switching problem by having two types of apple trees (Braeburn and Queen) with a finite life. My model optimises the land use policy with uncertainty and reversibility. I also model two competing switching strategies for orchardists to change their apple types. I implement my model for orchards in the Hawke's Bay region to examine if orchardists' behaviour follows the optimal land use policy that is proposed by my model. Orchardists have the option to keep the existing apple trees, or to switch to the competing apple type using the uprootplant or graft strategies. The former is a more expensive and slower process but results in a rejuvenated orchard whereas the latter is less expensive and faster but continues with the existing trees. My results suggest that orchardists employ the uproot-plant strategy when apple trees are no longer fully productive and are at the last stage of their life. In contrast, orchardists employ the graft strategy when orchards are young. In both strategies the cash flows begin to grow gradually until apple trees reach full productivity but the graft strategy has an earlier production than the uproot-plant strategy due to keeping the existing rootstock.

There are situations in my dataset where orchardists' actions do not match the optimal land use policy proposed by my real options model. According to the current model set up, on average an orchardist with a hectare land area loses 0.5% of the corresponding net revenue of land value that by assuming the same cost of \$50 for both apple types, varies from \$220.93 to \$382.43 for Braeburn and from \$272.98 to \$693.77 for Queen. Orchardists have a tendency towards keeping the existing trees and do not exercise their options to switch to the competing apple type either through the uproot-plant or the graft strategy. This hysteresis might govern their behaviour as long as there is not any force or a strong incentive to convince them to switch. Any financing or incentive from the government may lead orchardists to switch at the right time and decrease their inertia. My results

from the conditional logit model show that orchardists are biased towards inertia. However, these results assume that orchardists do the right thing in the next period of time and theses costs would be small unless they are really in the money. The differences between the optimal land value based on the real options analysis and the actual value that orchardists achieve through their current policy are a proxy for how far they are from the policy advised by my model. However, the model results are based on some assumptions and changing the parameters will minimise the difference between optimal and actual. I test this assertion by changing the feeding parameters in sensitivity analysis section and find lower difference. This shows that orchardists might use different parameters than what I have used in my model. Also, a range of simplifications such as choosing only two apple types and ignoring the fact that orchardists have a portfolio of apple varieties might explain the hysteresis among orchardists that results in difference between theory and practice. Comparing small orchardists with large orchardists reveals the same attitude of being less willing to exercise switching options.

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